

Investing in New York's Future: The Long-Term Benefits of Child Poverty Reduction Policies

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EXECUTIVE REPORT

The Child Poverty Reduction Advisory Council (CPRAC) was established in 2021 to propose policies that reduce child poverty in New York State by half over a ten-year period. In 2024, CPRAC recommended three policy packages that would strengthen state antipoverty policy, either through expansions to current policies or the establishment of new ones (CPRAC, 2024). Within CPRAC's policy packages are various combinations of: (a) expansions to the Empire State Child Credit—the state's version of the federal Child Tax Credit; (b) increases in Public Assistance benefit levels; (c) the establishment of a state food benefit; and (d) establishment of a state housing voucher program. These policy packages are projected to have a substantial effect on child poverty, with the first policy package estimated to meet CPRAC's goal of cutting child poverty in half.

In this report, we seek to understand the long-run benefits and costs of these policy packages. Our analysis finds that:

- The lifetime benefits of implementing these policy packages are roughly ten times greater than the fiscal costs. The greatest net economic benefits to society are derived from Policy Package 1, at an estimated \$94.1 billion (compared to \$8.9 billion in fiscal costs).
- Package 3 generates \$86.3 billion in net economic benefits to society compared to \$8.5 billion in fiscal costs and Package 2 generates \$65.3 billion in net benefits compared to \$6.6 billion in fiscal costs.
- The largest, long-term benefits derive from improvements to children's lifetime health and longevity—\$56.1 billion for Package 1 and increases in children's lifetime earnings—\$22.3 billion for Package 1.
- Taxpayers also experience long term gains—predominantly via reductions in criminal justice expenditures, reduced victim costs of crime, and increased future tax payments from children. The net gains to taxpayers from Policy Package 1, 2 and 3 are respectively \$5.1 billion, \$2.6 billion and \$4.3 billion.

Additionally, our benefit-cost analysis extends these estimates beyond one year of program implementation, aiming to understand the impact of these policy packages if their implementation was not temporary. Specifically, we estimate the long-term benefits and costs of 80 years of continued program implementation, finding that 80 years of continued program implementation could generate roughly \$2 trillion in net societal benefits, with exact estimates varying per policy package. Our benefit-cost analysis thus demonstrates the various ways that CPRAC's recommendations can impact New Yorkers, not just today but well into the future.

SECTION I. INTRODUCTION

In 2021, Governor Hochul established a statewide commission, the Child Poverty Reduction Advisory Council (CPRAC), to make recommendations about how to reduce child poverty in New York State by half within a decade. In 2022, CPRAC members were appointed and the CPRAC began convening to review research on child poverty, anti-poverty policies, and supporting evidence and data. CPRAC also established committees to host focused discussions on specific topic areas and identify initial policy priorities in those areas, which would form the basis for recommendations to reduce child poverty. After considering the priorities identified by its various committees and related reports and reviewing the data provided by the Urban Institute on the costs and poverty reducing effects of various individual proposals, CPRAC developed five packages, or combinations, of different proposals for further analysis. This includes obtaining data on the poverty reduction effects and costs of those packages from the Urban Institute, and requesting an assessment of the long term net social benefits of packages from the Columbia University Center on Poverty and Social Policy. The Urban Institute provided estimates of the costs and poverty reducing effects of the packages. Subsequently, informed by those estimates, CPRAC identified its top three among the five packages for advancement as initial recommendations to reduce child poverty in New York State by up to approximately 50 percent. This report builds on the Urban Institute findings to provide estimates of the long term net social benefits of the three selected packages.

CPRAC proposals would increase either benefit levels or eligibility levels or both of existing state programs and establish new benefits (state housing voucher and state food benefit). Proposals included in the three packages provide both cash and in-kind benefits. Cash benefits include the Empire State Child Credit and changes to New York State's Temporary Assistance benefits and eligibility. In-kind benefits include the state housing voucher and state food benefit. These proposed expansions and new policies are described in more detail in Section II.

Research indicates that increasing the incomes of poor and even middle-class families improves children's health and education and their earnings, health and longevity in adulthood. It also reduces their involvement in the child protective services system and their commission of crimes. Thus, there are good reasons to believe that the long term social benefits of cutting child poverty in half may be large.

One critical starting point for the estimation of net monetary social benefits is the Urban Institute estimates of benefit or transfer costs. Transfer costs (excluding administrative costs) measure the increase in the household incomes of families who receive cash or near-cash benefits from the policies included in the three proposed policy packages. As discussed in Section III below, we assume that near cash and other in-kind transfers increase household income by the amount spent on the transfer.

The Urban Institute estimates of transfer costs provide the estimates of increases in household income. Previous published research and research conducted at the Center on Poverty and Social Policy at Columbia University enable us to estimate how increases in household income translate into long term net social benefits via increases in health, earnings, longevity and decreases in crime, Child Protective Services (CPS) use, and health care costs.

This research is summarized briefly in Section III and in more detail in the technical appendix. In Section IV, we present our estimates of long-term social benefits. We begin with the Urban Institute estimates of transfer costs and poverty reduction and then present our estimates of long-term net social benefits for each of the three policy packages. Section IV also describes the time flow of net social benefits and costs up to 80 years after enactment of a permanent package of reforms.

SECTION II. SUMMARY OF PACKAGES

As part of the three packages it has advanced, CPRAC developed 11 policy proposals, each of which is designed to improve the economic circumstances of and alleviate poverty among families with children (CPRAC, 2024). These policy proposals include expansions to existing policies, such as the Empire State Child Credit and New York State's Temporary Assistance programs, as well as the establishment of new state benefit programs.

Some of these policies are designed to strengthen preexisting benefits by either expanding eligibility or by increasing benefit amounts. For example, some proposals focus on New York State's Empire State Child Credit (ESCC), the state's version of the federal Child Tax Credit, expanding eligibility by including children with ITINs (or children without a Social Security number) or increasing the credit's maximum benefit amounts. Other policy proposals are targeted towards New York State's Temporary Assistance program, specifically Family Assistance (FA) and Safety Net Assistance (SNA). These proposals could increase eligibility by applying earned income disregards or removing asset tests, or increase the value of the program's basic allowances.

Two of CPRAC's 11 policy proposals include the establishment of new statewide policies—a statewide food benefit and a housing voucher program. While these policies are similar to preexisting federal programs, such as SNAP and the Section 8 Housing Choice Voucher Program, these New York State proposals would differ by extending eligibility to families who may not be eligible for the federal programs (if, for example, their income or citizenship status does not meet the federal eligibility criteria). These 11 proposals are described in Table 1 below.

Table 1. Description of CPRAC policy proposals

Proposal #	Proposal Description
TP 1	Increase Maximum Empire State Child Credit Amount to \$500 for Children 6+, \$1000 for Children 6-, Fully Refundable
TP 2	Increase Maximum Empire State Child Credit Amount to \$1500 for Children 0-17, Fully Refundable
TP 3	Increase Maximum Empire State Child Credit Amount \$1500 for Children 6+, \$2000 for Children 6-, Fully Refundable
TP 4	Make Children with ITINs Fully Eligible for Maximum Empire State Child Credit Amount
PBP 1	Increase the Public Assistance Basic Allowances for Family Assistance and Safety Net Assistance by 50%
PBP 2	Increase the Public Assistance Basic Allowances for Family Assistance and Safety Net Assistance by 100%
PBP 3	Apply the Same Earned Income Disregards for Public Assistance Applicants as for Recipients
PBP 4	Remove the Public Assistance Assets Test
PBP 5	Remove Public Assistance Durational Sanctions in Areas Using That Policy
PBP 6	Create a State Food Benefit for Families with Children, No Noncitizen Restrictions
HP 1	Create a State Housing Voucher for Unsubsidized Income-Eligible Households, No Noncitizen Restrictions

Source: Child Poverty Reduction Advisory Council, October 2024.

From these 11 policy proposals, CPRAC has designed three different policy “packages”, or combinations of the above proposals. These policy packages are detailed in Table 2 below. All of the packages include a proposal to establish a statewide food benefit program, to create a state housing voucher program, as well as proposals to expand eligibility for the ESCC, FA, and SNA programs. Each package also includes one of three options to raise the maximum credit amount of the ESCC (labeled CTC in CPRAC’s vernacular), with a range of new maximum values of \$500-\$2000, and one of two options to expand the basic allowance of FA and SNA by either 50% or 100%. Together, these policy packages are designed to strengthen support for families with children in an array of economic circumstances and to reduce child poverty.

Table 2. Description of CPRAC policy packages

Recommended Policy Package 1	
<i>Proposal #</i>	<i>Proposal Description</i>
TP 2	Increase Maximum Empire State Child Credit Amount to \$1500 for Children 0-17, Fully Refundable
TP 4	Make Children with ITINs Fully Eligible for Maximum Empire State Child Credit Amount
HP 1	Create a State Housing Voucher for Unsubsidized Income-Eligible Households, No Noncitizen Restrictions
PBP 2	Increase the Public Assistance Basic Allowances for Family Assistance and Safety Net Assistance by 100%
PBP 3	Apply the Same Earned Income Disregards for Public Assistance Applicants as for Recipients
PBP 4	Remove the Public Assistance Assets Test
PBP 5	Remove Public Assistance Durational Sanctions in Areas Using That Policy
PBP 6	Create a State Food Benefit for Families with Children, No Noncitizen Restrictions
Recommended Policy Package 2	
<i>Proposal #</i>	<i>Proposal Description</i>
TP 1	Increase Maximum Empire State Child Credit Amount to \$500 for Children 6+, \$1000 for Children 6-, Fully Refundable
TP 4	Make Children with ITINs Fully Eligible for Maximum Empire State Child Credit Amount
HP 1	Create a State Housing Voucher for Unsubsidized Income-Eligible Households, No Noncitizen Restrictions
PBP 2	Increase the Public Assistance Basic Allowances for Family Assistance and Safety Net Assistance by 100%
PBP 3	Apply the Same Earned Income Disregards for Public Assistance Applicants as for Recipients
PBP 4	Remove the Public Assistance Assets Test
PBP 5	Remove Public Assistance Durational Sanctions in Areas Using That Policy
PBP 6	Create a State Food Benefit for Families with Children, No Noncitizen Restrictions
Recommended Policy Package 3	
<i>Proposal #</i>	<i>Proposal Description</i>
TP 3	Increase Maximum Empire State Child Credit Amount \$1500 for Children 6+, \$2000 for Children 6-, Fully Refundable
TP 4	Make Children with ITINs Fully Eligible for Maximum Empire State Child Credit Amount
HP 1	Create a State Housing Voucher for Unsubsidized Income-Eligible Households, No Noncitizen Restrictions
PBP 1	Increase the Public Assistance Basic Allowances for Family Assistance and Safety Net Assistance by 50%
PBP 3	Apply the Same Earned Income Disregards for Public Assistance Applicants as for Recipients
PBP 4	Remove the Public Assistance Assets Test
PBP 5	Remove Public Assistance Durational Sanctions in Areas Using That Policy
PBP 6	Create a State Food Benefit for Families with Children, No Noncitizen Restrictions

Source: *Child Poverty Reduction Advisory Council, October 2024.*

SECTION III. METHODOLOGY FOR ESTIMATING THE BENEFITS AND COSTS OF PERMANENT PACKAGES FOR THE FIRST YEAR OF PACKAGE IMPLEMENTATION

The transfer (cost) estimates used in this analysis are derived from Urban Institute tabulations of data from the ATTIS model, using 2019 as the base year. The policy packages will increase household income and generate short- and long-term impacts on children and adults. Estimates on household income are also derived from the microsimulation conducted by the Urban Institute.

Table 3 below presents the Urban Institute estimates of total transfers to recipients and child poverty reductions of the three packages.¹ Two features of the table stand out. First, only Package 1 reduces child poverty by 50%. Second, the poverty reductions are consistent with the costs of the packages. Package 3 costs a little less and reduces poverty a little less than Package 1. Package 2 is the least costly and reduces poverty the least. Not presented here are the numbers of children/adults that experience changes in family resources from the package. Each package affects roughly 3 million children (ages 0-17) and adults across 2 million families (see Tables A3 and A4 in [Appendix A](#)).

Table 3. Total costs of the income transfers for each policy package

	Total transfers of the new policies (\$2019 billions)	Percent child poverty reduction associated with each package (ages 0-17)
Package 1	\$8.944	-50.5%
Package 2	\$6.645	-41.0%
Package 3	\$8.477	-46.7%

Source: Urban Institute tabulations of data from the ATTIS model using CPRAC-SPM, using 2019 American Community Survey Data.

Note: Total transfers are derived from Urban Institute estimates on the total change in family resources broken out by family income level. The transfer totals are about .3% lower than the total costs of the packages estimated by the Urban Institute. The slight difference is attributable to different weighting methods used by the Urban Institute for the calculation of total transfers versus total costs.

To estimate the impacts of an increase in household income, we use Garfinkel et al. (2022, 2024). These papers present analysis of the latest literature and select the most rigorous experimental and quasi-experimental studies that provided causal evidence of the impacts of increased household income from cash and near-cash transfers on children and parents in low-income families. The majority of these studies examined income transfers targeted at low-income families, such as SNAP, and provided estimates of their effects on children's birthweight, neonatal mortality, health in childhood and adulthood, longevity, earnings in adulthood, educational attainment, and involvement in the criminal legal system and the child welfare system. For parents, these studies estimate the effects of transfers on parental health and longevity. All studies but one found that cash and near-cash transfers generate positive impacts throughout children's and parents' lifetimes. Given that many of these studies are concerned with different income transfer programs of different magnitudes of benefits and different lengths of program participation, for each study, Garfinkel et al. standardized its results to reflect the present discounted value of lifelong benefits and costs for children and parents

¹ As noted in the methodology section, the estimates of total transfers are about 0.3% lower than the total costs.

per \$1,000 increase in household income per year from cash and near-cash transfers.² Detailed summaries and standardization of these studies are included in [Appendix B](#).

In addition to estimating impacts, the dollar value of the impacts must be determined. Earnings impacts are already valued in dollar terms. How much is improved health and a longer life worth in dollar terms? How much is a reduction in crime worth in dollar terms? We follow Garfinkel et al. (2022, 2024) which relies on a vast literature in economics that assesses the dollar value of health, crime, and other impacts.

We use the standardized benefits and costs of an increase in household income of \$1000 per year calculated by Garfinkel et al. (2022, 2024) together with the Urban Institute estimates of increases in household income described above to calculate the present discounted value of the benefits and costs of the transfers for the first year of operation of the three policy packages. However, there are several assumptions we need to make. First, we assume that the impacts of near cash and in-kind transfers are the same as the impacts of cash transfers. Theoretically, these impacts may differ, as people may value near-cash and in-kind transfers less than their dollar values. However, we believe such differences are minor compared to the benefits of the policy packages. Under this assumption, each policy package is simply treated as a collection of cash transfers that increases household income. Second, we assume that the impacts calculated by Garfinkel et al. on parents apply to adults in general (including adults without children, who receive transfers from the policy package). In fact, one of the studies selected by Garfinkel et al. examines overall impacts on adults (including those without children). Finally, in accordance with the latest federal guidance, we use a discount rate of 2% for the discounting of future benefits and costs (OMB 2023).

Table 4 presents standardized mean estimates of the present discounted value of these lifetime benefits and costs for a one-child, one-parent, low-income household per \$1,000 increase in annual household cash income. A few examples are given in the text below of the present discounted value of different benefits that accrue to such a household.

² Discounting is the process of estimating future gains or losses in today's terms. Because a dollar today is worth more than a dollar next year (a dollar today can be invested at the current interest rate and will be worth more than a dollar by next year), expenditures today represent a greater cost than the same level of expenditures 10, 20, or 30 years from now. Conversely, a benefit of a certain level received in the future has a smaller monetary value in the present.

Table 4. Present discounted value of lifetime monetary benefits and costs for a one-child, one-parent, low-income household per \$1,000 increase in annual household cash income for a single year: Using mean estimates from quasi-experimental studies (\$2019)

	Direct + Participants	Indirect = Taxpayers	Total Society
Total transfer	\$ 1,000	\$ -1,000	\$ 0
Increased future earnings of children	\$ 1,940	\$ 0	\$ 1,940
Increased future tax payments by children	\$ -407	\$ 407	\$ 0
Increased children's health and longevity	\$ 4,892	\$ 0	\$ 4,892
Increased parents' and other adults' health and longevity	\$ 549	\$ 0	\$ 549
Avoided expenditures on other cash or near-cash transfers	\$ -26	\$ 26	\$ 0
Avoided expenditures on child protection	\$ 0	\$ 37	\$ 37
Avoided criminal justice expenditures	\$ 0	\$ 372	\$ 372
Reduced victim costs of crime	\$ 0	\$ 1,060	\$ 1,060
Increased costs of children's education	\$ -329	\$ -79	\$ -408
Avoided expenditures on children's health care costs	\$ 20	\$ 162	\$ 183
Avoided expenditures on parents' and other adults' health care costs	\$ 0.36	\$ 2.89	\$ 3.24
Increased payment due to increased children's longevity	\$ 450	\$ -450	\$ 0
Increased payment due to increased parents' and other adults' longevity	\$ 114	\$ -114	\$ 0
Decreased tax payments from parents and other adults^a	\$?	\$?	\$ 0
Administrative costs^b	\$ 0	\$ -70	\$ -70
Tax distortion costs for taxpayers (inefficiencies due to higher taxation)	\$ 0	\$ -304	\$ -304
Total^c	\$ 8,202	\$ 51	\$ 8,253

Source: Center on Poverty and Social Policy at Columbia University.

^a Decreased tax payments from parents and other adults are not yet estimated in this analysis as labor supply effects have not yet been provided. While decreased tax payments are typically small, they will be incorporated when these labor supply effects have been estimated.

^b Our estimate of administrative costs is based on the report of the Child Poverty Reduction Advisory Council (CPRAC, 2024). In the CPRAC report, it is estimated that the housing voucher proposal (HP1), which is included in all three policy packages analyzed in this report, could incur \$480 million to \$546 million of administrative costs. The public benefit proposals included in each of the three packages could incur a total of \$6.1 million of administrative costs. The tax policy proposals impose minimum or no administrative costs. We thus calculate the administrative cost of the three packages as \$546 million + 6.1 million = \$552.1 million, or \$0.6 billion if rounded. \$552.1 million of administrative cost are on average 7% of the cost of policy packages 1-3 (\$8.9, \$6.6, \$8.5 billion respectively). We thus calculate the administrative cost per \$1,000 to be \$70 (7% * \$1000=\$70).

^c Note that, due to rounding procedures, totals might not be an exact sum of their parts.

Table 4 shows that for every \$1000 increase in annual household cash income the long-term benefits to child beneficiaries are substantial, valued at \$8,202.³ The biggest benefits are in children's health and longevity, valued at \$4,892, representing nearly five times the initial increase in household cash income. These large values are based upon both the quasi-experimental research which indicates that increases in household income have a big impact on health and longevity, and the high value that human beings place on being healthy and living longer. The most common method for estimating the value of good health and living longer used today among researchers is to compare the wage rates of comparable people in high-risk versus low-risk jobs. For example, wage rates for miners (a dangerous occupation in terms of both health and longevity) are higher than wage rates for clerks of the same age, race,

³ These figures are drawn from Garfinkel et al 2024, which update the calculations in Garfinkel et al 2022. The biggest difference is the use of a discount rate of 2% rather than 3%. The choice of a 2% discount rate aligns with the latest recommendation made by The Office of Information and Regulatory Affairs (OIRA) on regulatory analysis (The White House, 2024).

and educational attainment in order to compensate the miner for the extra risk of mining. (For a longer discussion of how we combine the impacts in the quasi-experimental literature, with the valuation of health and life literature, see [Appendix B](#) of the technical report). In addition to health and longevity, the second largest benefit to recipient children is the increase of \$1,940 in their future earnings. We also note that there are also long-term benefits in health and longevity for adults, estimated at \$549.

Taxpayers also experience benefits in a few domains from this income transfer. The biggest taxpayer benefit, \$1,432 lifetime benefits per \$1,000 increase in income, comes from reductions in criminal justice expenditures and the victim costs of crime. These reductions are the result of fewer crimes being committed: Quasi-experimental research indicates that cash transfers to children, especially children in poverty, reduce crime, including both property crimes and violent crimes. We use the result of quasi-experimental research on crime reduction per year associated with cash transfers as well as studies that examine the age distribution of crime to calculate estimated reduction in crimes throughout children's lifetimes. We monetize the lifetime decrease in crime using standard literature estimates on the cost of crime. Twenty-six percent of this reduction is attributable to reductions in taxes and expenditures associated with police, courts, and jails and prisons, and 74 percent is attributable to reductions in victim costs. In the literature, victim costs are often estimated by how much people are willing to pay to avoid both property and violent crimes.⁴

The next largest taxpayer benefits are increased future earnings of children, which lead to increases in taxes they pay (\$407) and decreases in other transfers those children receive (\$26). Health care costs decrease by \$20 for beneficiaries and by \$162 for taxpayers. Child welfare spending declines as well, saving taxpayers \$26. Taxpayers also see increases in certain costs. Increased schooling of children imposes a cost on taxpayers of \$79. Increased longevity of both the child and parent increases Social Security and Medicare transfers that nearly offset the increase in taxes paid from increased earnings.

Valuation of different benefits for an impacted household are discussed in greater detail in [Appendix B](#). In Section V (and [Appendix C](#)), we discuss in more detail how we project each of the benefits and costs over time.

We conclude this section by noting that the lifetime benefits that accrue to program participants will underestimate the total long-term benefits to society if the first generation children's own children (i.e., the grandchildren of the parents receiving the transfers) that experience increases in health, education, and earnings themselves experience increases in health, education, and earnings as a consequence of the benefits to the first generation. We do not attempt to estimate the magnitude of these intergenerational benefits.

⁴ There are multiple estimates in the literature on the valuation of the victim costs of crime. Garfinkel et al., (2024) rely most heavily on Cohen (2020) and Miller et al., (2021).

SECTION IV. RESULTS ON LIFETIME SOCIAL BENEFITS AND COSTS GENERATED BY THE FIRST YEAR OF PACKAGE IMPLEMENTATION

To translate the standardized benefits and costs per \$1000 of program expenditure for a single parent and a single child from Table 4 into estimates of the social benefits generated by actual expenditures of particular programs and program packages, we need to take account of actual expenditures and the actual number of New York families and children that receive transfers from the program packages. As described in [Appendix A](#), we use estimates provided by the Urban Institute of the total transfers (approximately equal to the initial fiscal costs) of each program package and for estimates of the actual number of families and children being aided by each program package.

We also need to take account of the fact that the estimates from Garfinkel et al. (2022, 2024) apply to low-income families, whereas many of the families aided by the package are middle-income and upper-middle-income. Intuition and empirical evidence suggest that the social benefits generated by increases in household income will be lower the higher the income of the family. An increase of household income of \$2000 constitutes a 20% increase for those with incomes of only \$10,000, but an increase of only 8% for a family with \$50,000, and an increase of only 2% for a family with \$100,000. Though considerable correlational work in the US and elsewhere suggests income is more important for the poorest families, causal evidence on this point is surprisingly lacking; but based on causal analysis from Norway (Løken, Mogstad, and Wiswall, 2012), we assume that families with incomes under \$50,000 derive the full benefits, which then decline smoothly between \$50,000 and \$100,000, and that benefits to families above \$100,000 equal zero.. Thus, we use estimates from the Urban Institute of increases in household income for families with incomes below \$50,000, between \$50,000 and \$100,000, and above \$100,000. The full details of these increases in household income are described in [Appendix A](#).

Finally, because there is virtually no research that estimates the benefits to a second parent of transfers to low income families, we follow Garfinkel et al (2022) and assume that only one parent/adult in a family derives social benefits from the increase in family income from the program packages.

Whereas Table 4 presents the benefits and costs generated per \$1000 of program expenditures for a single parent and a single child Table 5 describes the present discounted value of the direct, indirect, and total net social benefits of each of the benefits generated by the actual program expenditures of Package 1 and the actual number of families and children in New York State who would receive the benefits of Package 1. The resulting figures in Table 5 are based on a single year of positive resource changes of \$8.9 billion for 2.1 million families, including 3.2 million children. A single year of operation of Package 1 has initial fiscal costs of \$8.9 billion and generates net social benefits of \$94.1 billion over the lives of all program participants. We assume the Package is enacted on a permanent basis. Each year would cost an additional \$8.9 billion and generate an additional \$94.1 billion over the lives of program participants.

Table 5. Present discounted value of lifetime monetary benefits and costs for Policy Package 1 generated by the first year of package implementation (\$2019 billions)

	Direct + Participants	Indirect = Taxpayers	Total Society
Total transfer	\$ 8.9	\$ -8.9	\$ 0
Increased future earnings of children	\$ 22.3	\$ 0	\$ 22.3
Increased future tax payments by children	\$ -4.7	\$ 4.7	\$ 0
Increased children's health and longevity	\$ 56.1	\$ 0	\$ 56.1
Increased parents' and other adults' health and longevity	\$ 4.0	\$ 0	\$ 4.0
Avoided expenditures on other cash or near-cash transfers	\$ -0.3	\$ 0.3	\$ 0
Avoided expenditures on child protection	\$ 0	\$ 0.4	\$ 0.4
Avoided criminal justice expenditures	\$ 0	\$ 4.3	\$ 4.3
Reduced victim costs of crime	\$ 0	\$ 12.1	\$ 12.1
Increased costs of children's education	\$ -3.8	\$ -0.9	\$ -4.7
Avoided expenditures on children's health care costs	\$ 0.2	\$ 1.9	\$ 2.1
Avoided expenditures on parents' and other adults' health care costs	\$0.003	\$ 0.02	\$0.023
Increased payment due to increased children's longevity	\$ 5.2	\$ -5.2	\$ 0
Increased payment due to increased parents' and other adults' longevity	\$ 0.8	\$ -0.8	\$ 0
Decreased tax payments from parents and other adults^a	\$?	\$?	\$ 0
Administrative costs^b	\$ 0	\$ -0.6	\$ -0.6
Tax distortion costs for taxpayers (inefficiencies due to higher taxation)	\$ 0	\$ -2.2	\$ -2.2
Total^c	\$ 88.9	\$ 5.1	\$ 94.0

Source: Center on Poverty and Social Policy at Columbia University.

^a Decreased tax payments from parents and other adults are not yet estimated in this analysis as labor supply effects have not yet been provided. While decreased tax payments are typically small, they will be incorporated when these labor supply effects have been estimated.

^b Our estimate of administrative costs is based on the report of the Child Poverty Reduction Advisory Council (CPRAC, 2024). In the CPRAC report, it is estimated that the housing voucher proposal (HP1), which is included in all three policy packages analyzed in this report, could incur a maximum of \$546 million of administrative costs. The public benefit proposals included in each of the three packages could incur a total of \$6.1 million of administrative costs. The tax policy proposals impose minimum or no administrative costs. We thus calculate the administrative cost of the three packages as \$546 million + 6.1 million = \$552.1 million, or \$0.6 billion if rounded.

^c Note that, due to rounding procedures, totals might not be an exact sum of their parts.

Some of the details in Table 5 are worthy of comment:

- The total transfer to program participants is \$8.9 billion. This benefit is offset by the equal fiscal cost to taxpayers, resulting in a net social benefit or cost of zero. The transfers, as noted above in the methods section, are not costless, however, because of administrative costs and the costs associated with distortion stemming from higher taxes required to finance the transfer.
- The largest benefits to program participants are increased future earnings of the children and improved health of the children—respectively \$22.3 billion and \$56.1 billion.
- Among all the benefits listed, the benefits to children are much larger than the benefits to adults, indicating that packages more heavily targeted to children will have larger benefits than packages that are less targeted to children.
- The largest benefits to taxpayers are the increase in taxes paid by child program participants who earn more and the reductions in both criminal justice costs and victim costs that result from less crime committed by children when they become teenagers and adults.
- The largest costs to taxpayers are the initial transfer of \$8.9 billion and the much later transfer to children and adults who live longer, including for example on government programs like Social Security, Medicare, and Medicaid that individuals might utilize more in their older years as they live longer lives.
- All of the relationships described above, such as much larger benefits for children than for adults, are common to all of the packages. Only the dollar amounts differ.

Table 6 presents the total net social benefits for all three packages. The first column also reproduces the initial fiscal costs of each package. The second and third columns break down the net social benefits to society as a whole by direct benefits to program participants and indirect benefits to taxpayers. The net social benefits and costs to society as a whole are given by the last column. Benefits and costs to society as a whole are the focus of benefit-cost analysis and of this report. The first package has the largest net social benefit: \$94.0 billion over the lives of program participants. Package 3 has the second-largest net social benefit, at \$86.3 billion. Package 2 generates a net social benefit of \$65.3 billion. In short, like poverty reduction, the net social benefits are directly related to the costs of the packages.

Table 6. Present discounted value of lifetime monetary benefits and costs for three policy packages generated by the first year of package implementation (\$2019 billions)

	Initial Fiscal Costs	Direct + Participants	Indirect = Taxpayers	Total Net Social Benefits
Package 1	\$ 8.9	\$ 88.9	\$ 5.1	\$ 94.0
Package 2	\$ 6.6	\$ 62.8	\$ 2.6	\$ 65.3
Package 3	\$ 8.5	\$ 82.0	\$ 4.3	\$ 86.3

Source: Center on Poverty and Social Policy at Columbia University.

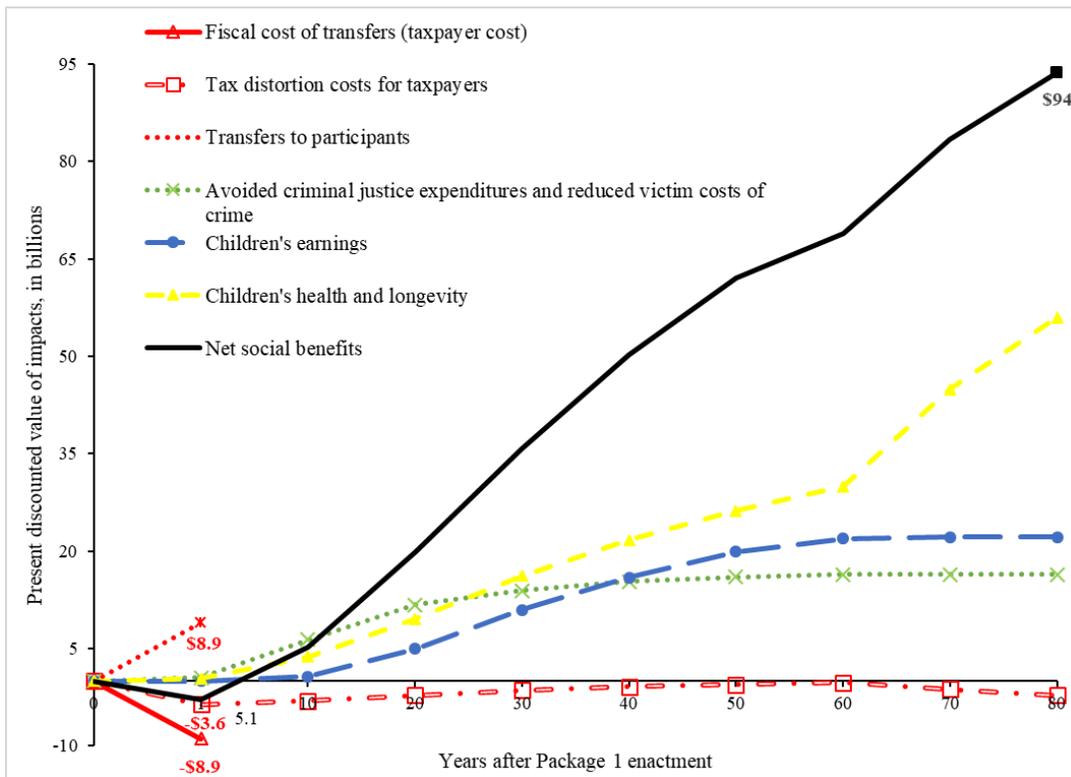
We also provide a rough breakdown of taxpayer benefits and costs that accrue to New York State residents compared to taxpayer benefits and costs that accrue to the federal government or to other states. This rough approach was required due to limitations in readily available data. For example, there is no data available about the number of families with children born in New York who move out of state and the ages of the children and parents at the time of that move, which would help determine with greater precision the proportion of taxpayer benefits and costs collected by New York state versus other states and the federal government, specifically when calculating the victim costs of crime, the biggest taxpayer benefit. We describe the sources of data we used and the calculations and results in [Appendix E](#) of the technical report. In summary, when accounting for all benefits, we estimate that for Policy Package 1, State taxpayers would experience a net gain of \$3.0 billion, and Federal taxpayers would experience a net gain of \$2.1 billion, for a total net taxpayer benefit of \$5.1 billion.

SECTION V. THE FLOW OF BENEFITS AND COSTS GENERATED BY EIGHTY YEARS OF PACKAGE IMPLEMENTATION

Up to this point we have focused on the present discounted value of a single year of Package implementation over the lifetimes of program beneficiaries. The initial fiscal costs of a program are incurred right away, while some benefits accrue much later. Figure 1 depicts this difference in the timing of lifetime benefits and costs by showing the flow of a few selected lifetime costs and benefits presented in Table 5 for the first year of Package 1 implementation. We begin with fiscal costs paid by the taxpayer. By the end of the first year the fiscal costs of the transfers reach \$8.9 billion. These costs are offset dollar for dollar by benefits to program participants of \$8.9 billion. The true social costs are the tax distortion costs, which equal 40% of the net changes in the present discounted values of taxes, or \$3.6 billion. While much smaller than fiscal costs, the tax distortion costs of Package 1 in the first year are still quite large. Note that these tax distortion costs decrease over time then increase again slightly as retirement and health benefits increase because of longer life of child recipients.

Figure 1 also depicts the three largest lifetime benefits (see Table D4 in [Appendix D](#) for details). Reduction in criminal justice expenditures and victim costs of crime to taxpayers is initially the largest benefit, but by 25 years after Package 1 implementation, lifetime health benefits to program participants exceed reductions in crime costs. By year 40, this is also true for lifetime earnings benefits. Note that all three lifetime benefits are much smaller in the first few years after Package 1 is first implemented than they eventually become. Finally, the figure depicts net lifetime social benefits for one year of Package 1 operation, which by year 80, reach \$94 billion, as presented in the last row of Table 5.

Figure 1. Selected lifetime benefits and costs for Policy Package 1 generated by the first year of package implementation (\$2019 billions)



Note that lifetime net social benefits are negative until 4.1 years after Package 1 is first implemented. In a permanent program, by the end of year 4, fiscal costs and therefore inefficiency costs of taxation will have been incurred for 4 years. To depict the flow of social benefits and costs of a permanent program, we shift to a new figure that incorporates the costs and benefits of multiple years of program operation.

In a permanent program, initial fiscal costs are repeated every year and generate both inefficiencies due to taxation and a new round of long run benefits. It is informative to describe that ongoing flow of costs and benefits over time. Doing so, for example, can help project how long it takes for social benefits from a continuing program to equal or exceed social costs: does it take 10 years, or 20 years? The Congressional Budget Office (CBO), for example, uses a ten-year window to examine the effects of policies on the federal budget, ignoring any costs or benefits of a program that occur more than 10 years in the future. In line with the practices of the CBO, we first examine the benefits and costs of the three packages for the first 10 years of program operation. Then we look at a more extended period of time, 80 years, which allows us to include effects on beneficiaries throughout their lives.

We assume that the number of children/adults benefitting from the program in each year and their real family incomes remain the same as in 2019 (the year of ACS data used for the analysis). The estimates therefore represent what the benefits and costs of the program package would be if the demographic composition and income distribution remained steady over 80 years. The oldest children and their parents in the first year of program operation will age out of the second year of operation, but are replaced by identical children and families. Similarly, newborns in the second year will be identical to newborns in the first year. Although incomes and family composition will undoubtedly change, how they will change is a separate research question beyond the scope of our analysis. The packages themselves will increase family incomes of children when they become adults which will reduce poverty and generate intergenerational effects on their children. As discussed above, we do not estimate these intergenerational effects. Because neither poverty reduction effects nor intergenerational benefits are incorporated into our analyses of the flow of benefits, they underestimate the long term benefits of the program packages.

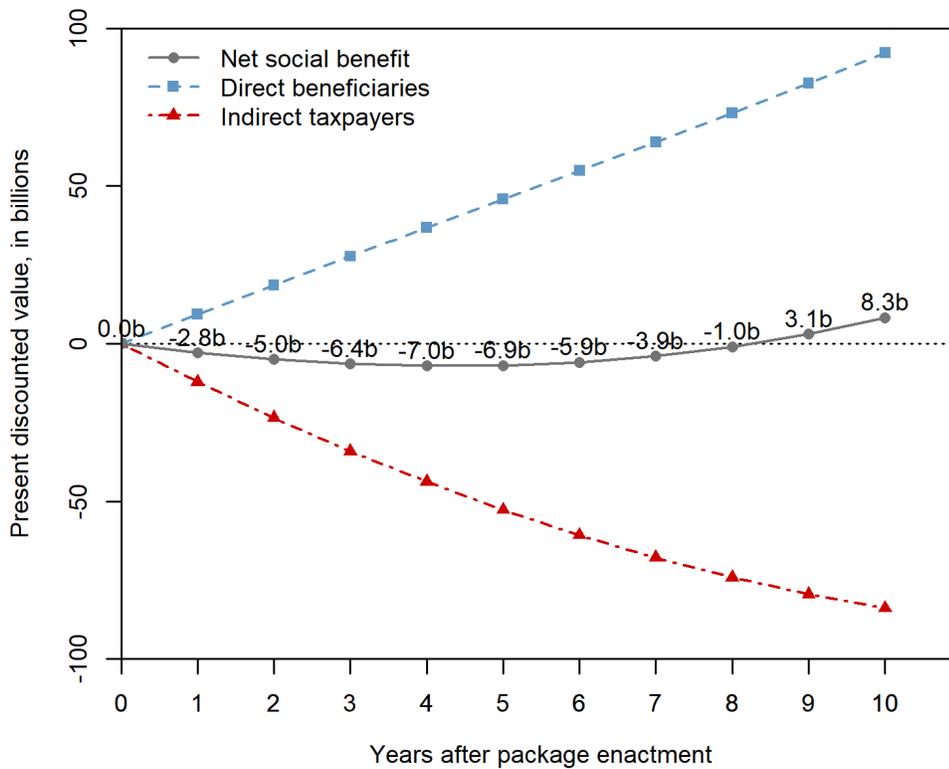
Each year of program implementation generates \$8.9 billion of fiscal costs. Fiscal costs (and benefits) that accrue in the future are discounted with a discount rate of 2 percent. The present discounted value of fiscal costs in year 2 is therefore $(0.98 \times \$8.9 \text{ billion})$, or \$8.72 billion. Adding up the present discounted value of the fiscal costs across years, we obtain the accumulated costs. In year 2, for example, the present discounted value of accumulated fiscal costs is \$8.9 billion plus \$8.72 billion. More important for social costs are the inefficiencies generated by taxation or 40% of these figures.

Each year of program implementation will also generate a set of benefits for beneficiaries, taxpayers, and society, as presented in Table 5. Benefits that accrue in the future are discounted. As described above, some benefits accrue relatively immediately for children and their parents—for example, increases in children's and parents' health. But even health benefits that accrue in the first year or two of program enactment will be much smaller than the present discounted value of lifetime health benefits. For example, assuming health benefits in years 1 and 2 are equal, cumulative health benefits will be nearly twice as large in year 2 as in year 1. But even twice the annual first year benefit is much lower than lifetime benefits. Other benefits take longer to realize—for example, increases in children's future earnings—and are thus more

heavily discounted. For newborn children, benefits from increased earnings will only begin to accrue 20 years after birth; even for 10-year-olds, on average, benefits to both program recipients and taxpayers take 10 years to begin to accrue. The biggest benefit to taxpayers is reduction in crime: while newborns do not commit crimes, in general they begin committing crimes at a younger age than they begin earning income, so these benefits are less heavily discounted. The cumulative benefits curves in Figures 2 through 5 below incorporate these differences in timing for each benefit. In Figures 2 through 5, we display the flow of cumulative benefits and costs over time for different periods of time since the enactment of the packages. As in Table 5, we display direct benefits to beneficiaries, indirect benefits to taxpayers, and net social benefits—the sum of direct and indirect benefits. In [Appendix C](#), we discuss in more detail how we project each of the benefits and costs over time. Figure 2 presents the flow of cumulative benefits and costs over time for the first ten years since the enactment of Package 1. The figures for Packages 2 and 3 are displayed in [Appendix D](#) of the technical report.

Figure 2 shows that the benefits to program participants are large from the start of the program, but the costs to taxpayers are even larger. Thus, net social benefits (benefits to participants and taxpayers minus costs to participants and taxpayers) are initially negative and remain so until year 8. After that, net social benefits increase rapidly and reach \$8.4 billion in the tenth year.⁵ (In packages 2 and 3, the pattern is the same, but net social benefits 10 years after enactment equal only \$2.2 billion and \$6.2 billion, respectively—see details in [Appendix D](#)).

Figure 2. The flow of cumulative benefits and costs over time in the first 10 years of Policy Package 1 implementation (\$2019)



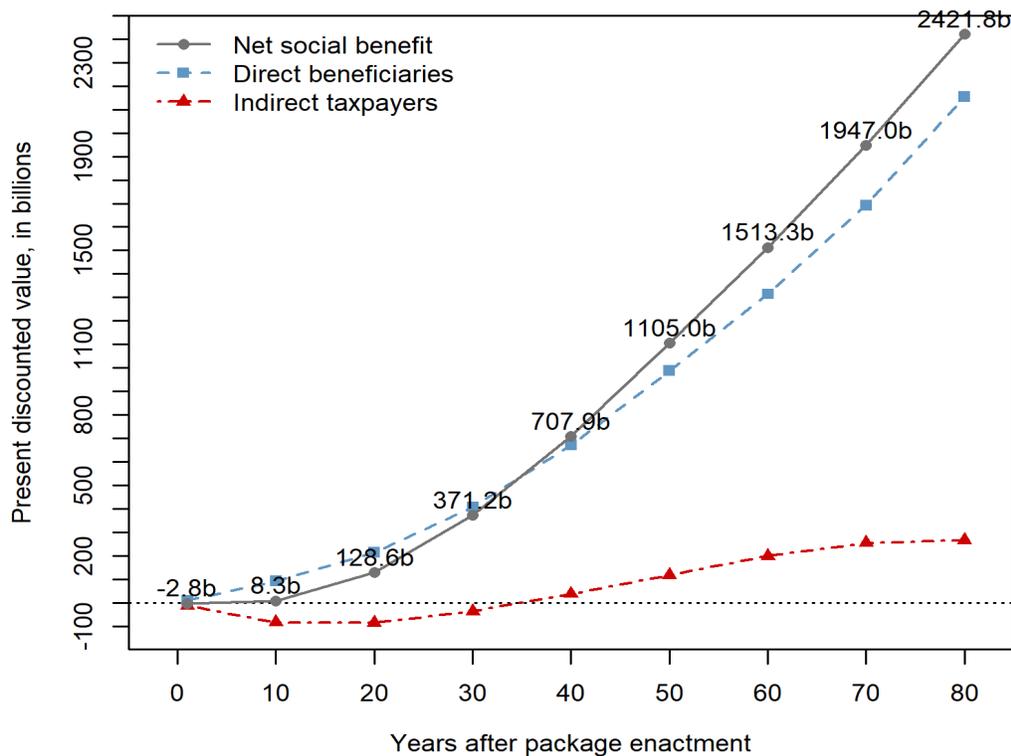
Source: Center on Poverty and Social Policy at Columbia University

⁵ The cumulative benefits relative to cumulative fiscal costs over 80 years will be smaller than the lifetime benefits and costs for one year of package implementation. This is because expenditures in yr 80 will generate only the first year of social benefits—the future benefits to those who receive the expenditures in yr 80 will accrue outside the time window within which we are analyzing. Similarly, expenditures in yr 79 will generate only 2 years of countable social benefits, yr 78 expenditures will generate only 3 years, and so on.

So, we have answered the question about how long it takes for cumulative social benefits to equal cumulative social costs—we estimate this to be about 8 years. But, we have not yet answered the question of how long it takes for the cumulative benefits to taxpayers to equal their cumulative social costs. Indeed, do benefits to taxpayers ever equal social costs? To do so, we need a longer window than 10 years.

Figures 3 through 5 display the flow of cumulative benefits and costs over time in the first 80 years of package implementation for the three packages. The first ten years of Figure 3 is the same as Figure 2. We first discuss Figure 3 in some detail and then briefly discuss the others.

Figure 3. The flow of cumulative benefits and costs over time in the first 80 years of Policy Package 1 implementation (\$2019)



Source: Center on Poverty and Social Policy at Columbia University.

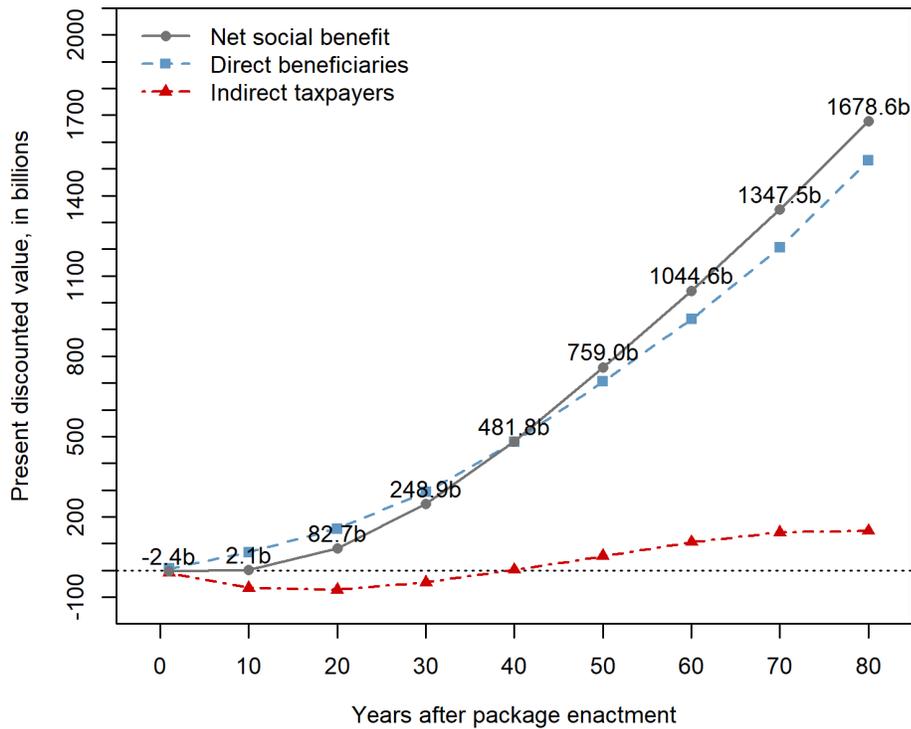
The most striking aspect of the figure is how different the picture looks when the analysis window is only ten years long as compared to when the window extends to 20 years and beyond. As shown in Figure 2, at year 10, benefits to society as a whole, i.e. the return on investment, is positive but relatively small, at only \$8.4 billion. By year 20, Figure 3 shows net social benefits increasing rapidly to about \$129 billion. Beyond year 20, benefits to program participants then continue to grow, while losses to taxpayers level off. By year 30, the social return grows dramatically: net social benefits exceed \$370 billion. Taxpayers as well as program beneficiaries experience net gains over the longer run, starting by year 32; the taxpayer gains are driven primarily by decreases in the victim costs of crime. Cumulative net social benefits, both direct and indirect, continue to grow steadily, such that, from 30 years to 80 years beyond initial program implementation, the net social benefits climb from \$370 billion to \$2,420 billion. Additional details for beneficiaries and taxpayers can be found in [Appendix D](#).

What are we to make of these stark differences in benefits and costs between windows of only 10 years as compared to windows of 20, 30, 50, and 80 years? The short answer is that in the case of investments in children, the 10-year window is too short to capture the benefits. If there was much greater uncertainty beyond 10 years, that would be a justification for putting less weight on the future years, emphasizing the 10-year rather than longer budget windows; indeed, the uncertainty of projections beyond 10 years has been used as justification for the CBO's focus on a 10-year window. But in this case, a great deal of the underlying empirical evidence of the impacts that we monetize is based on outcomes that have been measured well into adulthood and even throughout life, such as children's earnings in adulthood, health in adulthood, longevity, and crime. Even if there were greater uncertainty about the exact magnitude of these benefits beyond 10 years, treating these benefits as equal to zero is not justifiable. Rather, uncertainty about either the benefits or program duration should be handled explicitly, and would not in any case be expected to reduce net social benefits to zero (Boardman et al., 2018, chapter 11).

While the CBO's fiscal cost estimates focus on a standard ten-year window, this choice likely reflects political relevance rather than social. For an average person, a full lifetime of benefits and costs is arguably more relevant for policy implications. The CBO approach may offer tractable guidance in certain circumstances where the short-term payoff is important or there is great uncertainty over whether long-run impacts exist. However, the net social benefits by year illustrate how much society gains over varying ranges of plausible policy impact and longevity in ways that actually reflect social interests and not merely political interest in finances. And, after all, other government entities, including the trustees of the Social Security system and the Intergovernmental Panel on Climate Change, regularly make 70-plus year projections, due to the recognition that a 10-year window is not appropriate for many major societal concerns.

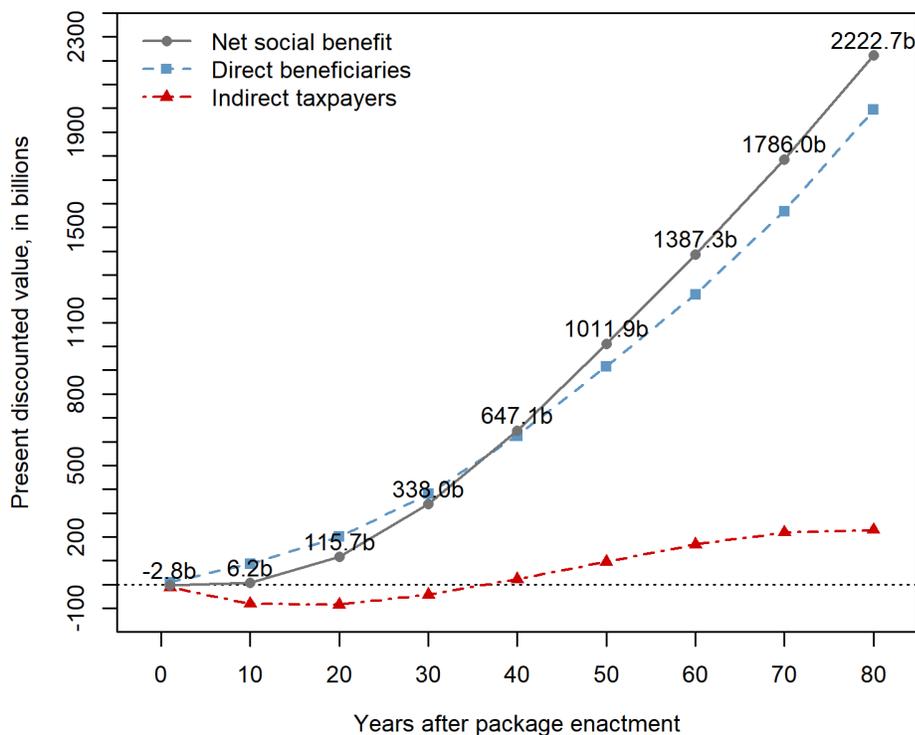
Figures 4 and 5 describe the time path of benefits and costs of Package 2 and Package 3, respectively. All the figures look, overall, alike. The benefits to society as a whole all become very large by 20 to 30 years after program enactment and continue to grow over time. As in Table 6, the benefits are smaller for Packages 2 and 3 as compared to Package 1. All the packages provide much lower returns during the first 10 years than after 20 or 30 years. Additional details can be found in [Appendix D](#).

Figure 4. The flow of cumulative benefits and costs over time in the first 80 years of Policy Package 2 implementation (\$2019)



Source: Center on Poverty and Social Policy at Columbia University.

Figure 5. The flow of cumulative benefits and costs over time in the first 80 years of Policy Package 3 implementation (\$2019)



Source: Center on Poverty and Social Policy at Columbia University.

ACKNOWLEDGEMENTS

This research was funded by the New York State Office of Temporary and Disability Assistance. We also acknowledge Robin Hood for their past and ongoing support of our Benefit-Cost Analysis work. The findings and conclusions presented in this report are those of the authors alone, and do not necessarily reflect the opinions of the funders or others acknowledged here. All errors are our own.

SUGGESTED CITATION

Garfinkel, Irwin, Elizabeth Ananat, Sophie Collyer, Robert Hartley, Anastasia Koutavas, Buyi Wang, and Christopher Wimer. 2025. [Investing in New York's future: The long-term benefits of child poverty reduction policies](#). New York: Center on Poverty and Social Policy, Columbia University.

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APPENDIX A. MICROSIMULATION OF THE THREE POLICY PACKAGES

The microsimulation is conducted by the Urban Institute using American Community Survey (ACS) and their Analysis of Transfers, Taxes, and Income Security (ATTIS) model. For more details on the microsimulation please read the Urban Institute's Methodology for this analysis. The first piece of information we need from the microsimulation is the total cost of net income transfers of the policy package. For example, this would include the new income transferred through the Empire State Child Credit or the State Food Benefit. The total cost of the package is equivalent to the total new resources transferred through the package. Table A1, produced by the Urban Institute, breaks out the net income transfers by three categories of baseline family resources: under \$50,000, \$50,000 to \$100,000, and above \$100,000, along with the total new resources transferred.⁶ Adjusting changes in resources by family income is motivated by the theory that the same increase in household income from transfers represents a smaller percentage increase for families with higher initial income, and may therefore generate smaller benefits. Though there is no empirical evidence in the U.S., Garfinkel et al., (2022) found a Norwegian study (Løken, Mogstad, and Wiswall, 2012) that provided causal evidence of this rate of decline. Garfinkel et al., (2022) concluded that families with initial income below \$50k enjoy the full benefits of increased household income, families with income between \$50k and \$100k enjoy partial benefits of increased household income, and families with income above \$100k enjoy no benefits from increased household income. We use this finding of Garfinkel et al., (2022). When calculating the benefits generated by the package, we count 100% of the resources received by families with initial resources below \$50k. We count less than 100% (56%) of the resources received by families with initial resources between \$50k and \$100k. We count 0% of the resources received by families with initial resources above \$100k.

Table A1. Total change in aggregate family SPM-defined resources, for each policy package, by baseline family (SPM-unit) pre-tax, pre-transfer, pre-expense resources (\$2019 billions)

	Under \$50,000 in pre-tax, pre-transfer, pre-expense resources	\$50,000 to \$100,000 in pre-tax, pre-transfer, pre-expense resources	More than \$100,000 in pre-tax, pre-transfer, pre-expense resources	Total (billions)
Package 1	\$6.484	\$1.531	\$0.930	\$8.944
Package 2	\$5.584	\$0.756	\$0.306	\$6.646
Package 3	\$5.917	\$1.539	\$1.021	\$8.477

Source: Urban Institute tabulations of data from the ATTIS model using CPRAC-SPM, using 2019 American Community Survey Data. Total change in aggregate resources rounded to the nearest million.

To decide what percentage of the full benefits we will assign to families whose initial resources fall into the \$50k-\$100k bin, we need to know the distribution of initial resources among these families.

⁶ Here, family is defined as the poverty unit under the Supplemental Poverty Measure (under this definition, childless family, or family composed of a single adult is also regarded as a family). Resources is defined as the SPM-unit level resources before accounting for tax credits and liabilities, cash and in-kind government transfers, and nondiscretionary expenses (medical, childcare, and work-related), and before accounting for resources transferred through the policy packages. That is, baseline SPM-unit level resources, pre-tax, pre-transfer, and pre-expenses. For presentation purposes, all numbers in the tables are rounded. For calculation, we use the unrounded numbers.

In Table A2 below, the Urban Institute shows the average baseline family resources in the three bins. According to the table, families in the \$50k-100k bin have an average of \$72,000 of baseline resources, which is 44% higher than the cutoff of \$50k. We assume that only 56% of the resources received by these families from the package would be generating benefits.

Table A2. Average (mean) baseline pre-tax, pre-transfer, pre-expense resource

	Pre-tax, pre-transfer, pre-expense resources			Total
	Under \$50,000	\$50,000 to \$100,000	More than \$100,000	
Number of families (in millions)	3.228	2.163	2.676	8.068
Average baseline pre-tax, pre-transfer, pre-expense resources	\$22,800	\$72,000	\$214,000	\$99,000

Source: Urban Institute tabulations of data from the ATTIS model using CPRAC-SPM, using 2019 American Community Survey Data. Average baseline pre-tax, pre-transfer, pre-expense resources rounded to the nearest thousand.

In addition to the adjustment of benefits with respect to family income, we also need to adjust the benefits by the number of children and parents in each family. Tables A3 and A4 present the total number of families and children that have experienced a change in family resources as a result of the policy package, calculated by the Urban Institute. A family with more children receives more transfers from the package and thus enjoys higher child-related benefits. The same logic applies to parents' benefits.

Table A3. Total number of families that experience changes in family resources from the package, by baseline family resources (\$2019 millions)

	Under \$50,000 in pre-tax, pre-transfer, pre-expense resources	\$50,000 to \$100,000 in pre-tax, pre-transfer, pre-expense resources	More than \$100,000 in pre-tax, pre-transfer, pre-expense resources	Total (millions)
Package 1	1.081	0.554	0.468	2.104
Package 2	1.081	0.553	0.390	2.024
Package 3	1.065	0.551	0.487	2.103

Source: Urban Institute tabulations of data from the ATTIS model using CPRAC-SPM, using 2019 American Community Survey Data. Number of families rounded to the nearest thousand.

Table A4. Total number of children (< 18) that experience changes in family resources from the package, by baseline family resources (\$2019 millions)

	Under \$50,000 in pre-tax, pre-transfer, pre-expense resources	\$50,000 to \$100,000 in pre-tax, pre-transfer, pre-expense resources	More than \$100,000 in pre-tax, pre-transfer, pre-expense resources	Total (millions)
Package 1	1.197	1.018	0.950	3.165
Package 2	1.197	1.016	0.794	3.006
Package 3	1.195	1.018	0.985	3.197

Source: Urban Institute tabulations of data from the ATTIS model using CPRAC-SPM, using 2019 American Community Survey Data. Total number of children rounded to the nearest thousand.

Table A5 sums the changes in family resources from the policy package across all children that experience changes in family resources from the package. As an example, suppose there is a family with 3 children, and that family's resources increased by \$300 under Package 1. This is a \$300 increase in the family's resources, but the total increase summed across all children in this family is \$900 because there were 3 children and each of them sees a \$300 increase in their family income.

Table A5. Total change in family resources from the policy package transfers at the child level, by baseline family resources (\$2019 billions)

	Under \$50,000 in pre-tax, pre-transfer, pre-expense resources	\$50,000 to \$100,000 in pre-tax, pre-transfer, pre-expense resources	More than \$100,000 in pre-tax, pre-transfer, pre-expense resources	Total (billions)
Package 1	\$9.432	\$3.651	\$2.311	\$15.393
Package 2	\$7.043	\$1.657	\$0.685	\$9.384
Package 3	\$8.483	\$3.712	\$2.564	\$14.760

Source: Urban Institute tabulations of data from the ATTIS model using CPRAC-SPM, using 2019 American Community Survey Data. Total change in family resources rounded to the nearest million.

With estimates of the total increase in household income brought by the policy package, we then estimate the impact of an increase in household income from the literature, discussed in the section below.

APPENDIX B. LITERATURE ON THE IMPACT OF CASH AND NEAR-CASH TRANSFERS

Literature Search Methodology and Selection Criteria

In order to ensure we have compiled a comprehensive literature on the benefits and costs of a child allowance, we followed a meta-analysis approach to gather evidence. We used a three-stage screening process to identify relevant studies for each benefit and cost. This methodology is described in detail in Garfinkel et al., 2022.

Literature on Impacts of Income

In this appendix, we present the detailed summaries and standardization of studies from Garfinkel et al., (2022, 2024), which we use to calculate the benefits and costs of the three policy packages. These studies examined the causal effects of cash and near-cash transfers on children's outcomes (future earnings, health in childhood and adulthood, longevity, educational attainment, and involvement in the child welfare system and criminal legal system) and parents' outcomes (health and longevity). Based on the increase in future earnings of children, Garfinkel et al., also calculated the increase in future tax payments of children and decrease in other transfers received by children. Based on the increase in health and longevity of children and adults, Garfinkel et al., calculated the resulting decrease in health expenditures and increase in longevity payments. Garfinkel et al., also estimated the increase in education cost from increased education, and reduction in expenditures on the child welfare and criminal legal system from less involvement in these systems. Garfinkel et al., standardized the findings across studies so they reflected the benefits and costs per \$1,000 increase in household income from cash and near-cash transfers per year. To calculate lifelong benefits and costs, Garfinkel et al., assigned the per-year benefits and costs throughout children's and parents' lives and calculated the present discounted value using a discount rate of 2%. If there were multiple studies for one outcome, Garfinkel et al., calculated the average. Table B1 presents the studies selected for the calculation and their standardized results. Only the impacts in Column A are used for the estimation of benefits and costs. Those listed in panel B and labeled supplementary studies are not used as that would involve double counting. The primary benefit of education is higher earnings, which are already counted. Similarly, the primary benefits of low birth weight are better health and eventually higher earnings, which are already counted. Improvements in mental health are encompassed by improvements in health

Table B1. Estimated impacts for low income families of a \$1,000 increase in household income as a result of a cash or near-cash transfer

Panel A: Impact studies used for the calculation of benefits	Panel B: Supplementary Impact studies
Children's earnings^a	Birthweight
Bailey et al. (2020)	Hoynes et al. (2015)
Bastian and Michelmore (2018)	Kehrer and Wolin (1979)
Aizer et al. (2016)	Almond et al. (2011)
Hoynes et al. (2016)	Markowitz et al. (2017)
Price and Song (2018)	Children's educational attainment
Children's health during childhood^b	Thompson (2019)
Averett and Wang (2018)	Bastian and Michelmore (2018)
Children's health during adulthood^b	Maxfield (2015)
Bailey et al. (2020)	Akee et al. (2010)
Hoynes et al. (2016)	Michelmore (2013)
Price and Song (2018)	Aizer et al. (2016)
Braga et al., (2020)	Child receiving high school diploma
Song (2019)	Thompson (2019)
Neonatal mortality	Akee et al. (2010)
Almond et al. (2011)	Bastian and Michelmore (2018)
Child longevity	Michelmore (2013)
Bailey et al. (2020)	Maxfield (2015)
Aizer et al. (2016)	Parents' and other adults' mental health
Crime	Gangopadhyaya et al. (2020)
Bailey et al. (2020)	Boyd-Swan et al. (2016)
Barr & Smith (2023)	
Child protection	
Berger et al. (2017)	
Rittenhouse (2022)	
Parents' and other adults' health^b	
Larrimore (2011)	
Morgan et al. (2020)	
Evans and Garthwaite (2014)	
Price and Song (2018)	
Parents' and other adults' longevity^b	
Price and Song (2018)	
Aizer et al. (2020)	
Chetty et al. (2016)	

Notes: *Results were statistically significant at the 5% level or lower.

+Includes both statistically significant and non-significant results for two or more measures of the same outcome.

^a All results are reported in 2019 dollars.

^b All results in the children's health section, parents' health and longevity section are calculated and expressed as a percentage of the full QALY value of \$126,628, as described later in the children's health section.

Children's Future Earnings

Aizer et al. (2016)

Aizer et al. (2016) found that in adulthood, sons whose mothers had received Mothers' Pensions between 1911 and 1935 experienced an increase in annual income of \$90.93 (s.e. 35.976), a 14% increase. As discussed in sections on children's longevity and children's educational attainment below, the authors also found an increase in longevity of 0.0158 (s.e. 0.007) or 1.16 years and an increase in educational attainment of 0.316 (s.e. 0.262) years. The authors matched administrative records, census records, and death records from 11 states to examine the long-term outcomes of male children who were raised in households who applied for the Mothers' Pensions between 1911 and 1935 (n=1,960). The authors compared the outcomes of children of accepted and rejected applicants using linear regressions. Rejected applicants were deemed to be an appropriate comparison group because like the accepted mothers, the rejected mothers were also economically constrained and sought aid, but they were somewhat better off (which is why they were rejected). So, in the absence of aid, their sons would have been expected to do somewhat better than the accepted sons, which implies that these estimates may somewhat understate the impact.

According to Aizer et al. (2016), Mothers' Pensions were \$3,684 (2019\$) annually and received for three years on average. A \$1,000 transfer for one year would thus increase children's future earnings by 1.27%⁷ ($0.14 \times ((1000/3684)/3)$). We believe that the level of future earning of children whose mothers received Mothers' Pensions during Aizer et al's study period approximates the 25th percentile income in 2019. According to the Current Population Survey, in 2019, annual earnings were on average \$10,000 at the 25th percentile of the working-aged⁸ earnings distribution (authors' calculations). Multiplying \$10,000 by 1.27% yields an annual increase in earnings of \$127. We calculate the present discounted value using equation 1 below. We assume a discount rate of $i=0.02$. According to our calculation above, the early benefit $B=\$127$. The average child beneficiary is assumed to be age 9. We use this assumption in the calculation of all child benefits. Increased earnings are assumed to begin at age 22 ($a=22$) and end at age 64 ($A=64$). We use this assumption for all estimates on children's future earnings. We conclude that the present discounted value of increased earnings in adulthood is \$2,863 as a result of a \$1,000 cash transfer during childhood.

$$PDV = \sum_{t=a}^A \frac{B}{(1+i)^{t-9}} = B \left(\frac{(1+i)^{9-(a-1)} - (1+i)^{9-A}}{i} \right) \quad (1)$$

Hoynes et al. (2016)

Hoynes et al. (2016) examine the long-term health and economic impact of exposure to food stamps between conception and age 5 using the Panel Study of Income Dynamics (PSID). They found that among individuals whose parents were without a high school diploma, exposure to food stamps from conception to age 5 increased earnings by \$3,610 (s.e. 5,064). \$3,610 (measured in 1995 dollars) is the equivalent of \$6,063 in 2019 dollars⁹. As discussed in the

⁷ Note that 1.27% is a rounded number. Even though in calculations we use the unrounded number, in the text we present the rounded number.

⁸ By working-aged, we refer to ages 25 to 64.

⁹ The paper starts measuring economic outcomes such as earnings in adulthood when individuals reach age 25. Since the sample includes individuals born between 1956-1981, this means that earnings in adulthood is first measured in 1981. The last wave of PSID data used by the paper is 2009. Thus, we assume that \$3610 is measured in 1995 dollars (the middle of the period 1981-2009).

section on children's health, they also found that among the full sample, exposure to food stamps from conception to age 5 decreased the probability of having metabolic syndrome by 0.438 (s.e. 0.204) standard deviations and increased the probability of reporting good health by 0.292 (s.e. 0.133) or 30 percentage points. The authors conducted difference-in-differences analyses taking advantage of variation in the introduction of the Food Stamp Program by county. The intent-to-treat group includes individuals whose parents were without a high school diploma and who did not receive food stamps as well as those whose parents were without a high school diploma but did receive food stamps. Models controlled for county, year of birth fixed-effects, year of interview, whether child was born to a female-headed household, education of head of household, family income, the child's gender, child's marital status, child's race, quadratic in age of child, state linear time trends, and 1960 county characteristics.

Hoynes et al. (2016) estimate that among families where heads had less than a high school degree, 43 percent participated in food stamps. Thus, in order to adjust results to reflect the impact on treated individuals we divide their results by 0.43, resulting in an estimate of \$14,100 ($\$6063/0.43$). Since individuals in the sample were exposed to food stamps for 7 years (from conception (age -1) to age 5), the estimate decreases to \$2,014. Average annual food stamps values per person in 1972 (near the midpoint of the study period) were \$994 per year in 2019 dollars (Department of Agriculture, 2021). Assuming average households have three individuals, the total household food stamps value would be \$2,982 on average. Thus, the impact decreases to \$675 ($\$2014 \times (1000/2982)$). As the paper studies the impact of exposure from conception (age -1) to age 5, we (conservatively) assume that individuals were exposed to food stamps through the entirety of childhood (from age -1 to age 17) but only derived benefits for future earnings during the first 7 years of payments. To measure the impact per year of payments, we multiply results by the 7/19 of years in which they derive benefits, yielding an estimate of \$249. Using equation 1, we conclude that the present discounted value of increased earnings in adulthood is \$5,624 as a result of a \$1,000 cash transfer during childhood.

Bailey et al. (2020)

Bailey et al. (2020) found that exposure to food stamps from conception to age 5 increased future earnings by 0.0114 (s.e. 0.0034) or 1.14 percent. The authors find no additional effects for exposure at ages 6-18. As discussed in the sections on children's health, children's longevity, and on crime reduction, they also discovered that as a result of exposure to food stamps, children's physical ability and health increased by 0.0013 standard deviations (s.e. 0.0013), children's longevity increased by 0.176 years (s.e. 0.030), children's future earnings increased by 1.14 percent (s.e. 0.34 percent), adult economic self-sufficiency increased by 0.0043 standard deviations (s.e. 0.0016), and the probability of being incarcerated decreased by 0.0008 (s.e. 0.0004) or 0.08 percentage points. Based on data from the 2001-2013 American Community Survey matched with the 2000 Census Long Form (n=7,705,000), the authors use a difference-in-difference framework exploiting the county-by-county introduction of food stamps. Models control for county of birth, birth year, and birth state fixed effects as well as 1960 county-level characteristics interacted with a linear birth-cohort trend.

Since children in the sample were exposed to food stamps for 7 years (conception to age 5), we divide 1.14 percent by 7, arriving at 0.16 percent. Average annual food stamps values per person in 1972 (near the midpoint of the study period) were \$994 per year in 2019 dollars (Department of Agriculture, 2021). Assuming average households have three individuals, the total household food stamps value would be \$2,982 on average. Thus a \$1,000 cash transfer would increase

earnings by 0.055 percent ($0.0016 \times (1000/2982)$). Then, we convert the intent-to-treat estimate to an estimate of the treatment effect on the treated. Using the Panel Study of Income Dynamics, the authors estimate that 16 percent of children participated in food stamps between 1975 and 1977. Thus, we divide 0.055 percent by 0.16, yielding 0.34 percent. The authors report that the natural log of the average labor income of the full samples is 10.57, which equals \$38,948.67. Income data spans from year 2000-2013 so the midpoint is year 2006. \$38,948.67 in 2006 dollars equals to \$49,169 in 2019 dollars. Thus, the estimate becomes \$168 (0.0034×49169) increase in income per year. As the paper studies the impact of exposure from conception (age -1) to age 5, we (conservatively) assume that child recipients were exposed to food stamps through the entirety of childhood (from age -1 to age 17) but only derived benefits for future earnings during the first 7 years of payments. We multiply results by the 7/19 of years in which they derive benefits, decreasing the impact to \$62. Using equation 1, we conclude that the present discounted value of increased earnings in adulthood is \$1,397.

Bastian and Micheltmore (2018)

Bastian and Micheltmore (2018) found that an additional \$1,097 in EITC (2019 dollars) exposure during childhood was associated with an increase in earnings of \$646.1 (s.e. 818.3) among children exposed between ages 0 and 5, an increase in earnings of \$42.4 (s.e. 415.1) among children exposed between ages 6 and 12, and an increase in earnings of \$564.0 (s.e. 244.9), among children exposed between ages 13 and 18. As discussed in the section on children's educational attainment, the exposure was also associated with a 0.012 (s.e. 0.003) or 1.2 percentage-point higher probability of completing high school, a 0.013 (s.e. 0.005) or 1.3 percentage-point higher probability of completing college and a 0.008 (s.e. 0.004) or 0.8 percentage-point higher chance of being employed in young adulthood among children exposed between ages 13-18. The 1968-2013 waves of the Panel Study of Income Dynamics (PSID) were used to examine the impact of exposure to the federal and state EITC between 1967 and 1995 ($n=3,495$). The authors measured EITC exposure using the maximum potential federal and state credit a household could receive based on the year, state, and number of children in the household. F-statistics using this maximum credit to predict increased family income were well above the critical value for weak instruments.

To simplify our calculations, we first determined an average impact across all ages by multiplying each of Bastian and Micheltmore's estimates for the three age groups times the proportion of children in that age group. According to Bastian and Micheltmore (2018), children exposed to EITC from ages 0-5, from ages 6-12 and from ages 13-18 make up 21.6%, 40.4% and 38% of their samples, respectively. Thus, the weighted average impact is \$371 ($[\$646.1 \times 0.216] + [\$42.4 \times 0.404] + [\$564.0 \times 0.38]$). Bastian and Micheltmore (2018) measure earnings in 2013 dollars. \$371 in 2013 dollars is \$407 in 2019 dollars. We find that \$1,000 of EITC, in 2019 dollars, increased children's earnings in adulthood by \$371 ($407 \times (1000/1097)$). However, these results are for multiple years of exposure to the EITC and include all children in states in which the maximum EITC increased, not just recipient children. We assume the child was exposed to the EITC from age 0-17 (a total of 18 years), yielding a \$21 ($\$371/18$) increase in earnings per year of exposure. To convert this intent-to-treat estimate to an estimate of the effects on the treated, we divide \$21 by the percentage of EITC-eligible households that received the EITC in 1990 (the middle of the study period), which was 83% (Scholz 1994)¹⁰, resulting in a

¹⁰ Scholz (1994) estimated an EITC participation rate of 80.5-86.4 percent. We use the average of this range of values, which is approximately 83%.

\$25 increase in earnings for a \$1000 transfer. Using equation 1, we conclude that the present discounted value of increased earnings in adulthood is \$561, as a result of a \$1,000 cash transfer during childhood.

Price and Song (2018)

Price and Song (2018) found that an additional \$2,962 (2019 dollars) in cash transfers annually for three to five years decreased children's future earnings by \$356 (s.e. 601). As elaborated in the sections on children's health, adult health and adult parent longevity, in consequence of the transfer, the probability of children applying for disability benefits (either means-tested and non-means-tested) in adulthood increased by 0.537 percentage points (s.e. 1.25), disability benefits application rate increased among parents by 0.063 (s.e. 0.0199) or 6.3 percentage points, and the likelihood of death rose among parents by 0.0138 (s.e. 0.0196) or 1.38 percentage points. They used long-term outcomes of the Seattle-Denver Income Maintenance Experiment (SIME/DIME) to examine the impact on families (n=52,867) who were randomized to receive cash transfers. On average, the treatment group received \$2,962 (2019 dollars) more in transfers annually than the control group for either three or five years, depending on treatment group. Long-term outcomes were measured by matching experimental data with data from the Social Security Administration and Washington State Department of Health. Regressions were conducted via least square, with the main independent variable being a dummy on treatment status. Other controls in the model included indicators for treatment location, race, family type, gender, manpower treatment status, birth date/age and year fixed effects.

\$356 is measured in 2013 dollars and in 2019 dollars it would be \$391. Adjusting for years of exposure, we divide -\$391 by 4 (the unweighted average of 3 and 5) and derive -\$98. Finally, to estimate the impact of a 1,000 transfer, we multiply -\$98 by (1000/2962). The final estimate is a \$33 decrease in earnings per year, or \$746 decrease in the lifetime.

Children's Health

Monetizing the Value of Life and Health:

We follow the standard practice to measure the monetary value of improvements in health using quality-adjusted life-years (QALY). QALY quantifies the impact of disease burden on a person's life expectancy as well as quality of life. In our analysis, we use a QALY value of \$126,628 for perfect health and \$0 for death, meaning that a year of living in perfect health has a monetary value of \$126,628. This value falls within the value of a QALY in the US recommended by the Institute for Clinical and Economic Review (Institute for Clinical and Economic Review 2020). We assume that the change in the value of a year of life as one moves from perfect health to death is linear.

It is also worth noting here that we use mother-reported measures of their own and their child's overall health status as our key measure of health impacts. Other measures available are consistent with the mother-reported measures, but mothers' reports are the only measures that are common to all of the studies we use. Self-reports of health status such as we are using have also been documented to be a good predictor of longevity (McGee et al. 1999; Miilunpalo et al. 1997). Overall health status is reported as either excellent, very good, good, fair, or poor, with excellent health having a rating of 5 and poor health a rating of 1. We add another status, "death", with a rating of 0, for self-reported health to fully capture the value of QALY. Excellent

health corresponds to a full QALY value of \$126,628 and death corresponds to a QALY value of \$0.¹¹

Averett and Wang (2018)

Averett and Wang (2018) found that increased income under the 1993 federal EITC expansion had the following effects among households with two or more children relative to one-child households with smaller EITC transfers: increased mother-rated reports of children's health by 0.2294 points (s.e. 0.2000) on a scale from 1 to 4 (with 1 being poor, 2 being fair, 3 being good and 4 being excellent), decreased accident rates by 0.0546 (s.e. 0.0275) or 5.46 percentage points, increased the frequency with which mothers sought medical attention for a child's illness by 0.0274 (s.e. 0.0410) or 2.74 percentage points, and increased reports of behavioral problems (measured by a standardized z-score) by 0.2435 standard deviations (s.e. 1.1146). The study used the 1979 National Longitudinal Survey of Youth (NLSY79) and the NLSY79 Child and Young Adult (NLSCYA) data and focused on mothers with fewer than 13 years of education as a proxy EITC eligibility (n=12,686). Difference-in-difference and mother fixed effects frameworks were employed to measure the increase in mothers' report of overall child health. Models controlled for characteristics of mothers and children and for state fixed effects.

We first use the result on mother-rated health for calculation. In order to examine the average impact of a \$1,000 cash transfer we first establish that as a result of the EITC expansion, families with two or more children received an increase in EITC benefits that are \$1,226 (2019\$) greater than families with one child.¹² Therefore, we estimate that mother-rated health of children would increase by 0.19 points ($0.2294 \times 1000 / 1226$) as a result of a \$1,000 cash transfer. Next, we divide the impact by years of exposure. We assume children were exposed to the EITC through the entirety of childhood or 18 years (from age 0-17), decreasing the impact to 0.01 points per year. We value the benefit of improved health by utilizing the QALY value of \$126,628, described in greater detail in the children's health section in the main text. To make sure that mother-rated health captures the full value of QALY, we add a category of "death" to the scale and give it a value of zero. Thus, death represents \$0 of QALY while excellent health (value of 4) represents the full value of \$126,628. A 0.01-point increase in absolute value of the scale thus represents a 0.26 percent improvement in health ($0.01/4$). We therefore estimate that a \$1,000 cash transfer increases quality of life by a value of \$329 every year ($\126628×0.0026) during childhood.

We then use the result on the frequency with which children suffer from accidents or injuries for the calculation. Having adjusted for \$1,226 EITC benefits in 2019 dollars and 18 years of exposure, we obtain a decrease in frequency of 0.25 percentage points per year. To keep the calculation consistent throughout the children's health section, we assume that moving from having no accidents or injuries to having accidents or injuries captures one-sixth of the value of QALY (this assumption is tied to the calculation of metabolic syndrome index in Hoynes et al. 2016, as discussed later). We estimate that a 0.25 percentage points decrease in frequency of

¹¹ Ideally, valuing quality of life is also done with a more detailed measure of current health but such data were not available in most studies and self reported health is considered to be very reliable. The National Center for Health Statistics has used similar measures of self-reported health (paired with physical limitations) to measure quality of life using a scale in which zero is death and 1 is excellent health (Gold et al., 1996).

¹² According to Averett and Wang (2018), families with one eligible child receive an average EITC benefit of \$822 in 1992 dollars pre-expansion and \$1374 post-expansion. Families with two or more eligible children receive an average EITC benefit of \$747 pre-expansion and \$1970 post expansion. The increase for families with two or more eligible children is \$671 in 1992 dollars greater than the increase for families with only one child.

accident and injuries per year translates into a \$52 increase in children's health per year ($(\$126628/6) * 0.0025$) during childhood.

The result on the frequency of illness that requires medical attention could suggest a decline in children's health. Having adjusted for \$1,226 EITC benefits in 2019 dollars and 18 years of exposure, we obtain a 0.12 percentage-point increase in the frequency per year. Assuming that having an illness that requires medical attention captures one-sixth of the value of QALY, we estimate that children's health decreased by \$26 per year during childhood.

The result on children's behavioral problems also suggests a decline in children's health. Having adjusted for \$1,226 EITC benefits in 2019 dollars and 18 years of exposure, the increase in behavioral problems becomes 0.01 standard deviation per year. Assuming that one standard-deviation of the behavioral health measure represents one-sixth of \$126,628, a 0.011 standard-deviation increase in behavioral problems per year represents a loss of \$233 per year.

We calculate the present discounted value using equation (1). We assume a discount rate of $i=0.02$. We do not discount the yearly benefit at age 9 and start discounting the yearly benefit from age $a=10$ to $A=21$. Giving all four results equal weight, we obtain a mean present discounted value of \$354.

Hoynes et al. (2016)

Hoynes et al. (2016) examine the long-term health and economic impact of exposure to food stamps between conception and age 5 using the Panel Study of Income Dynamics (PSID). They found that among the full sample, exposure to food stamps from conception to age 5 decreased the probability of having metabolic syndrome by 0.438 (s.e. 0.204) standard deviations and increased the probability of reporting good health by 0.292 (s.e. 0.133) or 29.2 percentage points. Metabolic syndrome is measured using an average standardized z-score of five binary components (obesity, high blood pressure, diabetes, heart disease, and heart attack). Report of good health is based on a self-reported health measure, with scale from 1 to 5 (with 1 being excellent, 2 being very good, 3 being good, 4 being fair, and 5 being poor). Both results on metabolic syndrome and report of good health are treatment-to-treated effects and suggest improvement in health.

We first calculate an estimate using the result on metabolic syndrome. As discussed above, average annual food stamps values per person in 1972 (near the midpoint of the study period) were \$994 in 2019 dollars (Department of Agriculture, 2021); assuming average households have three individuals, the total household food stamps value would be \$2,982 on average. Children in the sample were exposed to food stamps for 7 years (from conception (age -1) to age 5). As a result, a \$1000 cash transfer would lead the probability of metabolic syndrome to decrease by 0.021 standard deviations ($0.438*(1000/(2982*7))$). As the paper studies the health impact of exposure from conception (age -1) to age 5, we (conservatively) assume that child recipients were exposed to food stamps through the entirety of childhood (from age -1 to age 17) but only derived benefits for future health during the first 7 years of payments. To measure the impact per year of payments, we multiply results by the 7/19 of years in which they derive benefits, decreasing the impact to 0.0077 ($0.021*7/19$) standard deviations per year. We assume that six standard deviations of the metabolic syndrome index approximately capture the full range of quality of life: a standard deviation of -3 would equate to a QALY value of \$0 and a standard deviation of 3 would equate to a QALY value of \$126,628. We assume the distance between each standard deviation (an absolute value of 6) is equal so one standard deviation captures 1/6 of the value of \$126,628. If a \$1000 increase in household income during

childhood from a cash transfer decreases the probability of having metabolic syndrome by 0.0077 standard deviations per year, the benefit is then \$163 ($0.0077 \times (126628/6)$) per year.

An estimate can also be calculated based on the result on probability of reporting good health. A \$1,000 cash transfer would lead the probability of reporting good health to increase by 1.4 percentage points ($0.292 \times (1000 / (2982 \times 7))$). To measure the per year impact, we assume that children are exposed to food stamps during the entire childhood (age -1 to 17) but only derive health benefits for 7 years (age -1 to age 5). Thus, we multiply 1.4 percentage points by $(7/19)$ and obtain a per year impact of 0.52 percentage points. We measure improved health using QALYs. We measure quality of life using a scale that includes death and the five categories in the self-rated health measure (poor, fair, good, very good, or excellent). With death having a value of 0 and excellent health having a value of 5, the maximum increase in health would be an increase of 5 points. We equate death to a QALY value of \$0 and equate excellent health to a QALY value of \$126,628. Therefore, an increase of one unit of health quality for one year would be valued at $(126,628/5)$. A 0.52 percentage points per year increase in probability of good health would result in a benefit of \$131 ($0.0052 \times (126,628/5)$) per year.

We assume the average age of child beneficiaries to be 9 and the average age of death to be 78 and that the increase in physical health occurred from age 22 to age 78. Giving results on metabolic syndrome and probability of good health equal weight, we conclude that the present discounted value of improved health is \$3,917.

Bailey et al. (2020)

Bailey et al. (2020) found that exposure to food stamps from conception to age five increased physical ability and health, measured between the ages of 25 and 46 using a comprehensive index, by 0.0013 standard deviations (s.e. 0.0013).

The health index used by Bailey et al. (2020) is a standardized z-score and measures whether the respondent has a work disability, ambulatory difficulty, cognitive difficulty, independent living difficulty, vision or hearing difficulties, and/or self-care difficulty. The estimate of 0.0013 standard deviations improvement in physical health was derived from a sample that included all children who were exposed to food stamps rollout and was not limited to recipients. To adjust for this, we divide 0.0013 by the percentage of children in this age group who received food stamps, 16% (calculated by Bailey et al. using the PSID), increasing the impact to 0.008 standard deviations. Since children in the sample were exposed to food stamps for 7 years (conception to age 5), we divide 0.008 standard deviations by 7 and yield 0.001 standard deviations.

Again, the typical household food stamps value in 2019 dollars would be \$2,982. The impact of a \$1,000 benefit in 2019 dollars is thus 0.001 times the ratio of \$1000/\$2982, or 0.0004. As the paper studies the health impact of exposure from conception (age -1) to age 5, we (conservatively) assume that child recipients were exposed to food stamps through the entirety of childhood (from age -1 to age 17) but only derived benefits for future health during the first 7 years of payments. To measure the impact per year of payments, we multiply results by the $7/19$ of years in which they derive benefits, decreasing the impact to 0.0001 ($0.0005 \times 7/19$) standard deviations per year.

Assuming that one standard-deviation of the physical health measure represents one-sixth of \$126,628, a 0.0001 standard-deviation increase in physical health per year is equivalent to \$3 ($0.000143 \times 126628 / 6$) per year. Using equation (1), we conclude that the present discounted value of increased health in adulthood is \$81 as a result of a \$1,000 cash transfer during childhood.

Price & Song (2018)

Price and Song (2018) found that a transfer of \$2,962 (2019 dollars) for three to five years increased the probability of children applying for disability benefits (either means-tested and non-means-tested) in adulthood by 0.537 percentage points (s.e. 1.25).

Given that a treated family received an average of \$2,962 (2019 dollars) more transfer income for 3 to 5 years, we adjust the 0.537 percentage point impact on future disability assistance participation based on the unweighted average of 4 years of exposure. A cash transfer of \$1,000 would imply an impact of 0.181 percentage point ($0.537 \times 1000 / 2962$) for 4 years of exposure, or 0.0453 percentage points ($0.181 / 4$) higher disability assistance participation per year of childhood receipt. Assuming that moving from not applying for disability benefits to applying for disability benefits reflects a decline in one-sixth of value in health, we estimate that a \$1,000 cash transfer decreases quality of life by a value of \$10 every year ($-0.000453 \times 126628 / 6$). Using equation (1), we conclude that a \$1,000 transfer is associated with a present discounted value of -\$255 in children's health in adulthood.

Braga et al. (2020)

Authors found that exposure to EITC during childhood increased health between ages 22-27: increased the probability of having excellent or very good health by 0.017 (s.e. 0.004) or 1.7 percentage points, decreased the probability of being obese by 0.008 (s.e. 0.004) or 0.8 percentage points, decreased the probability of having functional limitation by 0.004 (s.e. 0.002) or 0.4 percentage points, and decreased the probability of having high blood pressure by 0.001 (s.e. 0.002) or 0.1 percentage points. Authors then examined whether the health impact between ages 22-27 would differ by the age of exposure. When health was measured by the probability of reporting excellent or very good health, exposure between birth and age 5 increased the probability by 0.010 (s.e. 0.004) or 1 percentage point, exposure between age 6-12 increased the probability by 0.004 (s.e. 0.003) or 0.4 percentage point, and exposure between 13-18 increased the probability by 0.005 (s.e. 0.002) or 0.5 percentage point. Authors also examined whether the beneficial health impact was persistent and found that the impact could last until age 51. Authors used 1968-2017 Panel Study of Income Dynamics (PSID) data. The analysis sample included 2,393 individuals. Authors took advantage of the variation in maximum EITC credits across states, time, and family size. Treatment was average annual maximum EITC credit during childhood, given the child's state of residency, year, and size of family. Other controls in the model included individual characteristics (year of birth, race, gender, parents' education, parents' marital status, sibling fixed effects), state characteristics (GDP per capita, unemployment rate, income tax rate, minimum wage, maximum welfare benefits, tax revenues), state fixed effects, year fixed effects, state-by year fixed effects, and state-specific time trends.

All the health impacts summarized above come from a \$100 increase in the average annual maximum EITC credit exposed (2017 dollars), or \$104.57 in 2019 dollars. If being exposed to an increase of \$104.57 in annual maximum EITC credit increased the probability of having excellent or very good health by 1.7 percentage points, then we linearly extrapolate to \$1,000

EITC benefits would increase the probability by 16 percentage points ($1.7 \times 1000 / 104.57$). To avoid overstating the per-year effect, we assume that children in the analysis sample are exposed to EITC through their entire childhoods (age 0-17), a total of 18 years. Dividing 16 percentage points by 18 gives us a 0.9 (16/18) percentage-point increase per year. To obtain a treatment-on-the-treated effect, we further divide the estimate by an estimated EITC take-up rate. The middle of the study period is 1993 and according to Scholz (1994) the EITC participation rate in 1990 was around 83 percent. Dividing 0.9 percentage points by 0.83 gives us a 1.09 (0.9/0.83) percentage-point increase. We value improvement in health using QALY. We measure quality of life on a scale of 0-5, with 0 corresponding to death and 5 corresponding to excellent health (full scale includes death, poor, fair, good, very good, or excellent, with each corresponding to 0, 1, 2, 3, 4 and 5 respectively, so the maximum increase is 5 points). If death has a QALY value of \$0 and excellent health has a full QALY value of \$126,628, then an increase in one unit of health corresponds to $1/5^{\text{th}}$ of the value of QALY. Valuing the 1.09 percentage-point increase by $1/5^{\text{th}}$ of QALY results in a \$276 increase in health per year. We assume that increased health in adulthood takes place from ages 22-78. We conclude that a \$1,000 increase in household income from cash transfer per year would increase the present discounted value of children's adulthood health by \$7,351.

Song (2019)

The author found that exposure to an increase of \$1,000 in the maximum EITC in utero through age 18 increased the probability of being in good health by 0.036 (s.e. 0.011) or 3.6 percentage points. The same increase led to a 0.075 (s.e. 0.021) or 7.5 percentage-point decrease in obesity, a 0.058 (s.e. 0.021) or 5.8 percentage-point decrease in smoking, and a 0.032 (s.e. 0.015) or 3.2 percentage-point increase in drinking. When differentiating exposure by age, the author found that an increase of \$1,000 in the maximum EITC in utero, ages 0-5, and ages 13-18 would increase the probability of being in good health by 0.035 (s.e. 0.015), 0.076 (s.e. 0.016), and 0.054 (s.e. 0.016), respectively. The same exposure at ages 6-12 would instead decrease the probability of being in good health by 0.031 (s.e. 0.023). The author used 1968-2017 PSID data and took advantage of the variation in maximum EITC credits across states, time, and family size. Treatment was average annual maximum EITC credit during childhood given the child's state of residency, year, and size of family. Other controls in the model included birth cohort fixed effects, state and year fixed effects, state-specific time trends, number of siblings, gender and race.

The 3.6 percentage-point increase in the probability of being in good health was a result of a \$1,000 increase (2017 dollars) in maximum EITC exposure, the equivalent of \$1,045.67 in 2019 dollars. A \$1,000 increase in EITC exposure in 2019 would thus lead to a 3.4 ($3.6 \times 1000 / 1045.67$) percentage-point increase in probability of being in good health. We divide 3.4 percentage points by 18 assuming the exposure effect is spread across all childhood years, and we obtain a per-year increase of 0.19 (3.4/18) percentage points. To obtain a treatment-on-the-treated effect, we further divide 0.19 by an estimated EITC take-up rate during the middle of the study period (year 1993), around 0.83. This yields a 0.23 percentage-point increase. Valuing the 0.23 percentage point increase by $1/5^{\text{th}}$ of QALY results in a \$58 increase in health. We conclude that a \$1,000 increase in household income from cash transfers would increase children's adulthood health by \$58 per year. Assuming that increased adulthood health occurs between ages 22-78, we obtain a present discounted value of \$1,557 in increased adulthood health, following a \$1,000 increase in household income from cash transfer per year.

Child Longevity

Aizer et al. (2016)

Aizer et al. (2016) found that in adulthood, sons whose mothers had received Mothers' Pensions experienced an increase in longevity of 0.0158 (s.e. 0.007) or 1.16 years. The \$20 monthly transfer in 1922 would be worth \$307 in 2019, or \$3,684 annually for on average three years. Therefore, a \$1,000 transfer for one year would increase children's life by 0.10496 years ($1.16 \times (1000/3684)/3$). Applying the QALY value, an increase in longevity of 0.105 years would be valued at \$13,291 ($0.10496 \times \$126,628$). Using assumptions described above, we calculate the present discounted value as $B/(1+i)^{t-9}$, with $B=13291$, $i=0.02$ and $t=78$. We conclude that the present discounted value of increased longevity in adulthood is \$3,390 as a result of a \$1,000 cash transfer during childhood.

Bailey et al. (2020)

Bailey et al. (2020) found that exposure to food stamps from conception to age 5 increased longevity by 0.176 years (s.e. 0.030). This estimate includes all children who were exposed to food stamps and is not limited to recipients. To adjust for this, we divide 0.176 by the percentage of children in this age group who received food stamps, 16% (calculated by Bailey et al 2020 using the PSID). Thus, the treatment-on-the-treated outcome is 1.1 years ($0.176/0.16$). Children in the sample were exposed to food stamps for 7 years (conception to age 5), so the impact decreases to 0.16 ($1.1/7$) years in longevity. As discussed above, family food stamps value in 2019 dollars was \$2,982 on average. Thus, the impact becomes 0.05 years ($0.16 \times 1000/2982$). As the paper studies the impact of exposure from conception (age -1) to age 5, we (conservatively) assume that child recipients were exposed to food stamps through the entirety of childhood (from age -1 to age 17) but only derived benefits for future health during the first 7 years of payments. To measure the impact per year of payments, we multiply results by the 7/19 of years, decreasing the impact to 0.0194 years in longevity per year. We value this increase in life expectancy using QALY. The increase in longevity is thus worth \$2458 ($0.0194 \times \$126,628$). Assuming that the extension of life occurred at age 78 (conservative, given that the mortality improvements in this study occurred at ages 25-64 so with our assumption we are discounting mortality improvements by more years), we calculate the present discounted value of increased longevity in adulthood as $2458/(1.02)^{69}$, which implies a benefit of \$627 as a result of a \$1,000 cash transfer during childhood.

Avoided Health Expenditures for Children

Healthcare Expenditure Elasticity

We rely on the results of three studies to determine the rate at which healthcare expenditures decrease in relation to increases in health status. Although the studies are not causal, no quasi-experimental study exists, to our knowledge, examining this relationship.

Chern, Wan, and Begun (2002) found that a one percent increase in SF-36 score was associated with 0.19 percent ($p < 0.001$; s.e. not available) decrease in health expenditures. The study sample included 4,255 randomly selected Trigon BlueCross/BlueShield policyholders in Virginia. The sample was limited to adults between the ages of 18 and 64. Health status was measured in 1994 using the MOS 36-Item Short-Form Health Survey (SF-36). The SF-36 measure used included five dimensions of health and ranged from 0-100 points (physical functioning, role limitations due to physical health problems, body pain, general health, and social functioning).

Healthcare expenditures were measured in 1995 and include out-of-pocket expenditures and expenditures covered by insurance. Using a structural equation modeling framework, the authors examined the association between health status and healthcare expenditures, controlling for age, gender, marital status, educational attainment, occupation, race, smoking behavior, high cholesterol, high blood pressure, diabetes, household income, point-of-service (POS) health plans, and Preferred Provider Organization (PPO) health plans. Since their estimate indicates that a one percent increase in health leads to a 0.19 percent decrease in health expenditures, we infer an elasticity of 0.19.

Lima and Kopec (2005) use the 1994-1998 Canadian National Population Health Survey (n=2,084) to examine the impact of health status on health care expenditures. Health status was measured using the Health Utilities Index (HUI). Using a multivariate log-linear model, controlling for sociodemographic characteristics, they found that a 0.1 improvement in HUI is associated with a 10% reduction in annual health care costs (s.e. not available). We infer an elasticity of one.

Desalvo et al. (2009) examine healthcare expenditures by self-rated health status using the 2003-2005 Medical Expenditure Panel Survey (n=7,948). The authors descriptively find that the average healthcare expenditures for individuals with excellent, very good, good, fair, and poor health were \$1,654, \$2,640, \$4,228, \$9,831, and \$12,709, respectively. To estimate the elasticity of healthcare expenditures in relation to health status we determine the change in healthcare expenditures associated with a one unit increase in health status; We measure the elasticity associated with moving from good to very good health because the average respondent in the sample reported their health to be good. We use DCG/HCC score to measure percentage change in health as it performs the best in predicting health expenditure according to the paper. Moving from good to very good health involves a 25 percent increase in health (good health corresponds to a DCG/HCC score of 1.66 and very good health corresponds to a score of 1.24). Improving health from good to very good is associated with a \$1,588 or 38%, decrease in healthcare expenditures. The results imply that the elasticity of health expenditures in relation to health status is 1.48 (38% /25%). In conclusion, the results indicate that the elasticity of healthcare expenditures in relation to health status ranges from 0.19 to 1.48, for an average elasticity of 0.84.

Decline in Healthcare Expenditures from 6 Months of Age Onwards

We rely on estimates from the Centers for Medicare & Medicaid Services (CMS, 2019) on average healthcare spending by age to estimate the value of decreased child healthcare expenditures. The CMS's most recent estimate of healthcare spending by age was from 2010. Age was divided into five categories: 0-18, 19-44, 45-64, 65-84, and 85 or older. Healthcare expenditures include costs to insurer and patient but exclude non-personal health care spending (government administration and the net cost of private health insurance, noncommercial research, investment in structures and equipment, and government public health activities). Average per capita spending was \$3,628 among children, \$4,422 among adults 19-44, \$8,370 among adults 45-64, \$15,857 among adults 65-84, and \$34,783 among adults 85 and older. CMS (2020) additionally estimates per capita healthcare expenditures projections for 2019. However, the projections were not disaggregated by age. In 2019, per capita healthcare

expenditures were \$9,825.¹³ This represents a 38% increase from 2010. Assuming growth in healthcare expenses was consistent by age group, we conclude that 2019 per capita health care spending was \$5,007 ($3628+(0.38*3628)$) among children, \$6,102 among adults 19-44, \$11,551 among adults 45-64, \$21,883 among adults 65-84, and \$48,001 among adults 85 and older.

In the section on children's health, we find that children's health in childhood (ages 9-21) would increase by 0.02 percent of QALY per year and children's health in adulthood (ages 22-78) would increase by -0.008 percent, 0.002 percent, 0.116 percent, 0.05 percent, or 0.2 percent of QALY per year as a result of \$1,000 cash transfer. We also find that one percent increase/decrease in health would lead to 0.84 percent decrease/increase in health expenditures. To calculate the monetary value of change in health at a certain age, we multiply the percentage change in health by the healthcare expenditure elasticity and then by per capita health care spending of that age. To calculate the present discounted value, we assume children experience avoided health expenditures from ages 9-21 as a result of increase in health during childhood and experience avoided health expenditures from ages 22-78 as a result of increase in health in adulthood. We conclude that a \$1,000 transfer is associated with a \$12 decrease in healthcare expenditures from ages 9-21 and an average of \$170 decrease in healthcare expenditures from ages 22-78.

Parents' and Other Adults' Health

Larrimore (2011)

Larrimore (2011) found that a \$1,541 (2019\$) increase in income led to a 0.0032 (s.e. 0.0028) or 0.32 percentage-point increase in the likelihood of having excellent health, a 0.0013 (s.e. 0.0011) or 0.13 percentage-point decrease in the probability of being in poor health and a 0.0328 (s.e. 0.01659) or 3.28 percentage-point decrease in the probability of having any functional limitation. The author examined the impact of income on health among parents 22-62 years-old with incomes below 200% of the federal poverty line using the maximum state plus federal EITC as an instrumental variable for income. Results were examined using the 1992-2005 Survey of Income and Program Participation (SIPP) panels (n=85,397). F-statistic results indicated the maximum EITC was a strong predictor of post-tax income. Health was measured using self-reported health status. The two-stage least squares regressions controlled for state of residence, whether the 1996 welfare reform had been enacted at the time of the observation, year, age, age-squared, race/ethnicity, gender, education, marital status, health insurance status, number of children in the household, and whether the respondent lives in a metropolitan area.

To stay consistent with other studies cited in the adult health section, we use the result on probability of excellent health for calculation. The midpoint of the study period is 1999 and the author was estimating the average marginal effect of a \$1000 increase. A \$1,000 increase in household income in 1999 is equivalent to an increase of \$1,541 in 2019. Thus, we find a 0.21 percentage point ($0.32*(1000/1541)$) increase in the likelihood of having excellent health. Larrimore measures self-rated health the same year as the transfer was received so no

¹³ Centers for Medicare & Medicaid Services (2020) estimate per capita spending using both person and non-personal health expenses (\$11,559). Results further indicate that aggregate personal healthcare expenses make up 85% of total healthcare spending. As a result, we assume that per capita personal healthcare spending is 85% of total per capital healthcare spending, providing a final per capita spending of \$9,825.

adjustment for years of exposure is needed. We measure improved health using QALYs valued at \$126,628, described in greater detail in the children's health section in the main text. We measure quality of life using a scale that includes death and the five categories in the self-rated health measure (poor, fair, good, very good, or excellent). With death having a value of 0 and excellent health having a value of 5, the maximum increase in health would be an increase of 5 points. Therefore, an increase of one unit of health quality for one year would be valued at \$25,326 ($\$126,628/5$). A 0.21 percentage point increase in the likelihood of having excellent health would then result in a benefit of \$53 ($0.0021 * \$126,628/5$) per year.

We calculate the present discounted value assuming the adult is 38 at the first transfer and that adult recipients receive health benefits for the remainder of their life. We conclude that a \$1,000 transfer improves the present discounted value of adult health by \$1,491.

Evans and Garthwaite (2014)

Evans and Garthwaite (2014) found that the 1994 EITC expansion increased the probability of mothers' reporting very good or excellent health by 0.0135 (s.e. 0.0075) or 1.35 percentage points, increased mothers' poor physical health days in the past month by 0.0105 (s.e. 0.039) or 1.05% (out of an average of 2.65 days among the treatment group), decreased mothers' poor mental health days in the past month by 0.0754 (s.e. 0.0328) or 7.54% (out of an average of 4.52 days among the treatment group), and decreased the total number of risky health conditions (i.e., total cholesterol, systolic blood pressure, any risky inflammatory condition) by 0.235 (s.e. 0.095) or 23.5% (out of an average of 1.108 conditions among the treatment group). The authors use a difference-in-difference framework, exploiting the 1994 EITC increase in the relative benefit for families with two or more children relative to those with one child. Analyses were conducted using the 1993-2001 Behavioral Risk Surveillance Survey (BRFSS) and the sample was restricted to mothers with a high school education or less ($n=82,907$). Models control for individual characteristics, state level fixed effects, and time fixed effects.

We use the 1.35 percentage-point increase in the probability of reporting very good or excellent health for the calculation. However, this result is not limited to EITC beneficiaries. In 1999 (near the midpoint of the study period of 1997), 75% of eligible households received the EITC (U.S. Government Accountability Office, 2001). Thus, we divide 1.35 percentage points by 75%, determining that the additional EITC cash transfer to larger families increased the probability of mothers having very good or excellent health by 1.8 percentage points among recipients. The authors find that conditional on receipt, on average, households with two or more children receive \$864 (2019\$) more in EITC benefits than families with one child, which suggests that a \$1,000 increase in EITC payments would increase the likelihood of having very good or excellent health by 2.08 percentage points ($1.8 * (1000/864)$). The authors measure improved health for 6 years after the increase in transfers. Therefore, per year of exposure the probability of having excellent health increases by 0.35 percentage points ($2.08/6$). We measure the monetary value of improvements in health using QALYs of \$126,628. QALYs are measured using a scale which includes death and the five self-rated health categories. Evans and Garthwaite (2014) examine the probability of having very good or excellent health, so we measure the benefit of moving up one unit, from good to very good health. An increase of one unit of quality would be valued at ($\$126,628/5$). A 0.35 percentage point increase in the likelihood of having very good health then results in a benefit of \$88 ($0.0035 * 126,628/5$) per year.

We also use the 1.05 percent increase in the number of bad physical health days for the calculation, because unlike other results, it reflects a decline in mother's health. To obtain the effect among recipients only, we divide 1.05 percent by 75% (the percent of eligible household that received EITC in 1999), yielding 1.4 percentage points. Adjusting the impact for a \$1,000 increase, we increase the impact to 1.62 percentage points ($1.4 \times (1000/864)$). Lastly, to obtain the impact of an exposure of one year, we divide 1.62 percentage points by 6, yielding 0.27 percentage points. The increase in the number of bad physical health days can be considered as an increase of the probability of having bad health. We measure the loss of moving down one unit, from fair to poor health. Multiplying 0.27 percentage points by $(\$126,628/5)$, we yield a loss of \$68 per year.

We do not use the decrease in total number of risky conditions for the calculation because we are unsure whether such bio-marker measures the continuous health degradation or the probability of future health risks. We do not use the decrease in the number of poor mental health days for the calculation because it is a partial measure of overall health and we remain consistent in not counting the mental health outcome equally with the more comprehensive outcome (in this case the probability of reporting very good or excellent health).

We calculate the present discounted value assuming the adult is 38 at the first transfer and that adult recipients receive health benefits for the remainder of their life. Since the paper has four results on health, three of which suggest improvement in health and one suggests a deterioration in health, we give the result on probability of good/excellent health a weight of 3/4 and the result on bad physical health day a weight of 1/4. The weighted present discount value is \$1,385.

Morgan et al. (2020)

Morgan et al. (2020) found that a 10-percentage point higher state EITC (relative to the federal EITC) was associated with a decrease in the prevalence of individuals having frequent poor mental health by 97 [-237.2, 42.6] individuals per 100,000 and a decrease in the prevalence of individuals with frequent poor physical health by 150 [-284.4, -14.9] individuals per 100,000, with 95-percent confidence intervals shown in brackets. The prevalence of having suboptimal overall health increased by 31 per 100,000 individuals [-123.3, 185.9]. The authors use a difference-in-difference framework to examine the impact of increased state EITC transfers on health using state and year variation in the EITC. Analyses were conducted using the 1993-2016 Behavioral Risk Factor Surveillance System survey. The sample was limited to adults with a high school education or GED equivalent or less ($n=2,884,790$). Frequent poor physical health is measured as whether the respondent reported having had 14 or more days in the past month in which they would describe their physical health as "not good." Overall health is measured by a self-reported health scale from 1 to 5 (with 1 being excellent, 2 being very good, 3 being good, 4 being fair, and 5 being poor). Respondents with fair or poor health are regarded as having suboptimal health. States with non-refundable EITCs were considered to not have an EITC and were lumped with non-EITC states. Models control for state minimum wage, state GDP, adoption of Medicaid expansion, state fixed effects, and year fixed effects.

Results indicated that a 10-percent higher state EITC (relative to the federal EITC) was associated with a decrease in the number of people reporting frequent poor physical health of 150 per 100,000 individuals, or 0.15 percentage points. In 2004 (the approximate midpoint of the study period), the average federal EITC was \$1,834 (Kneebone, 2007). Therefore, a 10

percent increase in the state EITC, relative to the federal EITC, would be \$183.4, the equivalent of \$248 in 2019 dollars. Thus, we estimate that a \$1,000 increase in cash transfers would decrease frequent poor physical health by 605 $((1000/248)*150)$ per 100,000 individuals, or by 0.605 percentage points. However, these results include all households with a high school education or less and are not limited to treated individuals. According to Internal Revenue Service (2013), 20% of households with a high school education or less are eligible for the EITC. To estimate the effect of the treatment on the treated, we divide 0.605 by 0.2, yielding 3.03 percentage points. Lastly, we adjust results for years of exposure. The authors do not describe the average years of exposure. We assume state EITC programs were implemented or expanded at the midpoint of the study period on average, 2004. This would indicate the average years of exposure was 12. Thus, we conclude that per year, a \$1,000 state EITC decreases the prevalence of frequent poor physical health by 0.25 $(3.03/12)$ percentage points. To remain consistent with the valuing procedures used in the remainder of adult health estimates, we assume the frequent poor physical health measure captures the equivalent of moving one-fifth of a QALY. Thus, results indicate a 0.25 percentage-point decrease would result in a benefit of \$64 $(0.0025*126628/5)$ per year.

Following the same process, we monetize the result on suboptimal health, which indicates a decline in adult health. Results indicate that a 10-percent higher state EITC was associated with a 0.031 percentage-point increase the prevalence of suboptimal health. A \$1,000 transfer would increase the prevalence by 0.125 percentage points $(0.031*(1000/248))$. The treatment on the treated effect would be 0.625 percentage point $(0.125/0.2)$. Lastly, to adjust for years of exposure, we divide 0.625 by 12, yielding 0.05 percentage points per year. Multiplying -0.05 percentage point by $(\$126628/5)$, we arrive at a loss of -\$13 in adult health per year.

We do not use the result on poor mental health for the calculation because it is a partial measure of overall health and we remain consistent in not counting the mental health outcome equally with the more comprehensive outcome (in this case the prevalence of having suboptimal overall health).

We calculate the present discounted value assuming the adult is 38 at the first transfer and that adult recipients receive health benefits for the remainder of their life. Since the paper reports three results on health, two of which suggest an improvement of health and one of which suggests a decline, we give the result on physical health a weight of 2/3 and the result on suboptimal health a weight of 1/3. The weighted present discounted value is \$1,082.

Price and Song (2018)

Price and Song (2018) found that a \$2,962 (2019 dollars) transfer for three to five years resulted in a 6.28 percentage-point increase in disability benefits application rate among parents (s.e. 0.0199).

Having adjusted the result to reflect the impact of a \$1,000 dollars cash transfer, the impact decreases to 2.12 percentage points $(6.28* (1000/2962))$. Having adjusted for 4 years (an unweighted average of 3 and 5) of exposure the impact further decreases to 0.53 percentage points $(2.12/4)$. We value the benefit of improved health by utilizing QALY, valued at \$126,628. To remain consistent with other impact studies in this section, we assume that moving from not applying for disability benefits to applying for disability benefits reflects a decline in one-fifth of the value of QALY. We multiply $\$126628/5$ by 0.53%, yielding a loss of \$134 in adult health per year. We calculate the present discounted value assuming the adult is 38 at the first transfer and

that adult recipients receive health benefits for the remainder of their life. We conclude that a \$1,000 transfer is associated with a present discounted value of -\$3,806 in adult's health.

Parents' or Other Adults' Longevity

Price and Song (2018)

Price and Song (2018) found that an additional \$2,962 (2019 dollars) in cash transfers annually for three to five years increased the likelihood of death among adults by 0.0138 (s.e. 0.0196) or 1.38 percentage points.

When we adjust their estimate to measure the impact of a \$1,000 transfer, the impact decreases to 0.47 percentage points ($1.38 \times (1000/2962)$). To account for years of exposure, we divide by 4 (an unweighted average of 3 and 5), decreasing our estimate to 0.12 percentage points ($0.47/4$). We value the change in mortality by utilizing QALY, valued at \$126,628. As a result, a 0.12 percentage point increase in the probability of death represents a QALY decrease of \$147 (126628×0.0012). We calculate the present discounted value assuming that adults received the first transfer benefit at age 38 (based on the assumption that a parent is 29 at their child's birth (based on the mean age of mothers at birth as of 2019 according to CDC Vital Statistics)) and that the extension of life occurred at age 78 (life expectancy in the U.S. in 2018). We plug zero into the equation from ages 38-77 and -\$147 for age 78. We conclude that \$1,000 transfer is associated with a present discounted value of -\$67 in adult longevity.

Aizer et al. (2020)

Aizer et al. (2020) found that the Mothers Pension Program increased mother's longevity by 0.247 years (s.e. 0.494). Sample was restricted to mothers who had applied for the program only once no later than 1930 ($n=16,228$). The causal impact of the program was evaluated through OLS regressions that controlled for county fixed effect, application year fixed effect, individual control, county control and state control.

Mothers participated in the program for an average of three years. The \$20 monthly transfer in 1922 would be worth \$307 in 2019, or \$3,684 annually for on average three years. Adjusting the impact for a \$1000 transfer, mother's longevity would rise by 0.07 years ($0.247 \times (1000/3684)$). Adjusting for years of exposure, mother's longevity would rise by 0.02 years ($0.07/3$). We value the change in mortality using QALYs, valued at \$126,628. Therefore, a 0.02-year of increase in longevity is equivalent of \$2,830 (0.02×126628) (\$2,830 per year is around 2.23% of QALY per year) and the present discounted value is \$1,282.

Chetty et al. (2016)

Chetty et al. (2016) found that an increase in income from \$14,000 to \$20,000 (moving from the 10th to the 15th income percentile) was associated with an increase in longevity of 0.7-0.9 years (s.e. not available). The authors used population-level tax records and Social Security death records between 1999 and 2014 to examine the relationship between pre-tax income and life expectancy. The study included all individuals with incomes above zero between the ages of 40 and 76 with a valid Social Security Number in the specified years, and measured income using tax records. The authors defined income as adjusted gross income plus tax-exempt interest income minus taxable Social Security and disability benefits. Respondents were assigned a percentile rank from 1 to 100 based on their income relative to all other individuals with the same sex and age during this period. The relationship between income percentile and life

expectancy was approximately linear. Life expectancy was measured using the expected length of life for a hypothetical individual who faced a mortality rate at each age that matched those in the cross-section during a given year. Results were analyzed separately for men and women. To help mitigate concerns regarding reverse causality, mortality was measured two years after income (2001-2014). Their estimate of increased longevity implies that per \$1,000 increase in household income, the present discounted value of longevity is \$234. Even though the paper is not causal, its results do assist in establishing an upper-bound estimate.

Chetty et al. found that an increase in income from \$14,000 to \$20,000 (moving from the 10th to the 15th income percentile) was associated with an increase in longevity of 0.7-0.9 years (s.e. not available). Income was measured in 2012 dollars. The \$6,000 increase is the equivalent of \$6697 increase in 2019. Extrapolating from these results, a \$1,000 increase in income was associated with between a 0.1 ($0.7 \times (1000/6697)$) and 0.13 ($0.9 \times (1000/6697)$) year increase in life expectancy. We use the approximate midpoint of this range for our final estimate, 0.12 years. A QALY is valued at \$126,628 (described in greater detail in the children's health section in the main text). As a result, a 0.12-year increase in life represents a QALY increase of \$15,127. To account for years of exposure, we divide results by 21.5 (assuming that on average parents will have 2 children. Through CPS data we found that in two-children families, siblings are spaced 3.5 years apart on average, leading to parent's eligibility to receive a transfer for 21.5 years), decreasing our estimate to \$704. We calculate the present discounted value assuming that adults received the first transfer benefit at age 38 (based on the assumption that a parent is 29 at their child's birth (based on the mean age of mothers at births as of 2019 according to CDC Vital Statistics)) and that the extension of life occurred at age 78 (life expectancy in the U.S. in 2018). We conclude that \$1,000 transfer is associated with a present discounted value of \$319 in increased longevity.

Using three studies above, we conclude that the average present discounted value of increased adult longevity is \$511.

Avoided Health Expenditures for Parents and Other Adults

We rely on the Centers for Medicare & Medicaid Services (2019) to estimate average health expenditures. As explained above, using the Centers for Medicare & Medicaid Services (2019)'s results, we estimate that 2019 per capita health care spending averaged \$6,102 among adults 19-44, \$11,551 among adults 45-64, and \$21,883 among adults 65-84.

According to table 2, a \$1,000 cash or near-cash transfer increases an adult's health by -0.106 percent, 0.03 percent, 0.039 percent or 0.042 percent of QALY per year. We use the same calculation method in the section on children's health expenditures to convert this percentage change in health into monetary value of change in healthcare expenditures. We calculate the present discounted value of health expenditures, assuming that parents experience benefits in avoided health expenditures from ages 38-78. We conclude that a \$1,000 transfer is associated with an average present discounted value of \$3.24 in reduced health expenditures.

Children's Educational Attainment

Based on six studies below, we found that a \$1,000 cash transfer would increase children's years of schooling by 0.002-0.03 years. The mean increase is 0.01 years. We do not count increased education as a benefit because any benefit from increased education is already

counted in the benefit of increased earnings. However, as detailed later, we do use the results of these studies to calculate the increased cost posed by increased educational attainment.

Akee et al. (2010)

Akee et al. (2010) found among American Cherokee children, receiving tribal casino payments led to an increase in years of education of 0.379 (s.e. 0.447) and 0.117 (s.e. 0.304) years among the age 9 and age 11 cohorts, respectively, and an increase in the probability of graduating high school by age 19 of 0.156 (s.e. 0.073) or 15.6 percent for the age 9 cohort and 0.042 (s.e. 0.066) or 4.2 percent for the age 11 cohort. The authors used data from the Great Smoky Mountains Study (GSMS), which began in 1993 and included a representative sample of children aged 9, 11, and 13 in 11 counties in North Carolina (n=1,185). American Cherokee children within the included counties were oversampled (350). In 1996, the Eastern Band of Cherokees opened a casino. Each tribal member received a proportion of the casino's profits. The two youngest age cohorts (ages 9 and 11) were identified as "after-treatment" cases and the oldest cohort (age 13) functioned as the "before-treatment" case. Casino payments began in 1997, when children were 13, 15, and 17; therefore, each age cohort lived in homes in which the parent(s) received payments for 6, 4, and 2 years, respectively. Linear regression models controlled for the number of adults in the household eligible for the casino payments, the age cohort of the child, an interaction term of the age cohort and number of adults, household poverty status prior to the opening of the casino, sex of child, race of child, and education levels of both parents. Outcomes were measured at ages 19 or 21.

According to Akee et al. (2010), annual payments were an average of \$4,000 starting in 1996, which is the equivalent of \$6,538 in 2019. Therefore, to determine the change in educational attainment associated with a \$1,000 cash transfer, we divide results by (1000/6538). Age 9 cohort experienced a 0.06-year ($0.379 \times (1000/6538)$) increase in years of education and a 2.39 percent ($15.6 \text{ percent} \times (1000/6538)$) increase in the probability of graduating high school. Age 11 cohort experienced a 0.02-year ($0.117 \times (1000/6538)$) increase in years of education and a 0.64 percent ($4.2 \text{ percent} \times (1000/6538)$) increase in the probability of high school graduation. The age 9 cohort was exposed to transfers for a total of 6 years and the age 11 cohort was exposed for 4 years. A one year \$1,000 transfer increased the probability of graduating high school by between 0.16 percent ($0.64 \text{ percent} / 4$) and 0.4 percent ($2.39 \text{ percent} / 6$). It increased years of education by between 0.004-year ($0.02 / 4$) and 0.01-year ($0.06 / 6$). Since age 9-cohort makes up 54 percent of the post-treatment group and age-11 cohort makes up of 46 percent of the post-treatment group, the weighted increase in high school graduation is 0.29 percent ($0.16 \text{ percent} \times 0.46 + 0.4 \text{ percent} \times 0.54$) and the weighted increase in educational attainment is 0.007 years ($0.004 \times 0.46 + 0.01 \times 0.54$). If we want to express the increase in years of education in percentage, then given that the control group (households with no American Indian parent) has an average of 11.96 years of education, a one year \$1,000 transfer increased years of education by between 0.04 percent ($0.004 / 11.96$) and 0.08 percent ($0.01 / 11.96$). The weighted increase in years of education is 0.06 percent ($0.04 \text{ percent} \times 0.46 + 0.08 \text{ percent} \times 0.54$).

Maxfield (2015)

Maxfield (2015) found that a \$1,000 increase in maximum EITC led to a 0.0139 (s.e. 0.0078) or 1.39 percentage-point increase in the probability of completing one or more years of college at age 19, a 0.0207 (s.e. 0.0099), or 2.07 percentage-point increase in the probability of receiving a high school diploma or GED at age 19 and a 0.0295 (s.e. 0.0301) increase in years of schooling at age 19. The author used the 1979 National Longitudinal Survey of Youth (NLSY) and

corresponding child file. The data included children of all ages linked to their mother between 1988 and 2000, covering all major federal EITC expansions, and long-term outcomes for the children as young adults between 1994 and 2010. The sample was limited to children whose family was ever eligible to receive the EITC and to children who have a sibling in the sample (n=2,720). EITC exposure is measured based on the maximum federal and state EITC the household was eligible for by year, number of children, and state. Ordinary least squares (OLS) analyses controlled for child age and age squared, mother's score on the Armed Forces Qualification Test (AFQT), indicators for race, sex, birth order, and birth year, mother's age and age squared, mother's marital status, age of mother at birth of child, mother's educational attainment, the age the child would be expected to graduate high school, number of children in the household, maximum welfare benefit by state and year for a family of three, per pupil spending on K-12 public education in state and year, and state, year, and family fixed effects.

The author found that the average probability of obtaining a high school diploma or GED at age 19 is 75 percent, the average probability of completing one or more years of college at age 19 is 25 percent, and the years of schooling completed at age 19 on average is 12.07 years. Therefore, a 2.07 percentage-point increase in the probability of obtaining a high school diploma or GED represents a 2.76 percent increase, a 1.39 percentage-point increase in the probability of completing one or more years of college represents a 5.56 percent increase, and a 0.0295-year increase in years of schooling represents a 0.24 percent increase. The author states that a \$1,000 increase in maximum EITC benefit increased average estimated EITC payments by \$328 in 2008 dollars, the equivalent of \$384 in 2019 dollars. Therefore, a \$1,000 increase in real EITC payments increases the probability of receiving a high school diploma or GED by 7.19 percent (2.76 percent*(1000/384)), increases the probability of completing one or more years of college by 14.48 percent (5.56 percent*(1000/384)) and increases years of schooling by 0.64 percent (0.24 percent*(1000/384)). Children in the sample were an average of 8 years old, meaning they were exposed to the increased EITC for an average of 10 years (age 8-17). Thus, a one-year, \$1,000 increase in EITC payments increases the probability of receiving a high school diploma or GED by 0.72 percent, increases the probability of completing one or more years of college by 1.45 percent, and increases years of schooling by 0.06 percent. Lastly, since the sample includes children whose families were ever eligible for EITC, not those whose families have actually received EITC, we divide all impacts by 75 percent (U.S. Government Accountability Office, 2001), which is the percentage of eligible households that claimed EITC in 1999 (near the middle of the study period of 1994). The impact on receiving a high school diploma or GED becomes 0.96 percent, on completing one or more years of college becomes 1.93 percent, and on years of schooling becomes 0.08 percent. If we express the increase in schooling in terms of years, then the increase is 0.01 years (((0.0295*(1000/384))/10)/0.75).

Micheltmore (2013)

Micheltmore (2013) found that a \$1,000 increase in the maximum state EITC increased the likelihood of being enrolled in college by 0.015 (s.e. 0.012), increased the likelihood of ever being enrolled in college by 0.027 (0.012), increased years of educational attainment by 0.107 years (s.e. 0.051) or 0.89 percent (0.107 years out of an average of 11.97 years), and increased the probability of high school completion by 0.023 (s.e. 0.012) or 3.29 percent (2.3 percentage points out of an average of 70 percent). Data was derived from the Survey of Income and Program Participation (SIPP), pooling panels from 1990 through 2008. Data when individuals were 18-23 years old was used for this analysis, with parental educational attainment as a proxy for EITC eligibility. Participants living with parents who had no schooling beyond a high school

degree were considered EITC-eligible (n=25,337). The study employed a difference-in-differences analysis with variation in treatment dosage to determine the impact of state EITCs on educational attainment. Analyses examined the impact of within state EITC expansions on educational attainment relative to changes in outcomes among individuals in untreated states. Models controlled for demographic characteristics and year and state fixed effects.

To remain consistent with other literature, we focus on results for high school completion and years of education. The author does not describe the change in dollars received associated with a \$1,000 increase in the maximum EITC. We assume the increase is \$384 in 2019 dollars, as found by Maxfield (2015). This results in a 8.57 percent ($0.0329 \times (1000/384)$) increase in the likelihood of completing high school and a 2.32 percent ($0.0089 \times (1000/384)$) increase in years of education. Individuals were exposed to the EITC for between 7 and 18 years (individuals in the sample were younger than 12 years old). We take the midpoint of the range of exposure, 12.5 years. Thus, a one-year, \$1,000 increase in EITC benefits increased the likelihood of completing high school by 0.69 percent (8.57 percent/12.5) and increased years of education by 0.19 percent (2.32 percent/12.5). Next, we adjust results to apply only to the 75 percent of EITC-eligible households that are “treated” through actual receipt of an income transfer (U.S. Government Accountability Office, 2001). We find that per \$1,000 cash transfer the probability of completing high school increased by 0.91 percent and increased years of education by 0.25 percent. If we express the increase in education in terms of years, then the increase is 0.03 years ($((0.107 \times (1000/384))/12.5)/0.75$).

Hoynes et al. (2016)

Hoynes et al. (2016) found that exposure to food stamps in early childhood increased the probability of receiving more than a high school education by 0.184 standard deviations (s.e. 0.108). We do not use this result to measure the impact of a \$1,000 transfer on educational attainment because the outcome differs slightly from remaining literature and we are unable to convert results presented in z-scores to percentage terms due to absence of the standard deviation of the mean.

Aizer et al. (2016)

Aizer et al. (2016) found that in adulthood, sons whose mothers had received Mothers' Pensions experienced an increase in years of schooling of 0.316 years (s.e. 0.262), a 3.4 percent increase.

As previously stated, the \$20 monthly transfer in 1922 would be worth \$307 today, or \$3,684 annually for on average three years. Therefore, a \$1000 transfer for one year would increase years of education by 0.31 percent [$0.034 \times ((1000/3684)/3)$]. If we express the increase in education in terms of years, then the increase is 0.03 years ($((0.316 \times (1000/3684))/3)$).

Bastian and Micheltmore (2018)

Bastian and Micheltmore (2018) found that an exposure of \$1,097 in EITC benefit (2019 dollars): prior to age five, decreased the probability of graduating high school by 0.005 (s.e. 0.005) or -0.5 percent and decreased educational attainment by 0.024 years (s.e. 0.071) or -0.18 percent (based on an average of 13.7 years); between the ages of 6 and 12, decreased the probability of graduating high school by 0.003 (s.e. 0.003) or -0.3 percent, and increased schooling by 0.008 years (s.e. 0.022) or 0.06 percent (based on an average of 13.7 years); and between the ages of 13 and 18, increased the probability of graduating high school by 0.012 (s.e. 0.003) or 1.2

percent and increased schooling by 0.081 years (s.e. 0.025) or 0.59 percent (based on an average of 13.7 years).

To simplify calculations, we first calculate an average impact across all ages by multiplying each of Bastian and Micheltore's estimates for the three age groups times the proportion of children currently in that age group. According to Bastian and Micheltore (2018), children exposed to EITC from ages 0-5, from ages 6-12 and from ages 13-18 make up 21.6%, 40.4% and 38% of their samples, respectively. This results in an increase in the probability of graduating high school by 0.23 percent and increase the average years of schooling by 0.21 percent. However, these results are for multiple years of exposure to EITC and include all children in states in which the maximum EITC increased, not just recipient children. We assume the child was exposed to the EITC in all 18 years, yielding a 0.01 percent increase in the probability of graduating high school and a 0.01 percent increase in the average years of schooling. To convert this intent-to-treat estimate to an estimate of the effects on the treated, we divide each estimate by 83 percent (Scholz 1994), which was the EITC participation rate in 1990 (the middle of the study period). It results in a 0.02 percent increase in the probability of graduating high school and a 0.01 percent increase in schooling for a \$1000 transfer. Finally, to obtain the impact of a \$1,000 increase, we multiply the estimate by (1000/1097), resulting in a 0.01 percent increase in the probability of graduating high school and a 0.01 percent increase in years of schooling. If we express the increase in schooling in terms of years, then the increase is 0.002 years, calculated as the following: $((((-0.024 * 0.216) + (0.008 * 0.404) + (0.081 * 0.38)) / 18) / 0.83) * (1000 / 1097)$.

Thompson (2019)

Thompson (2019) found that exposure to an average-sized casino operation over the entirety of childhood increased the probability of receiving an associate's degree by 0.057 (s.e. 0.027) or 5.7 percentage points, increased the probability of receiving a bachelor's degree by 0.010 (s.e. 0.009) or 1 percentage point, increased educational attainment by 0.328 years (s.e. 0.070), and increased the probability of high school graduation by 0.041 (s.e. 0.011) or 4.1 percentage points. The author examined educational attainment among self-identified American Indians in 36 counties where a tribal casino was opened during respondents' childhood. Analyses were conducted using the 2000 Decennial Census IPUMS samples and American Communities Survey (ACS). A difference-in-differences framework was used to compare the educational attainment of American Indian individuals from the same county with differing levels of exposure to tribal casino payments. The within-county differences were then compared to determine whether results differed based on the size of the county's casino operations. The sample was limited to self-identified American Indian individuals from counties that opened a casino between 1987 and 2004. The author was able to identify respondents' current county of residence (during adulthood) and state of birth but was unable to identify what county the individual resided in throughout childhood. As selective migration might bias findings, the author limited the sample to individuals currently living in a county in the same state as they were born (n=11,647). Casino exposure was measured by dividing the number of slot machines operated by the American Indian casino in the county by the size of the American Indian population per county and year. The casino exposure measure was then scaled so that a one-unit increase corresponded to an individual spending their full childhood in a county with an average-sized gaming operation and American Indian population. Outcomes were measured between the ages of 22 and 40. Analyses controlled for county and cohort fixed effects, age at time of survey, tribal identity, and gender.

Transfer income increased by \$304.9 (s.e. 47.1) in the average American Indian household. Hourly wage increased by \$1.56 (s.e. 0.093) and unemployment decreased by 0.020 (s.e. 0.004) or 2 percentage points. Results indicate that increased educational attainment of children was likely a result of both improved labor market opportunities and transfer payments for the family. Of the total \$3,548 increase in income among American Indian families, 8.6 percent (\$305 in 2016 dollars or \$325 in 2019 dollars) was a result of transfer income. We assume the same proportion of increased educational attainment was a result of transfer income. Therefore, the income transfer increased educational attainment by 0.03 (0.086×0.328) years or 0.23 percent (out of the control group mean of 12.33 years) and increased the probability of high school graduation by 0.35 (0.086×0.041) percentage points or 0.46 percent (out of the control group mean of 76 percent). A \$1,000 transfer would then increase educational attainment by 0.7 percent ($0.0023 \times (1000/325)$) and the probability of high school graduation by 1.43 percent ($0.0046 \times (1000/325)$). Results represent the impact of exposure to the transfer for 18 years, so we adjust results for years of exposure. We find that a \$1,000 cash transfer increased educational attainment by 0.04 percent ($0.007/18$) and the probability of high school graduation by 0.08% ($0.0143/18$). If we express the increase in educational attainment in terms of years, then the increase is 0.005 years ($((0.328 \times 0.086) \times (1000/325))/18$).

Child Welfare

Berger et al. (2017)

Berger et al. (2017) found that \$1,000 in potential EITC is associated with a decrease in the probability of neglecting a child, a decrease in the probability of abusing a child, and a 0.0027 (s.e. 0.0038) or 0.27 percentage-point decrease in the probability of being investigated by Child Protective Services (CPS). Using the Fragile Families and Child Wellbeing Study (4,040 family-wave observations), the authors use an instrumental variable strategy to examine the effect of income on child maltreatment. Income is instrumented using state and national variation in EITC generosity. The sample is limited to unmarried families with AGIs of no more than \$45,000 per year (nominal dollars). EITC generosity is measured using TAXSIM based on year, income, filing status, number of dependents, and state of residence. F-statistic results indicated the EITC was a strong predictor of post-tax income. Child maltreatment is measured using mothers' self-reports of having been investigated by CPS, indicators of physical violence, and indicators of neglect. Analyses control for race/ethnicity, maternal education, number of biological children in household, family structure, age of youngest child, mother's age, whether the mother reported no household income, average of lagged and current household income, census tract unemployment rate, census tract public assistance receipt rates, wave of observation, and state of observation.

The authors find that a \$1,000 increase in potential EITC benefits is associated with a \$936 to \$1,030 increase in income in 2009 dollars; We use the average of this range of earnings, \$983, which is the equivalent of \$1,167 in 2019 dollars. Thus, the impact on the probability of CPS investigation becomes 0.23 percentage points ($0.27 \times (1000/1167)$). Fang et al. (2012), based on federal, state, and local expenditures on child welfare activities (CPS investigations and foster care) and the number of CPS-involved children, estimate that the average per-year cost per investigated child is \$7,728 (2010 dollars), the equivalent of \$9,082 (2019 dollars). Therefore, we estimate that a \$1,000 transfer is associated with \$21 (9082×0.0023) in decreased spending on child welfare investigations. Berger et al. (2017)'s sample is limited to unmarried mothers.

However, based on correspondence with Berger, who reported finding similar, but much less precisely estimated, effects for married mothers, we assume the impact of a \$1,000 transfer on CPS involvement does not differ among married mothers. We conclude that a \$1,000 transfer decreased the present discounted value of expenditures on child welfare by \$37.

Rittenhouse (2022)

Authors found that being eligible for larger child-related tax benefits during infancy led to a 0.001 (s.e. 0.00149) or 0.1 percentage-point decrease in having any referrals to Child Protective Services (CPS), a 0.000897 (s.e. 0.00140) or 0.0897 percentage-point decrease in having any CPS investigations through age 2, a 0.000775 (s.e. 0.000569) or 0.0775 percentage point decrease in foster care placement through age 2, a 0.00729 (s.e. 0.00344) decrease in the number of referrals to CPS through age 2, a 0.00658 (s.e. 0.00263) decrease in the number of CPS investigations through age 2, and a 0.612 (s.e. 0.290) decrease in days spent in foster care through age 2. Effects are larger for low-income households. Eligibility of larger benefits led to a 0.00579 (s.e. 0.00301) or 0.579 percentage-point decrease in having any referrals to CPS, a 0.00547 (s.e. 0.00286) or 0.547 percentage-point decrease in having any CPS investigations through age 2, a 0.00255 (s.e. 0.00124) or 0.255 percentage-point decrease in foster care placement, a 0.0190 (s.e. 0.00731) decrease in the number of CPS referrals, a 0.0171 (s.e. 0.00561) decrease in the number of CPS investigations, and a 1.880 (s.e. 0.646) decrease in the number of days spent in foster care. Authors used data from the Children's Data Network, which housed data on birth records, death records, and CPS records in California. The sample included children born within 60 days of January 1st to first-time mothers between November 1999-March 2017 (n=1,181,675). Low-income households were defined as those whose predicted incomes were below 200% of the federal poverty line. To estimate the causal impact of cash transfer, authors used a regression-discontinuity design, where children born in December (treatment group) are eligible for tax benefits in the following year when they are age 0-1 but children born in January (comparison group) are not. However, the treated children would also lose tax benefits one year earlier than the comparison children. Other controls in the model included re-centered birth year fixed effect.

We use the results for low-income children because they are more likely to be eligible for EITC. We use the results on the probability of having CPS investigations as it is also examined by the other child welfare literature we use—Berger et al., 2017. Using ACS data, authors estimated that among low-income household, children born in December received \$2,881 (\$2017) more child-related tax benefits during the first year of life than children born the next month in January, the equivalent of \$3,012.59 in 2019 dollars. However, the treated children would also lose tax benefits one year earlier than the control children. Receiving \$3,012.59 at age 0 is more valuable than receiving \$3,012.59 18 years later at age 17, which only has a present discounted value of \$2,151 using a 2% discount rate. We contribute the effects discovered by the paper to the difference between \$3,012.59 and \$2,151, around \$861. If \$861 led to a 0.547 percentage-point decrease in the likelihood of having CPS investigations, then a \$1,000 benefit would lead to a 0.64 (0.547*1000/861) percentage-point decrease in the likelihood of having CPS investigations. The effect discovered by the paper was an intent-to-treat effect. We further divide the 0.64 percentage-point decrease by 0.7945, which was the estimated EITC participation rate in 2008 (the middle of the study period) according to Jones (2014). This adjustment yields a 0.81 percentage-point decrease in CPS investigation. According to Fang et al. (2012), the average per-year cost per investigated child is \$7,728 (2010 dollars), the equivalent of \$9,082 (2019 dollars). A 0.81 percentage-point decrease in CPS investigation is

thus worth \$73. We conclude that a \$1,000 increase in household income from cash transfers would bring \$73 worth of benefits in reduced expenditures on child welfare per year.

Avoided Expenditures and Victim Costs of Crime

In order to calculate avoided expenditures and victim costs of crime from cash transfers, we need to know: 1) the average monetary cost per crime, 2) lifetime distribution of criminal activities, 3) the impact of cash transfers on crime. In the section below, we discuss the evidence we have collected on these three components and how we use them for the calculation. The calculation below is documented in Garfinkel et al., (2024), published on the website of Columbia University Center on Poverty and Social Policy. It is based on Garfinkel et al., (2022), published in the *Journal of Benefit and Cost Analysis*, but improves it further by using more accurate estimates of cost per crime and the age distribution of crime and also including new quasi-experimental evidence on the impact of cash transfers on crime.

Quasi-experimental research indicates that cash transfers to children, especially children in poverty, reduce crime, including both property crimes and violent crimes. We use the result of quasi-experimental research on crime reduction per year and studies that examine the age distribution of crime to calculate reduction in crimes throughout children's lifetimes. We monetize the lifetime decrease in crime using standard literature estimate on the cost of crime.

Cost per crime

Cost of crime includes both victim cost and criminal legal system costs (ex: police, incarceration, court). We use Cohen (2020)'s estimate on total cost per crime (minus the cost of lost productivity of criminals to avoid double counting as we are already counting future earnings increases of children).

Because incarceration only applies to people ages 18 and above, for cost per crime committed before age 18, we need to further subtract incarceration cost from the total cost. Since Cohen (2020) did not provide an estimate on incarceration cost per crime, we estimate incarceration cost by calculating what percentage of criminal legal system cost is incarceration cost and what percentage of total cost is criminal legal system cost. We first estimate the percentage of criminal legal system cost that is incarceration cost. According to Table 4.1 of Cohen (2020), in 2015, incarceration cost per capita is \$261 and criminal legal system cost per capita is \$855. Incarceration cost is thus 30.5% ($261/855$) of criminal legal system cost. We then estimate the percentage of crime cost that is the criminal legal system cost. Table 5 of Miller et al. (2021) presents that per violent crime, criminal legal system cost is \$5,529 ($\$2328 + \3201), and total cost minus perpetrator work loss is \$90,401. Criminal legal system cost is thus 6.1% of total cost for violent crime. Per non-violent crime, criminal legal system cost is \$707 ($\$274 + \433), and total cost minus perpetrator work loss is \$2,250. Criminal legal system cost is thus 31.4% of the total cost for non-violent crime. According to Table 4 of Miller et al. (2021), there are 24,117,831 violent crimes and 120,999,583 total crimes, suggesting that there are 96,881,752 non-violent crimes. Weighting 6.1% and 31.4% by the percentage of total crimes that are violent and non-violent, we conclude that for all crimes, criminal legal system cost is 26.4% of the total cost of crime. Since total cost of crime includes criminal legal system cost and victim cost, our calculation based on Miller et al. (2021) suggests that victim cost is 73.6% of the total cost of all crimes.¹⁴

¹⁴ An alternative calculation based on Cohen (2020) suggests that criminal legal system cost is 10.1% of total cost of crime. Table 6.2 of Cohen (2020) presents that for all crimes committed in the United States in 2017, criminal legal system cost is worth a total of \$211,764 million. Total cost minus perpetrator work loss is worth \$2,094,702 million. Criminal legal system cost is thus 10.1% of total cost of crime. The Cohen estimate on the percentage of total cost that is criminal legal system cost would suggest much higher social benefits for reducing crime, so we cautiously rely

Our calculations suggest cost per murder is \$8,158,816 in 2019 dollars (\$8,006,490 pre-age-18), cost per robbery is \$29,070 (\$28,527 pre-age-18), and cost per assault is \$41,224 (\$40,454 pre-age-18). We use the unweighted average of Cohen's estimate on rape and on other sexual assault for cost per rape- \$119,001 (\$116,779 pre-age 18). Cohen doesn't have an estimate for property crime. The FBI considers burglary, larceny, motor-vehicle theft, and arson as property crime. Our calculations suggest cost per burglary is \$2,887, per larceny is \$4,344, per motor vehicle theft is \$8,499 and per arson is \$35,245.

In order to calculate the cost per property crime, we need to know what percentage of property crimes are burglary, larceny, motor vehicle theft, and arson. According to the FBI (2019), there are 1,245,410 violent crimes (including number of rapes under revised definition) and 6,959,072 property crimes (including arson, which FBI imperfectly estimates to be around 33,395). Among violent crimes, 16,425 (1.32%) are murder and manslaughter, 139,815 are rape (11.23%), 267,988 (21.52%) are robbery, and 821,182 (65.94%) are aggravated assault. Numbers of simple assaults are not reported and thus not included in the calculation. Among property crimes, 1,117,696 (16.06%) are burglary, 5,086,096 (73.09%) are larceny-theft, 721,885 (10.37%) are motor vehicle theft, and 33,395 (0.48%) are arson. FBI likely underestimates the true level of crime because not all crimes are reported to the police. We thus obtain data on the percentage of victimizations that are reported to the police from self-reported victimization data. According to Table 5 of the report of Department of Justice, in 2021, among violent crimes, only 25% ($0.3/(0.3+0.9)$) of rape victimizations are reported to the police, followed by 58.82% ($1/(1+0.7)$) of robbery and 62.96% ($1.7/(1.7+1)$) of aggravated assault. Data on murder is not available. We assume that 100% of murder are reported to the police. 31.20% ($27.8/(27.8+61.3)$) of property victimizations are reported to the police. Among property crimes, only 41.30% ($5.7/(5.7+8.1)$) of burglary are reported to the police, 76.74% ($3.3/(3.3+1)$) of motor vehicle theft are reported to the police. Data on larceny and arson is not available. We assume that larceny has the same report rate of other theft- 26.48% ($18.8/(18.8+52.2)$). We assume that all arson is reported to the police. We combine two data sources to estimate the true level of crime. We assume that the percentage of victimizations reported to the police stay the same from 2019 to 2021. In 2019, there should be 16,425 ($16425/1$) murder, 559,260 rape ($139815/0.25$), 455,580 robbery ($267988/0.5882$) and 1,304,230 ($821182/0.6296$) aggravated assault. Thus, within the true level of violent crime, 0.7% are murder, 23.95% are rape, 19.51% are robbery, and 55.84% are aggravated assault. There should be 2,706,000 burglary ($1117696/0.413$), 19,208,129 larceny-theft ($5086096/0.2648$), 940,638 motor vehicle theft ($721855/0.7674$) and 33,395 arson ($33395/1$). Thus, within the true level of property crime, 11.82% are burglary, 83.92% are larceny, 4.11% are motor vehicle theft and 0.15% are arson. Using numbers we have calculated, we conclude that the cost per property crime is \$4,388 (\$3,967 pre-age-18).

Age-crime Relationship

We use the age-crime relationship discovered by Schulman et al. (2013). Authors found that the proportion of youth engaging in any type of crime peaks in adolescence (ages 15-16) and decreases as youth enter adulthood. The authors used self-reported crime data from waves 1-7 of the National Longitudinal Survey of Youth, 1997 Cohort (NLSY97) survey. Two measures of criminal behaviors were created: offending and index offending. Offending was constructed based on categories for whether the respondent had committed assault, property damage, other

on the more moderate results based on Miller et al. (2021), which would give us smaller estimates on cost per-crime before age 18.

property crime, theft below \$50, theft above \$50, and the selling of drugs in the past year. Given that queries on these crimes could sometimes confuse serious offenses with minor ones, authors constructed another measure on index offending, replacing responses on theft with responses on the details of theft. Index offending was constructed for categories of whether the respondent had committed assault, shoplifting above \$50, the stealing of a purse or wallet, stealing of things from a locked building, stealing of cars and other motor vehicles, stealing of things using a weapon, and the selling of drugs in the past year. The authors first analyzed the age pattern of offending and index offending with descriptive statistics, then through structural equation modeling.

We use figure 1 of Schulman et al. (2013), which presents the proportion of NLSY97 youth committing any type of offenses or index offenses from ages 12-22. We focus on the distribution of index offenses since this measure looks at more serious crimes and is less likely to confuse trivial offenses with serious ones. To estimate the proportion of youth committing index offenses beyond age 22, we assume that the proportion is half of the proportion at age 22 from ages 23-44, a quarter of the proportion at age 22 from ages 45-64, and zero from ages 65 and beyond. We approximate the age-crime relationship of violent crime using the age-crime relationship of assault. Since we calculate that within violent crimes, 0.7% are murder, 23.95% are rape, 19.51% are robbery, and 55.84% are aggravated assault (see the previous section for the calculation), we attribute 0.7%, 23.95%, 19.51% and 55.84% of the age-crime relationship of violent crime to murder, rape, robbery, and assault. For instance, if figure 1 shows that 9% of youth aged 12 commit violent crimes, given that 0.7% of violent crimes are murders, we estimate that 0.6% ($9\% \times 0.7\%$) of youth aged 12 commit murders. We assume that the proportion of youth committing property crime is the sum of the proportion of youth committing all crimes in figure 1 except assault.

Bailey et al. (2020)

Bailey et al. (2020) find that exposure to food stamps at age five or younger decreased the probability of being incarcerated by 0.0008 (s.e. 0.0004) or 0.08 percentage points. Based on data from the 2001-2013 American Community Survey matched with the 2000 Census Long Form ($n=7,705,000$), the authors use a difference-in-difference framework exploiting the county-by-county introduction of food stamps. Models control for county of birth, birth year, and birth state fixed effects as well as 1960 county-level characteristics interacted with a linear birth-cohort trend.

We use Bailey et al. (2020)'s results to measure the impact of a \$1,000 increase in household income on the present discounted value of crime. To translate their estimate of the intent to treat to an estimate of the treatment on the treated, we divide 0.08 percentage points by the percentage of children in this age group who received food stamps, 16 percent. Thus, the treatment-on-the-treated outcome is a 0.5 percentage-point increase for a cumulative exposure of 7 years, or a 0.07 percentage-point decrease in the probability of being incarcerated per year. The average annual food stamps value per person in 1972 (near the midpoint of the study period) was \$994 per year in 2019 dollars (Department of Agriculture, 2021). Assuming average households have three individuals, the total household food stamps value would be \$2,982, on average. The impact of a \$1,000 benefit in 2019 dollars is thus 0.07 percentage points times the ratio of $\$1000/2982$, or 0.024 percentage points. Since the paper studies the impact of exposure from conception to age 5, to avoid overstating long-run benefits we assume that child recipients were exposed to food stamps through the entirety of childhood (in utero through age

17) but only derived benefits for future earnings during the first 7 years of payments. We multiply results by the 7/19 of years in which they derive benefits, decreasing the impact to a 0.009 percentage-point decrease in the probability of being incarcerated. We conclude that a \$1,000 increase in household income from cash transfers per year would decrease the chance of incarceration by 0.0088 percentage points.

We calculate reduction in costs of crimes using two methods. In the first method, we start with the standardized impact on incarceration and convert it into an impact on the level of crime. We first convert it into an impact on arrests by dividing it by the arrest-incarceration ratio estimated by the Vera Institute (2019): 0.99 incarcerations per arrest.¹⁵ The result is a decrease on arrest probability of 0.0089 percentage points (0.0088/0.99). To be consistent with our crime calculation based on Barr & Smith (2023), we decompose the impact on arrests into impacts on arrest of a specific type of crime. Based on statistics from the FBI (2019),¹⁶ we calculate that 0.2% of all crimes are murder or manslaughter, 1.7% rape, 3.3% robbery, 10% aggravated assault, and 84.8% property crimes. We thus distribute 0.2%, 1.7%, 3.3%, 10%, and 84.8% of the 0.0089 percentage-point reductions in arrests to reductions in arrest of murder (0.00002 percentage points), rape (0.0002 percentage points), robbery (0.0003 percentage points), aggravated assault (0.0009 percentage points), and property crimes (0.008 percentage points). Not all crimes lead to arrest and not all crimes are reported to the police. For each type of crime, we further divide the impact on arrests by the percentage of that type of crime that lead to arrest (FBI 2019)¹⁷ and percentage of that type of crime reported to the police (Department of Justice 2022)¹⁸ to arrive at the impact on the level of crime. Having adjusted for the percentage of crime that leads to arrests and is reported, we conclude that there would be a 0.00003 percentage-point reduction in murder, a 0.002 percentage-point reduction in rape, a 0.002 percentage-point reduction in robbery, a 0.003 percentage point reduction in aggravated assault, and a 0.14 percentage-point reduction in property crime. We multiply these percentage-point decrease of crime by the cost of crime calculated above to get the dollar value of reduction in crimes. To calculate the present discounted value, we multiply the dollar value by the distribution of crime from ages 0-78 and discount the benefit with a discount rate of 2%. We conclude that following a \$1,000 increase in household income from cash transfer, the present discounted value of reduced costs of crime over the lifetime (ages 0-78) is \$6.

In the second method, we start with the standardized impact on incarceration. We then follow Bailey et al.'s method to monetize such impact. We multiply the standardized impact by the average length of incarceration (2.6 years according to Bailey et al.) and by the cost of incarceration (\$33,985 in 2019 dollars according to Bailey et al.) and yield a result of \$7.8

¹⁵ This ratio may seem high, but using it has the virtue of giving us an estimate of smaller, more conservative magnitude

¹⁶ According to the FBI (2019), there are 1,245,410 violent crimes and 6,959,072 property crimes, including arson. Among these crimes, 16,425 (0.2%) are murder and manslaughter, 139,815 rape (1.7%), 267,988 (3.3%) robbery, 821,182 (10%) aggravated assault, and 6,959,072(84.8%) property crimes. Among violent crimes, 11.23% (139815/1245410) are rape.

¹⁷ According to the FBI (2019), 61.4 percent of murder offenses, 52.3 percent of aggravated assault offenses, 30.5 percent of robbery offenses, 32.9 percent of rape offenses, and 17.2 percent of property crimes were cleared by arrest or exceptional means.

¹⁸ FBI statistics are likely underestimates because not all crimes are reported to the police. According to the Department of Justice, in 2021, 27.8 out of 89.1 property victimizations are reported to the police (31%), 1 out of 1.7 robberies are reported to the police (59%), 0.3 out of 1.2 rapes are reported to the police (25%), and 1.7 out of 2.7 aggravated assaults are reported to the police (63%). We assume that 100% of murders are reported to the police.

($0.000088 \times 2.6 \times 33985$). Bailey et al.'s sample for incarceration ranges from 22-54 years old. For simplicity, we assume that the \$7.8 reduction in cost of incarceration takes place at age 38 (the midpoint of the age range) and that average child beneficiary is 9 years old. The present discounted value of reduction in incarceration cost is thus \$3.3 ($\$7.8 / (1.03)^{29}$). Reduction in the cost of incarceration is only part of the reduction in the cost of crime. As calculated in the previous section, incarceration cost is 30.5% of criminal legal system cost. We thus divide \$3.3 by 30.5% to estimate the reduction in criminal legal system cost, a total of \$10.8. As calculated in the previous section, victim cost is 74% of the total cost of crime and criminal legal system cost is 26% of the total cost of crime. We divide \$10.8 by 26% to estimate reduction in the total cost of crime and arrive at \$41. We multiply \$41 of reduction in cost of crime by the distribution of crime from ages 0-78 and discount the benefit with a discount rate of 2%. We conclude that following a \$1,000 increase in household income from cash transfer, the present discounted value of reduced costs of crime over the lifetime (ages 0-78) is \$104.

We use the unweighted average of the two present discounted value, \$55, as the final result calculated from Bailey et al. (2021). We conclude that following a \$1,000 increase in household income from cash transfer, the present discounted value of reduced costs of crime over the lifetime (ages 0-78) is \$55.

Barr & Smith (2023)

Authors found that being exposed to Food Stamps in utero through age 5 reduces the probability of any criminal conviction by age 24 by 0.013 (s.e. 0.007), or 1.3 percentage points, reduces probability of violent-crime conviction by 0.005 (s.e. 0.002) or 0.5 percentage points, and reduces the probability of property-crime conviction by 0.003 (s.e. 0.003) or 0.3 percentage points. Being exposed to Food Stamps between ages 0-5 reduces the arrest rate of violent crime between ages 18-24 by 0.151 (s.e. 0.048) or 15.1 percent, and reduces the arrest rate of property crime by 0.128 (s.e. 0.091) or 12.8 percent. Within violent-crime, it reduces arrest rate of murder by 0.032 (s.e. 0.014) or 3.2 percent, rate of aggravated assault by 0.064 (s.e. 0.030) or 6.4 percent, and rate of robbery by 0.042 (s.e. 0.014) or 4.2 percent.

Authors used administrative data from North Carolina on convictions, nationally representative Uniform Crime Reporting (UCR) data on arrests, and linked them with information on Food Stamps availability within a county and month for various birth-month cohorts. The North Carolina data covered all individuals convicted in North Carolina from 1972-2015. UCR data covered individuals arrested in a county (more than counties in North Carolina) and year. The sample on convictions is restricted to those born between 1964-1974 and includes 13,173 observations. The sample on arrests is restricted to those aged 18-24 and the number of observations vary from 30,453 to 96,386 depending on the type of crime. Regressions were conducted via ordinary least squares, exploiting within-county differences in the availability of Food Stamps in the 1960s and 1970s. The dependent variable was the crime rate of individuals born in a certain county and birth-cohort. The main independent variable was Food Stamps exposure in that county and birth-cohort. Other variables included birth county fixed effects, birth cohort fixed effects and interactions between pre-treatment county characteristics and time trends.

We first calculate the per year impact of a \$1,000 increase in household income from Food Stamps on convictions. The average conviction rate for any type of crime is 9 percent. Thus, the 1.3 percentage-point decrease in any conviction is a 14.4% decrease ($1.3/9$). Children in the sample are exposed to Food Stamps for 5.75 years (0.75 year for the nine months in utero, and 5 years between ages 0-5), thus, the per year decrease in any crime conviction is 2.5% ($14.4\%/5.75$). The average annual food stamps value per person in 1972 (near the midpoint of the study period) was \$994 in 2019 dollars (Department of Agriculture, 2021); assuming average households have three individuals, the total household food stamps value would be \$2,982, on average. A 1,000 increase in Food Stamps would thus cause any crime conviction to decrease by 0.84% ($2.5\% * 1000/2982$). As the paper studies the crime impact of exposure in utero through age 5 (a total of 5.75 years according to the authors), we cautiously assume that child recipient exposure to food stamps is spread over the entirety of childhood (from age -1 to age 17) but only derived benefits for future crime reduction during the first 5.75 years of payments. To measure the impact per year of payments, we multiply results by the 5.75/19 of years in which they are assumed to derive benefits, decreasing the impact to 0.25% ($0.84\% * 5.75/19$). Finally, we adjust for an estimate of the Food Stamps participation rate to obtain the treatment-on-the-treated effect. Using a participation rate of 16%, the treatment-on-the-treated effect is 1.59%. We thus conclude that a one-year increase of household income of \$1,000 from the in-kind value of Food Stamps reduces crime conviction by age 24 by 1.59%.

Through the same standardization process, we conclude that a one-year increase of \$1,000 from Food stamps reduces violent-crime conviction by 3.68% and reduces property-crime conviction by 0.14%. It reduces the arrest rates of violent crime by 1.67%, property crime by 1.4%, murder by 0.35%, aggravated assault by 0.7%, and robbery by 0.46%. The paper does not provide any estimate on rape. Our calculation based on FBI statistics and statistics from the Department of Justice suggests that 11.23% of violent crimes are rape (for calculation details, see footnote 5). Given the 1.67% reduction in the arrest rate of violent crime we have calculated, we infer that the arrest rate of rape decreased by 0.19%.

To calculate reduction in the costs of crime, we follow the first method of Bailey et al. (2021) converting the impacts on arrests into impacts on levels of crime. To make the conversion, for each type of crime, we divide the impact on arrest by the percentage of that type of crime cleared by arrest, and by the percentage of that type of crime reported to the police. For instance, according to the FBI (2019), 30.5% of robberies lead to arrest and, according to the Department of Justice, 59% of robbery victimizations are reported to the police. We thus divide 0.46% reduction in robbery arrests by 30.5% and again by 58.8% to obtain a reduction in robbery of 2.58%. Then we multiply by cost per crime estimated from Cohen (2020) and by the age-crime relationship from Schulman et al. (2013) to calculate the present discounted value. We conclude that following a \$1,000 increase in household income from cash transfer, the present discounted value of reduced costs of crime over the lifetime (ages 0-78) is \$2,808.

Other Transfers

As there were not published findings appropriate to our question, we conducted our own analyses on the relationship between earnings and other transfers using the 2014 Survey of Income and Program Participation. Respondents were interviewed annually between 2013-2016. Transfers are measured as the sum of annual EITC, housing subsidies, disability, workers compensation, WIC, unemployment compensation, TANF, SSI, general assistance, and food stamps. SIPP respondents report the amount received for each transfer excluding the EITC and housing subsidies. EITC transfers are estimated using the NBER's Taxsim. Respondents report receipt of housing subsidies but not the amount received. The amount of housing subsidies received is estimated based on the difference between the amount of rent paid and the fair market rent in the state (estimated separately for urban and rural areas) for the corresponding household size. The sample includes individuals between the ages of 18 and 64. The sample is limited to individuals with a high school education or less, a proxy for eligibility for mean-tested programs. Linear regressions were conducted controlling for race, the number of children in the household, and marital status ($n=634,678$). Further analyses were conducted expanding the sample to include individuals with less than a college degree; results, as expected, are smaller, increasing our confidence in the findings (Regression results presented in table A4.1 and A4.2 below).

In children's earnings section in the main text, we find that a \$1,000 cash transfer during childhood increases earnings in adulthood by -\$33, \$25, \$62, \$127, and \$249 per year. We estimate that \$1,000 in earnings reduces transfers by \$13.61. Therefore, we find that that the corresponding decrease in transfers is $-\$0.45$ ($13.61*(-33/1000)$), $\$0.34$ ($13.61*(25/1000)$), $\$0.84$ ($13.61*(62/1000)$), $\$1.72$ ($13.61*(127/1000)$), and $\$3.39$ ($13.61*(249/1000)$). We estimate the present discounted value of the decrease in other transfers, assuming to begin at age 22, and end at age 64. The average child beneficiary is assumed to be age 9. Using the mean estimate, we conclude that the present discounted value of decreased transfers is \$26 in adulthood as a result of a \$1,000 cash transfer during childhood.

Increased Payments Due to Increased Children's and Adult's Longevity

With the increased children's longevity comes a cost. Two major components of the cost are Social Security and Medicare payments. According to the Social Security Administration (2019), retired workers received an average of \$1,461 in Social Security per month in 2018. This means that the annual Social Security payment was \$17,532 in 2018, the equivalent of \$17,821 in 2019 dollars. According to the Kaiser Family Foundation, Medicare spending per enrollee was \$10,536 in 2019. We thus assume that one year of increase in longevity requires \$28,357 of payments from Social Security and Medicare. To estimate the total increase in Social Security and Medicare payments, we turn to our previous estimates on longevity.

Our estimates indicate that a \$1,000 increase in cash transfer for one year would increase children's longevity by 0.0194 (Bailey et al., 2020) or 0.105 years (Aizer et al., 2016). A 0.0194-year increase in longevity would thus require \$551 ($0.0194*28357$) more Social Security and Medicare payments. Since we assume that the extension of longevity occurs at age 78, we assume that payments are made to children at age 78 as well. The present discounted value of increased payments is \$140. A 0.105-year increase in longevity would require \$2976 ($0.105*28357$) more Social Security and Medicare payments. The present discounted value is \$759. Using the mean of these two present discounted values, we conclude that as a result of

the impact of a \$1,000 cash transfer on children's longevity, there would be a \$450 increase in Social Security and Medicare payments to children once in adulthood.

For adults' longevity, the average increase in longevity based on three studies is 0.0089-year. A 0.0089-year increased longevity would require \$253 in increased payments (0.0089×28357). The present discounted value is \$114. We conclude that due to the impact of the \$1,000 cash transfer on adults' longevity, there would be a \$114 increase in Social Security and Medicare payments made to parents.

Increased Costs Due to Increased Education of Children

Increased schooling poses direct costs to child beneficiaries in the form of tuition and fees and to taxpayers in the form of tax payments used to support national and local educational systems. Our estimates on increased schooling suggest that a \$1,000 dollar increase in household income from a child allowance for one year would increase years of schooling by 0.0018-0.0297 years. Since for most of our impact studies, an average child in the sample has 12 years of education, we regard the 0.0018-0.0297 increase as an increase in postsecondary education. We use data provided by Abel and Deitz (2014) to calculate the increased direct costs. The study estimated that for a 4-year bachelor degree, the price charged for one year was \$14,750 but students paid only \$6550, with \$8,200 offset by grants, scholarships and tax benefits to students. For a 2-year associate degree, the price charged for one year was approximately \$3,000, but was completely offset by grants, scholarships and tax benefits to students that summed up to \$4,300, implying that students gained \$1,300 in tuition and fees. Taking an average of \$6,550 and -\$1,300, the average direct costs to child beneficiaries in the form of tuition and fees are worth \$2,625, the equivalent of \$2,880 in 2019 dollars. Taking an average of \$8,200 and \$4,300, the direct costs to taxpayers in the form of tax payments used for grants and scholarship are worth \$6,250, the equivalent of \$6,856 in 2019 dollars. Multiplying 0.0018-0.0297 years of schooling by \$2,880 yields an increase in yearly cost for child beneficiaries of \$5-\$86. Multiplying 0.0018-0.0297 years of schooling by \$6,856 yields an increase in yearly taxpayers cost of \$12-\$204. Assuming that increase in schooling takes place at age 18, the average of the present discounted values of child beneficiaries' and taxpayers' costs would be \$33 and \$79 respectively.

Increased schooling also poses opportunity cost for child beneficiaries in the form of lost wages while attending school. We again use the opportunity cost of college estimated by Abel and Deitz (2014). The study found that students would forgo \$96,000 in annual earnings (in 2013 dollars) over a 4-year bachelor degree and \$46,000 in annual earnings over a 2-year associate degree. Thus, per year, students would forgo \$24,000 in annual earnings for a bachelor degree and \$23,000 for an associate degree, yielding an average of \$23,500, the equivalent of \$25,778 in 2019 dollars. We thus assume that individuals would give up \$25,778 in earnings in the labor market for every one-year increase in postsecondary education. Multiplying our estimates on increased schooling by \$25,778 gives us \$45-\$766 of opportunity cost. Assuming that child beneficiaries are age 9 and increase in schoolings happen at age 18, the present discounted values of the opportunity cost of schooling range from \$38-\$641, with an average of \$296.

Thus, for child beneficiaries, total costs of increased schooling amounts to an average of \$329. For taxpayers, total costs of increased schooling amounts to an average of \$79.

Standardized Benefits and Costs per \$1,000 Increase in Household Income from Cash and Near-cash Transfers

Table B3 below comes from Garfinkel et al., (2024) and summarizes the calculations explained above. Displayed in the table is the present discounted value of monetary benefits and costs for single child, single parent low-income families per \$1,000 increase in household income, using a discount rate of 2%¹⁹. For detailed calculations see Garfinkel et al. (2022)(2024). The three columns represent benefits and costs for participants (in this case, children and parents who experience changes in household income from the policy), taxpayers, and the society respectively. A positive number indicates a benefit while a negative number indicates a cost. For children, the biggest benefit lies in health and longevity. The present discounted value of increased lifelong health and longevity is \$4,892 per child, over four times the \$1,000 increase in household income. The benefit in increased lifelong earnings is also substantial, with a present discounted value of \$1,940. There are costs for children as well. With higher earnings comes more tax payments (-\$407) and less transfers (-\$26). With more education comes more expenditures on education (-\$329). For parents, the largest benefit is increased health and longevity, valued at \$549. Increased household income also benefits taxpayers in various ways. For taxpayers, the biggest gain comes from the saved expenditures on criminal legal system and reduced victim costs of crime, valued at \$1,432. The improved health of children will reduce taxpayers' share of healthcare expenditures by \$170. The increased earnings of children will generate \$407 more payments to taxpayers. The biggest cost for taxpayers is increased longevity payment (ex: social security) to children due to children's increased longevity, amounting to -\$450.

¹⁹ The latest federal guidance recommends a discount rate of 2% (OMB 2023).

Table B3. Present discounted value of monetary benefits and costs for one child and one parent in low-income families per \$1,000 increase in household income (\$2019)

	Direct + Participants	Indirect = Taxpayers	Total Society
Total transfer	\$ 1,000	\$ -1,000	\$ 0
Increased future earnings of children	\$ 1,940	\$ 0	\$ 1,940
Increased future tax payments by children	\$ -407	\$ 407	\$ 0
Increased children's health and longevity	\$ 4,892	\$ 0	\$ 4,892
Increased parents' and other adults' health and longevity	\$ 549	\$ 0	\$ 549
Avoided expenditures on other cash or near-cash transfers	\$ -26	\$ 26	\$ 0
Avoided expenditures on child protection	\$ 0	\$ 37	\$ 37
Avoided criminal justice expenditures	\$ 0	\$ 372	\$ 372
Reduced victim costs of crime	\$ 0	\$ 1,060	\$ 1,060
Increased costs of children's education	\$ -329	\$ -79	\$ -408
Avoided expenditures on children's health care costs	\$ 20	\$ 162	\$ 183
Avoided expenditures on parents' and other adults' health care costs	\$ 0.36	\$ 2.89	\$ 3.24
Increased payment due to increased children's longevity	\$ 450	\$ -450	\$ 0
Increased payment due to increased parents' and other adults' longevity	\$ 114	\$ -114	\$ 0
Decreased tax payments from parents and other adults^a	\$ 0	\$ 0	\$ 0
Administrative costs^b	\$ 0	\$ -70	\$ -70
Tax distortion costs for taxpayers (inefficiencies due to higher taxation)	\$ 0	\$ -304	\$ -304
Total^c	\$ 8,202	\$ 51	\$ 8,253

Source: *Center on Poverty and Social Policy at Columbia University.*

^a Decreased tax payments from parents and other adults are not yet estimated in this analysis as labor supply effects have not yet been provided. While decreased tax payments are typically small, they will be incorporated when these labor supply effects have been estimated.

^b Our estimate of administrative costs is based on the latest report of the Child Poverty Reduction Advisory Council (CPRAC, 2024). In the CPRAC report, it is estimated that the housing voucher proposal (HP1), which is included in all three policy packages analyzed in this report, could incur a maximum of \$546 million of administrative costs. The public benefit proposals included in each of the three packages could incur a total of \$6.1 million of administrative costs. The tax policy proposals impose minimum or no administrative costs. We thus calculate the administrative cost of the three packages as \$546 million + 6.1 million = \$552.1 million, or \$0.6 billion if rounded.

APPENDIX C. PROJECTION OF THE ACCUMULATED BENEFITS AND COSTS OF THE POLICY PACKAGE OVER TIME

The section above details how we estimate the benefits and costs of the policy package over the lifetimes of program participants. In this section we discuss how we estimate the accumulation of benefits and costs of the policy package over time. This type of projection is frequently conducted by the Congressional Budgetary Office (CBO) in order to examine the long-term budgetary effect of a policy. The most common projection is done over ten years. But the Social Security Administration projects benefits and revenues over 75 years.

When policies are permanently implemented, each year of implementation will incur fiscal costs, and each year of fiscal costs will generate benefits, leading to a stream of fiscal costs and social benefits and costs over time. The analysis of the flow of benefits and costs over time requires the same inputs as the analysis of the benefits and costs of the policy package over the lifetime of program participants: the fiscal cost of the policy package, total increase in household income under the package, and the impact of increases in household income from transfers. But the projection of the flow of social benefits and costs over time is slightly more complicated because the benefits (and the costs derived from the transfers) occur at different times.

Each year the policy is implemented, the policy generates a set of benefits for program participants taxpayers, and the society, leading to a stream of benefits over time. Some benefits occur in the long-run, such as increased future earnings of children, which accrue when children become adults. Some benefits occur as soon as the policy is enacted and household income increases, such as increases in children's health, and avoided criminal justice expenditures and reduced victim costs of crime. But even the benefits that accrue in the first year will be much smaller than the present discounted value of the lifetime increases in those benefits. In addition to the different timing of benefits, we also need to take into consideration that program spending in earlier years generate more years of benefits than program spending in later years. As an example, suppose a policy has a fiscal cost of \$2 billion per year. Over ten years of implementation, the policy will incur a fiscal cost of \$2 billion ten times, with each \$2 billion generating a set of benefits. Some benefits, like increase in children's health, accrue at year 1 (the year the policy is first implemented), while other benefits accrue at year 2, year 3...year 10. By the end of 10 years, the spending in year 1 has been generating benefits for ten years. The spending in year 2 has been generating benefits for nine years. The spending at year 10 has been generating benefits for only one year. Summing up the benefits accumulated within ten years and discounting future benefits, we arrive at the accumulated benefits of the policy over ten years.

Below we discuss in more detail how we project each of the benefits and costs over time, especially the timing of the benefits. In Figures D1 – D6 in the Appendix, we illustrate what the accumulated benefits and costs are of the three policy packages proposed by CPRAC.

Increased Future Earnings of Children

Following the assumption of Garfinkel et al., (2024), we assume the benefit of increased earnings begins when children reach age 19 and start working, and continues until children reach age 65 and retire. Under each of the policy package studied, children are 0-17 years old at the start of the policy, which means that in the first year of the policy, no children would reach age 19 and so there are zero benefits in children's future earnings in the first year. After ten years of policy enactment, the benefits are realized for some children but not all. For instance, children that are 10-17 years old at the start of the policy have reached age 19 while children that are 0-9 years old have yet to do so. Meanwhile, after ten years of policy enactment, there have been ten years of program spending, with the earlier spending generating more years of benefits: the spending in year 1 (the start of the policy) has been generating benefits for children ages 0-17 for ten years, while the spending in year 10 has been generating benefits for children ages 0-17 for only one year. Within each age group, we first sum up the benefits accumulated within 10 years and calculate the present discounted value of the sum by discounting benefits that come after the first year of the policy by 2%. We then calculate a weighted average of the sum by assuming that there is an equal number of children in each age group, giving benefits accumulated by each age group a weight of 1/18. The result is the sum of present discounted value of benefits accumulated by children ages 0-17 within 10 years, or in simpler term, the benefits accumulated after 10 years of policy enactment. As we project benefits 20-80 years into the future, all children will have reached age 19 and start experiencing benefits in increased future earnings, causing the accumulated benefits to grow. Forty years after enactment, the oldest children will have reached their peak earnings years.

Increased Future Tax Payments by Children

Similar to the earning benefit, increase in children's future tax payments starts at age 19 and continues throughout the children's work life (ages 19-65). Increased future tax payments is a cost for children but a benefit for taxpayers. By the end of 10 years of policy enactment, not all children ages 0-17 have started paying higher taxes. For instance, children that are ages 10-17 at the start of the policy have reached age 19 and started paying higher taxes while children ages 0-9 have not. As we move beyond 10 years of policy enactment, more children would have started paying higher taxes, benefitting the taxpayers.

Avoided Expenditures on Other Cash or Near-Cash Transfers

Similar to the earning benefit, avoided expenditures on other cash and near-cash transfers start at age 19 and continue throughout the children's work life (ages 19-65). Reduced transfers is a loss for children but gains for taxpayers. By the end of 10 years of policy enactment, not all children ages 0-17 have experienced reduced transfers. For instance, children that are ages 10-17 at the start of the policy have reached age 19 and have started receiving less transfers, while children ages 0-9 have not. As we move beyond 10 years of policy enactment, more children would have started receiving less transfers due to the increase in their earnings, leading to bigger savings for taxpayers.

Avoided Criminal Justice Expenditures and Reduced Victim Costs of Crime

Following the assumption of Garfinkel et al., (2024), we assume that reduction in crime and thus the avoided criminal justice expenditures on crime and reduced victim costs of crime start at age 12 and continue until age 65. These savings are benefits for taxpayers. When the policy package is first implemented, children that are 12-17 years old would experience reduction in

crime relative to what they would otherwise have committed, so crime benefits would accrue immediately for taxpayers. In the first 10 years of program enactment, crime benefits would accrue for some children but not all. For instance, children that are 1-17 years old at the start of the policy would have non-zero probability of committing crime, while children age 0 still have zero probability of committing crime. Over time, more children would reach age 12 and start experiencing reductions in crime, generating an increasing amount of savings for taxpayers.

Increased Children's Health and Longevity

Following the assumption of Garfinkel et al., (2024), we assume that health benefit starts at age 0 and lasts throughout the entirety of a child's life. The benefit in longevity however, comes much later in life. The average life expectancy in the US is 78 years old (Kochanek, Xu, and Arias 2020). Thus, we count improvement in children's health from ages 0-78 and count increase in longevity at age 78. Health benefits start as soon as the policy is implemented. After 10 years of policy enactment, all children ages 0-17 have accumulated some health benefits from multiple years of policy spending. There are no longevity benefits yet because no children have reached the age of 78. As the policy goes beyond 10 years, children accumulate more health benefits. Longevity benefit first accrues after 62 years of policy enactment for children that are 17 years old at the start of the policy. Afterwards, children in the younger age group will reach age 78 and begin to accumulate longevity benefits as well.

Increased Parents' and Other Adults'²⁰ Health and Longevity

Following the assumption of Garfinkel et al., (2024), we assume that health benefits start when adults (including both parents and adults without children) are age 29 and continue until they reach age 65. Longevity benefit occurs at age 78. Adults are 29-46 years old at the start of the policy so the health benefits will accrue as soon as the policy is enacted. As we go from 10-80 years, more adults would accumulate health benefits. Benefit in adult longevity first accrues after 33 years of policy enactment for adults that are 46 years old at the start of the policy. Afterwards, longevity benefits will accrue for the younger adults.

Avoided Expenditures on Children's and Adults' Health Care Costs

Similar to the health benefit, reduction in healthcare expenditures occur as soon as the policy is enacted. Savings in healthcare expenditures not only benefit the children throughout their lives (ages 0-78), but also the taxpayers. By the end of 10 years of policy enactment, all children ages 0-17 have seen reductions in their healthcare expenditures, benefitting both the children and taxpayers. As we project the benefit beyond 10 years, healthcare savings for children and taxpayers will continue to grow.

Similar to the benefit for children, reduction in healthcare expenditures occur for adults as soon as the policy is enacted. Savings in healthcare expenditures not only benefit the adults throughout their lives (ages 29-78), but also the taxpayers. By the end of 10 years of policy enactment, adults of all ages have seen some reductions in their healthcare expenditures. As we project the benefit beyond 10 years, healthcare savings for adults and taxpayers will continue to grow.

²⁰ Adults include both parents and adults without children.

Increased Payment Due to Increased Children's and Adults' Longevity

Similar to the benefit of increased longevity, increased longevity-payment doesn't accrue until children and adults reach age 78. Such payment benefits children and adults but has to be paid by taxpayers. In the first year of the policy, no children or adults would have reached 78 years old, so there would be no longevity payment. This is also the case after only 10 years of policy enactment. Longevity payment for adults first accrues after 33 years of policy enactment for parents that are 46 years old at the start of the policy. Longevity payment first accrues for children after 62 years of policy enactment for children that are 17 years old at the start of the policy.

Avoided Expenditures on Child Protective Services

Following the assumption of Garfinkel et al., (2024), we assume that avoided expenditures on child protective services accrue immediately when the policy is first enacted and last for only one year. These savings are benefits for taxpayers. By the end of 10 years of policy enactment, all children ages 0-17 would have benefitted from less involvement in the child welfare system, generating some savings for taxpayers. As we go beyond 10 years of policy enactment, more savings will be generated for taxpayers.

Increased Costs of Children's Education

Following the same assumption of Garfinkel et al., (2024), we assume increased education and its costs occur at age 18 and do not last beyond age 18. Increased education poses costs both for children and taxpayers. In the first year the policy is enacted, no children would be age 18 so there would be no costs during the first year. After ten years of policy, some children (ex: those that are 9-17 years old at the start of the policy) would start accumulating these costs but not all. As we go beyond 10 years, these costs would grow for children and taxpayers as more children reach age 18.

APPENDIX D. ADDITIONAL RESULTS

Table D1. Aggregate benefits and costs of Policy Package 1 generated by a single year of package implementation: Present discounted value (\$2019 billions)

	Direct + Participants	Indirect = Taxpayers	Total Society
Total transfer	\$ 8.9	\$ -8.9	\$ 0
Increased future earnings of children	\$ 22.3	\$ 0	\$ 22.3
Increased future tax payments by children	\$ -4.7	\$ 4.7	\$ 0
Increased children's health and longevity	\$ 56.1	\$ 0	\$ 56.1
Increased parents' and other adults' health and longevity	\$ 4.0	\$ 0	\$ 4.0
Avoided expenditures on other cash or near-cash transfers	\$ -0.3	\$ 0.3	\$ 0
Avoided expenditures on child protection	\$ 0	\$ 0.4	\$ 0.4
Avoided criminal justice expenditures	\$ 0	\$ 4.3	\$ 4.3
Reduced victim costs of crime	\$ 0	\$ 12.1	\$ 12.1
Increased costs of children's education	\$ -3.8	\$ -0.9	\$ -4.7
Avoided expenditures on children's health care costs	\$ 0.2	\$ 1.9	\$ 2.1
Avoided expenditures on parents' and other adults' health care costs	\$0.003	\$ 0.02	\$0.023
Increased payment due to increased children's longevity	\$ 5.2	\$ -5.2	\$ 0
Increased payment due to increased parents' and other adults' longevity	\$ 0.8	\$ -0.8	\$ 0
Decreased tax payments from parents and other adults^a	\$ 0	\$ 0	\$ 0
Administrative costs^b	\$ 0	\$ -0.6	\$ -0.6
Tax distortion costs for taxpayers (inefficiencies due to higher taxation)	\$ 0	\$ -2.2	\$ -2.2
Total^c	\$ 88.9	\$ 5.1	\$ 94.0

Source: Center on Poverty and Social Policy at Columbia University, 2024.

^a Decreased tax payments from parents and other adults are not yet estimated in this analysis as labor supply effects have not yet been provided. While decreased tax payments are typically small, they will be incorporated when these labor supply effects have been estimated.

^b Our estimate of administrative costs is based on the latest report of the Child Poverty Reduction Advisory Council (CPRAC, 2024). In the CPRAC report, it is estimated that the housing voucher proposal (HP1), which is included in all three policy packages analyzed in this report, could incur a maximum of \$546 million of administrative costs. The public benefit proposals included in each of the three packages could incur a total of \$6.1 million of administrative costs. The tax policy proposals impose minimum or no administrative costs. We thus calculate the administrative cost of the three packages as \$546 million + 6.1 million = \$552.1 million, or \$0.6 billion if rounded.

^c Note that, due to rounding procedures, totals might not be an exact sum of their parts.

Table D2. Aggregate benefits and costs of Policy Package 2 generated by a single year of package implementation: Present discounted value (\$2019 billions)

	Direct + Participants	Indirect = Taxpayers	Total Society
Total transfer	\$ 6.6	\$ -6.6	\$ 0
Increased future earnings of children	\$ 15.5	\$ 0	\$ 15.5
Increased future tax payments by children	\$ -3.2	\$ 3.2	\$ 0
Increased children's health and longevity	\$ 39.0	\$ 0	\$ 39.0
Increased parents' and other adults' health and longevity	\$ 3.3	\$ 0	\$ 3.3
Avoided expenditures on other cash or near-cash transfers	\$ -0.2	\$ 0.2	\$ 0
Avoided expenditures on child protection	\$ 0	\$ 0.3	\$ 0.3
Avoided criminal justice expenditures	\$ 0	\$ 3.0	\$ 3.0
Reduced victim costs of crime	\$ 0	\$ 8.4	\$ 8.4
Increased costs of children's education	\$ -2.6	\$ -0.6	\$ -3.3
Avoided expenditures on children's health care costs	\$ 0.2	\$ 1.3	\$ 1.5
Avoided expenditures on parents' and other adults' health care costs	\$ 0.002	\$ 0.02	\$ 0.02
Increased payment due to increased children's longevity	\$ 3.6	\$ -3.6	\$ 0
Increased payment due to increased parents' and other adults' longevity	\$ 0.7	\$ -0.7	\$ 0
Decreased tax payments from parents and other adults^a	\$ 0	\$ 0	\$ 0
Administrative costs^b	\$ 0	\$ -0.6	\$ -0.6
Tax distortion costs for taxpayers (inefficiencies due to higher taxation)	\$ 0	\$ -1.8	\$ -1.8
Total^c	\$ 62.8	\$ 2.6	\$ 65.3

Source: Center on Poverty and Social Policy at Columbia University, 2024.

^a Decreased tax payments from parents and other adults are not yet estimated in this analysis as labor supply effects have not yet been provided. While decreased tax payments are typically small, they will be incorporated when these labor supply effects have been estimated.

^b Our estimate of administrative costs is based on the latest report of the Child Poverty Reduction Advisory Council (CPRAC, 2024). In the CPRAC report, it is estimated that the housing voucher proposal (HP1), which is included in all three policy packages analyzed in this report, could incur a maximum of \$546 million of administrative costs. The public benefit proposals included in each of the three packages could incur a total of \$6.1 million of administrative costs. The tax policy proposals impose minimum or no administrative costs. We thus calculate the administrative cost of the three packages as \$546 million + 6.1 million = \$552.1 million, or \$0.6 billion if rounded.

^c Note that, due to rounding procedures, totals might not be an exact sum of their parts.

Table D3. Aggregate benefits and costs of Policy Package 3 generated by a single year of package implementation: Present discounted value (\$2019 billions)

	Direct + Participants	Indirect = Taxpayers	Total Society
Total transfer	\$ 8.5	\$ -8.5	\$ 0
Increased future earnings of children	\$ 20.5	\$ 0	\$ 20.5
Increased future tax payments by children	\$ -4.3	\$ 4.3	\$ 0
Increased children's health and longevity	\$ 51.7	\$ 0	\$ 51.7
Increased parents' and other adults' health and longevity	\$ 3.7	\$ 0	\$ 3.7
Avoided expenditures on other cash or near-cash transfers	\$ -0.3	\$ 0.3	\$ 0
Avoided expenditures on child protection	\$ 0	\$ 0.4	\$ 0.4
Avoided criminal justice expenditures	\$ 0	\$ 3.9	\$ 3.9
Reduced victim costs of crime	\$ 0	\$ 11.2	\$ 11.2
Increased costs of children's education	\$ -3.5	\$ -0.8	\$ -4.3
Avoided expenditures on children's health care costs	\$ 0.2	\$ 1.7	\$ 1.9
Avoided expenditures on parents' and other adults' health care costs	\$0.002	\$ 0.02	\$ 0.02
Increased payment due to increased children's longevity	\$ 4.7	\$ -4.7	\$ 0
Increased payment due to increased parents' and other adults' longevity	\$ 0.8	\$ -0.8	\$ 0
Decreased tax payments from parents and other adults^a	\$ 0	\$ 0	\$ 0
Administrative costs^b	\$ 0	\$ -0.6	\$ -0.6
Tax distortion costs for taxpayers (inefficiencies due to higher taxation)	\$ 0	\$ -2.1	\$ -2.1
Total^c	\$ 82.0	\$ 4.3	\$ 86.3

Source: Center on Poverty and Social Policy at Columbia University, 2024.

^a Decreased tax payments from parents and other adults are not yet estimated in this analysis as labor supply effects have not yet been provided. While decreased tax payments are typically small, they will be incorporated when these labor supply effects have been estimated.

^b Our estimate of administrative costs is based on the latest report of the Child Poverty Reduction Advisory Council (CPRAC, 2024). In the CPRAC report, it is estimated that the housing voucher proposal (HP1), which is included in all three policy packages analyzed in this report, could incur a maximum of \$546 million of administrative costs. The public benefit proposals included in each of the three packages could incur a total of \$6.1 million of administrative costs. The tax policy proposals impose minimum or no administrative costs. We thus calculate the administrative cost of the three packages as \$546 million + 6.1 million = \$552.1 million, or \$0.6 billion if rounded.

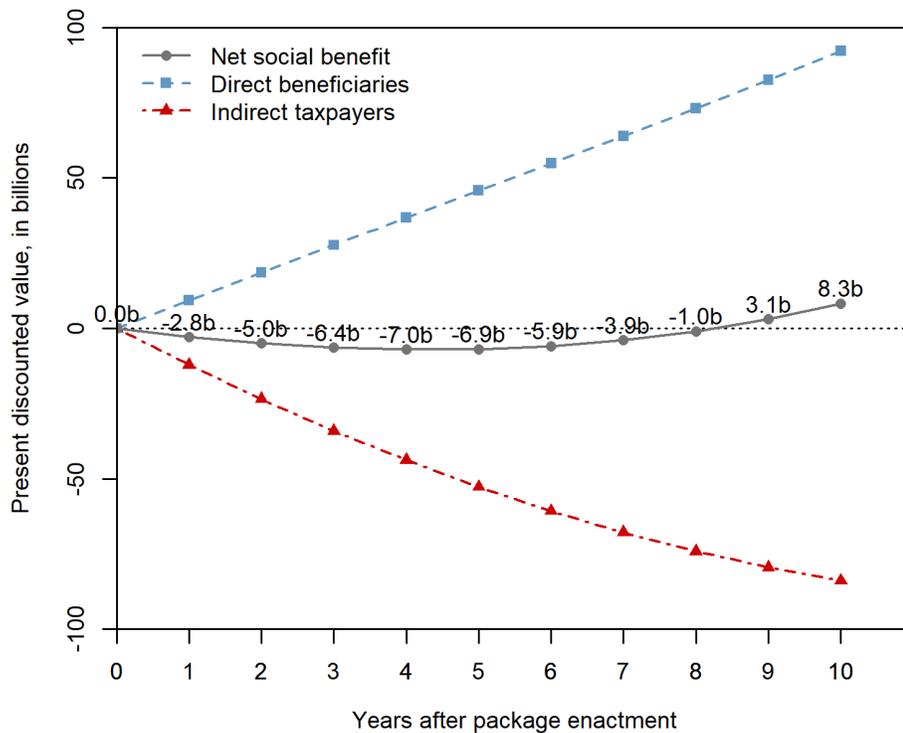
^c Note that, due to rounding procedures, totals might not be an exact sum of their parts.

Table D4. Selected lifetime benefits and costs for Policy Package 1 generated by the first year of package implementation (\$2019 billions)

Years after program implementation	Children's earnings	Children's health and longevity	Avoided criminal justice expenditures & reduced victim costs of crime	Fiscal cost of transfers (taxpayer cost)	Tax distortion costs for taxpayers	Transfers to participants	Net social benefits
0	\$0	\$0	\$0	\$0	\$0	\$0	\$0
1	\$0	\$0.35	\$0.5	-\$8.9	-\$3.6	\$8.9	-\$3
10	\$0.71	\$3.74	\$6.33		-\$3.1		\$5
20	\$4.96	\$9.56	\$11.73		-\$2.24		\$20
30	\$10.92	\$16.19	\$13.91		-\$1.41		\$36
40	\$15.9	\$21.7	\$15.28		-\$0.91		\$50
50	\$19.93	\$26.21	\$16.04		-\$0.57		\$62
60	\$21.92	\$29.92	\$16.39		-\$0.26		\$69
70	\$22.16	\$44.87	\$16.43		-\$1.27		\$83
80	\$22.16	\$55.96	\$16.43		-\$2.19		\$94

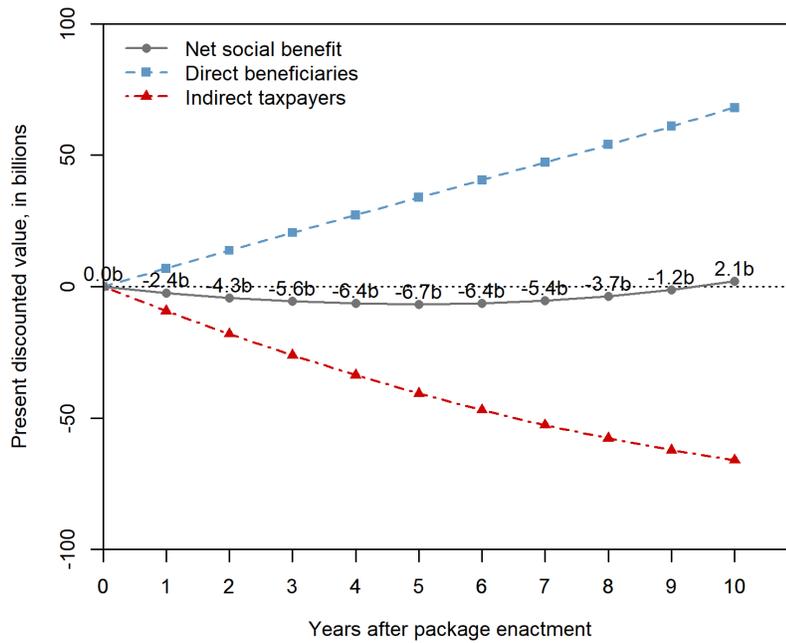
Source: Center on Poverty and Social Policy at Columbia University, 2024.

Figure D1. The Flow of Cumulative Benefits and Costs Over Time in the First 10 Years of Policy Package 1 Implementation (\$2019)



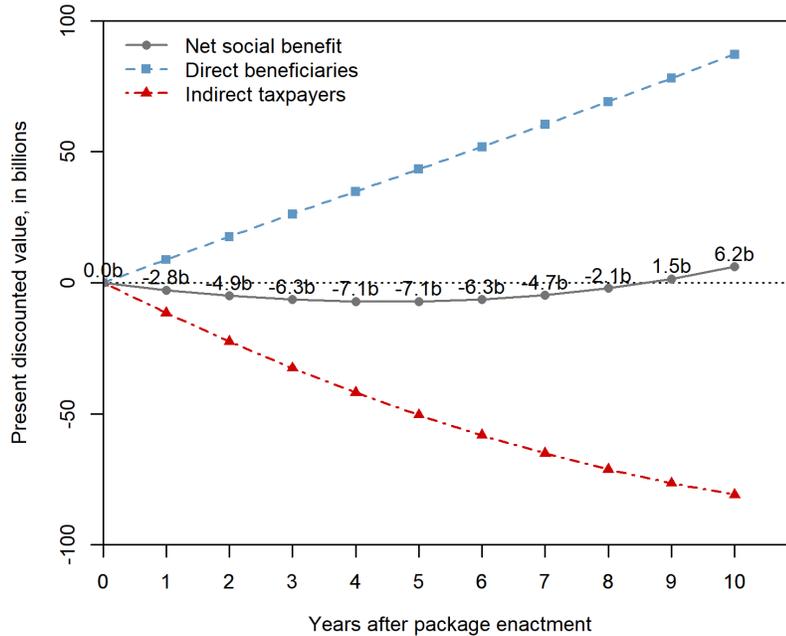
Source: Center on Poverty and Social Policy at Columbia University, 2024.

Figure D2. The Flow of Cumulative Benefits and Costs Over Time in the First 10 Years of Policy Package 2 Implementation (\$2019)



Source: Center on Poverty and Social Policy at Columbia University, 2024.

Figure D3. The Flow of Cumulative Benefits and Costs Over Time in the First 10 Years of Policy Package 3 Implementation (\$2019)



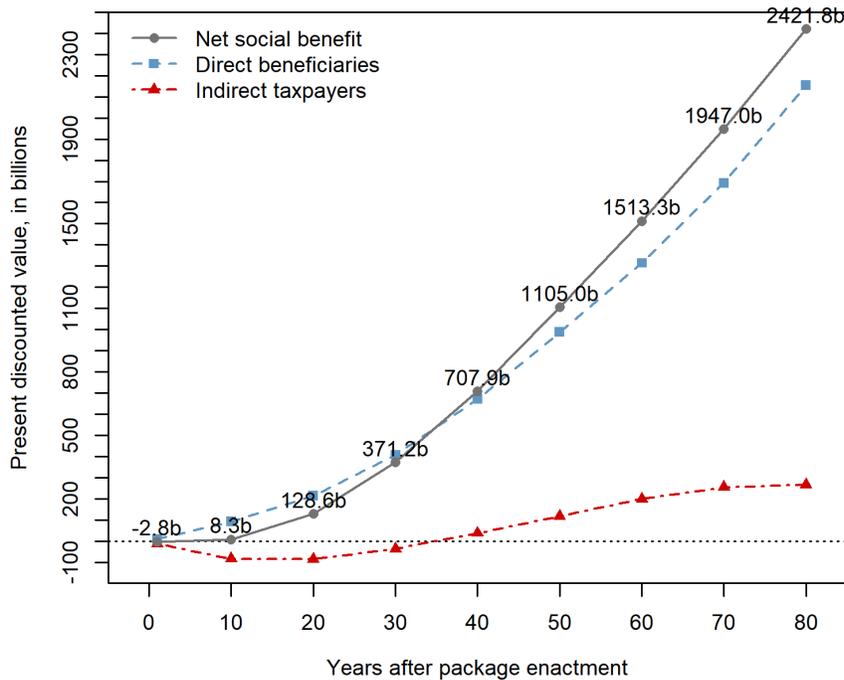
Source: Center on Poverty and Social Policy at Columbia University, 2024.

Table D5. The Flow of Cumulative Benefits and Costs Over Time in the First 10 Years of Policy Package Implementation (\$2019 billions)

Years after program implementation	Package 1			Package 2			Package 3		
	Total society	Indirect taxpayers	Direct beneficiaries	Total society	Indirect taxpayers	Direct beneficiaries	Total society	Indirect taxpayers	Direct beneficiaries
1	-\$2.84	-\$12.14	\$9.31	-\$2.38	-\$9.28	\$6.90	-\$2.77	-\$11.58	\$8.81
2	-\$4.97	-\$23.51	\$18.54	-\$4.26	-\$18.00	\$13.74	-\$4.89	-\$22.44	\$17.55
3	-\$6.37	-\$34.07	\$27.70	-\$5.62	-\$26.14	\$20.52	-\$6.33	-\$32.55	\$26.21
4	-\$7.02	-\$43.79	\$36.78	-\$6.45	-\$33.69	\$27.24	-\$7.08	-\$41.88	\$34.80
5	-\$6.89	-\$52.67	\$45.79	-\$6.73	-\$40.64	\$33.90	-\$7.11	-\$50.43	\$43.32
6	-\$5.89	-\$60.68	\$54.79	-\$6.40	-\$46.96	\$40.56	-\$6.33	-\$58.16	\$51.83
7	-\$3.94	-\$67.81	\$63.87	-\$5.40	-\$52.66	\$47.25	-\$4.68	-\$65.08	\$60.40
8	-\$0.97	-\$74.05	\$73.08	-\$3.69	-\$57.73	\$54.04	-\$2.08	-\$71.18	\$69.10
9	\$3.09	-\$79.41	\$82.50	-\$1.21	-\$62.17	\$60.96	\$1.51	-\$76.46	\$77.97
10	\$8.29	-\$83.88	\$92.17	\$2.07	-\$65.98	\$68.05	\$6.16	-\$80.92	\$87.08

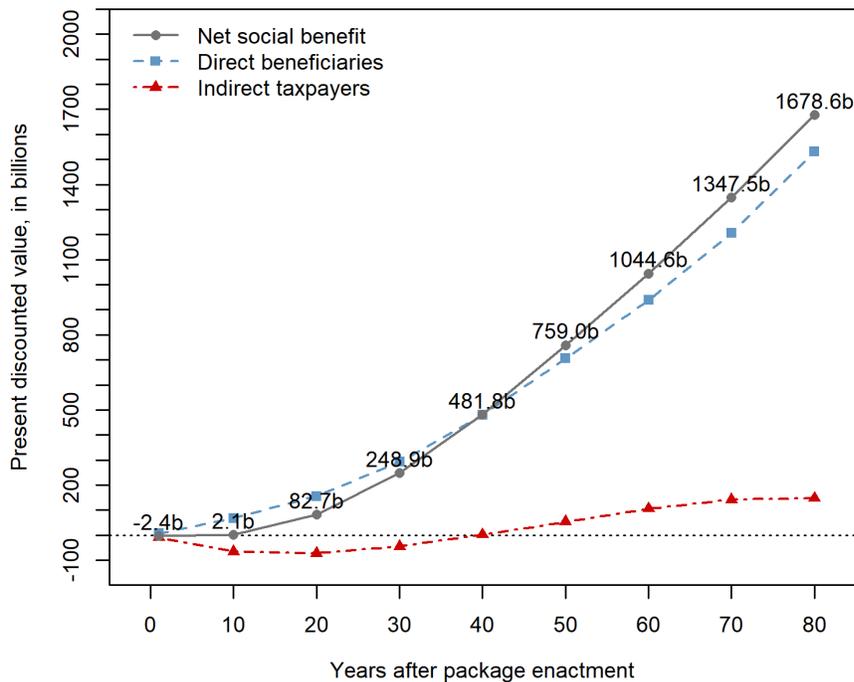
Source: Center on Poverty and Social Policy at Columbia University, 2024.

Figure D4. The Flow of Cumulative Benefits and Costs Over Time in the First 80 Years of Policy Package 1 Implementation (\$2019)



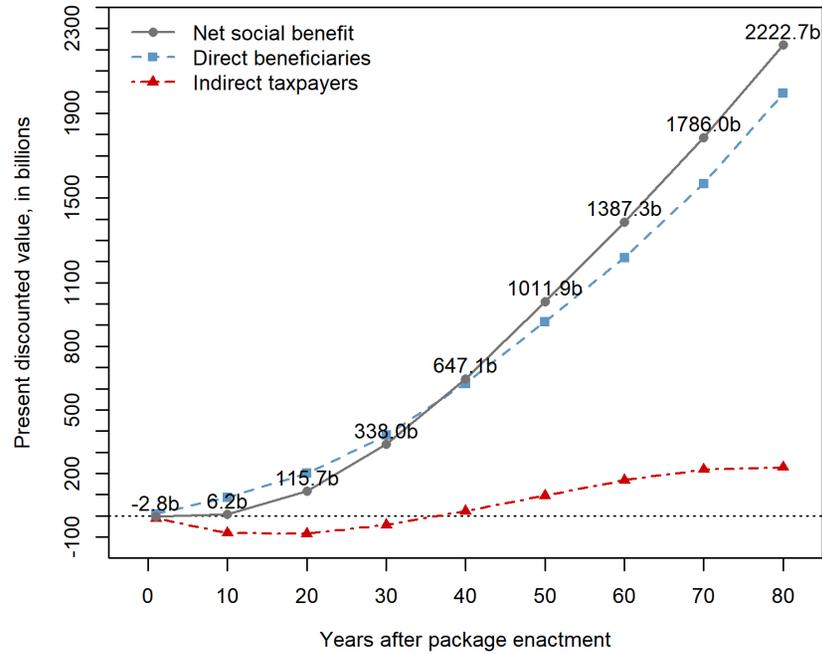
Source: Center on Poverty and Social Policy at Columbia University, 2024.

Figure D5. The Flow of Cumulative Benefits and Costs Over Time in the First 80 Years of Policy Package 2 Implementation (\$2019)



Source: Center on Poverty and Social Policy at Columbia University, 2024.

Figure D6. The Flow of Cumulative Benefits and Costs Over Time in the First 80 Years of Policy Package 3 Implementation (\$2019)



Source: Center on Poverty and Social Policy at Columbia University, 2024.

Table D6. The Flow of Cumulative Benefits and Costs Over Time in the First 80 Years of Policy Package Implementation (\$2019 billions)

Year	Package 1			Package 2			Package 3		
	Total society	Indirect taxpayers	Direct beneficiaries	Total society	Indirect taxpayers	Direct beneficiaries	Total society	Indirect taxpayers	Direct beneficiaries
1	-\$2.84	-\$12.14	\$9.31	-\$2.38	-\$9.28	\$6.90	-\$2.77	-\$11.58	\$8.81
10	\$8.29	-\$83.88	\$92.17	\$2.07	-\$65.98	\$68.05	\$6.16	-\$80.92	\$87.08
20	\$128.63	-\$85.05	\$213.68	\$82.73	-\$73.11	\$155.85	\$115.72	-\$85.04	\$200.76
30	\$371.22	-\$36.28	\$407.50	\$248.91	-\$44.42	\$293.33	\$338.01	-\$42.65	\$380.66
40	\$707.89	\$37.10	\$670.79	\$481.82	\$1.99	\$479.83	\$647.07	\$22.82	\$624.25
50	\$1,104.96	\$117.66	\$987.29	\$759.01	\$53.48	\$705.53	\$1,011.92	\$95.25	\$916.66
60	\$1,513.35	\$198.94	\$1,314.40	\$1,044.64	\$106.00	\$938.65	\$1,387.31	\$168.65	\$1,218.66
70	\$1,947.01	\$254.63	\$1,692.37	\$1,347.47	\$141.44	\$1,206.03	\$1,786.04	\$218.75	\$1,567.29
80	\$2,421.80	\$266.77	\$2,155.03	\$1,678.58	\$147.22	\$1,531.37	\$2,222.70	\$228.98	\$1,993.72

Source: Center on Poverty and Social Policy at Columbia University, 2024.

APPENDIX E. NEW YORK STATE/LOCAL AND FEDERAL BREAKDOWN OF BENEFITS AND COSTS

Tables E1, E2, and E3 break down the benefits and costs of taxpayers from policy packages 1, 2 and 3, respectively, into those that are New York state/local and those that are federal. The last column reproduces the taxpayer column in Table 5, except we calculate the total benefits and costs twice, first excluding and then including the benefits from reduced victim costs of crime and the cost associated with inefficiencies/distortion stemming from higher taxes. Because neither of these benefits and costs are relevant to the government budget, they are considered “non-fiscal.” To simplify we include benefits and costs to other states than New York in the federal category.

Benefits accrued to New York State vs Federal government: The biggest “fiscal” benefit relevant to the government’s budget at the State and Federal levels comes from increased future tax payments by children, at \$4.7 billion. Using data on taxes from New York State (New York State Department of Taxation and Finance 2023) and the Internal Revenue Service (IRS 2023), we estimate that \$3.68 billion, or 78% of the total \$4.7 billion in increased future tax payments would be increases in federal tax payments, leaving \$1.02 billion, or 22% of the increase for State/local tax payments.²¹ Savings resulting from reduced healthcare spending would also be majority federal. Using New York state Medicaid data (Office of the NEW YORK 2023a), we estimate that 57% of the savings from reduced healthcare spending would accrue to the Federal government.²² Savings resulting from reduced criminal justice expenditures and public assistance accrue mainly to State/local government. Data suggests that 83% (Urban Institute 2021)²³ of the reduction in criminal justice expenditures and 74% (Office of the NEW YORK 2023b)²⁴ of reduction in public assistance expenditures would accrue to New York State/local

²¹ According to the New York State Department of Taxation and Finance (2023), New York state collected \$80,445,164,911, \$118,528,148,966, and \$108,593,023,191 in taxes (including personal income tax, corporation and business tax, sales, excise and user tax, property transfers tax, and others) in the fiscal year 2021, 2022, and 2023 respectively. According to Internal Revenue Service (IRS 2023), the federal government collected \$330,143,910,000, \$393,135,616,000, and \$373,831,039,000 in taxes (including individual income tax and employment taxes, estate tax, corporation income tax, estate and trust income tax, gift tax, excise taxes, and others) from New York state in the fiscal year 2021, 2022, and 2023 respectively. Adding up the taxes collected by New York state and the taxes collected by the federal government from New York state, we arrive at a total of \$410,589,074,911, \$511,663,764,966, and \$482,424,062,191 of state and federal taxes in New York. Dividing the federal taxes in New York by the total taxes in New York, we arrive at the estimate that federal taxes are 80.4%, 76.8%, and 77.5% of total taxes in New York during the fiscal year 2021, 2022, and 2023 respectively. The average of the three percentages is 78%. We use the 78% for our calculation. Multiplying \$4.7 billion increase in future taxes from Package 1 by 78%, we conclude that \$3.68 billion increase would be federal taxes. The rest of the \$1.02 billion would be state/local taxes.

²² According to Office of the NEW YORK (2023), in the state fiscal year 2022-23, New York state Medicaid spending was a total of \$92.3 billion, \$53.1 billion of which came from federal funds. We thus approximate that the federal share of total healthcare spending in New York is around 57% ($53.1/92.3$). Thus, for a total of \$1.9 billion in savings in children’s healthcare costs from Package 1, we calculate that \$1.083 billion (1.9×0.57) is federal and the rest of the \$0.817 billion is state/local.

²³ According to the Urban Institute (2021), state and local governments spent \$135 billion on police, \$87 billion on corrections, and \$52 billion on courts in the fiscal year of 2021, summing up to a total of \$274 billion on criminal justice expenditures. When it comes to federal contributions, the federal government spent \$30 billion on police, \$7 billion on corrections, and \$15 billion on courts in the fiscal year of 2017, summing up to a total of \$52 billion. We adjust the 2017 federal contributions for inflation and arrive at \$57.5 billion in 2021 dollars. Adding up the spending of state and federal government, we arrive at a total spending of \$331.5 billion, 17% of which are federal ($57.5/331.5$) and 83% of which are local ($274/331.5$). We use these percentages in our calculation. Multiplying the \$4.3 billion of savings on crime from Package 1 by 83% and 17%, we arrive at \$3.569 and \$0.731 billion of savings for the state and federal taxpayers respectively.

²⁴ According to Office of the NEW YORK (2023b), in the state fiscal year 2022-23, of the \$2,363 millions of public assistance, \$614 millions are primarily federally-funded. \$614 million is close to 26% of \$2,363 million. We thus assume that 74% of savings in transfers are state and local while the rest of the 26% are federal. There are \$0.3

government. Finally, using data from New York State Office of Children and Family Services (2024), we estimate that savings in child protection costs would accrue to State/local and Federal governments equally.²⁵

Costs accrued to State vs Federal Government: On the cost side, the fiscal cost and administrative cost of the package, and the taxpayer portion of the increase in children's education cost that results from improved school attendance and performance are sole responsibilities of local government. Increased expenditures due to increased longevity, including for example on government programs like Social Security, Medicare, and Medicaid that individuals might utilize more in their older years as they live longer lives, are financed mostly by federal taxes. We estimate that 6% of these expenditures are paid by the State while 94% are paid by the Federal government.²⁶

Totaling costs and benefits accrued to State vs Federal governments: The first total sums up the "fiscal" benefits and costs, excluding the non-fiscal estimates related to benefits from the reduced victim costs of crime and costs associated with inefficiencies/distortion stemming from higher taxes. Excluding these non-fiscal estimates, the first total cited indicates Package 1 incurs a net cost of \$4.72 billion in State/local taxes and leads to a net benefit of \$0.14 billion in Federal taxes. The next two rows after the first total show the distribution of the benefits from reduced victim costs of crime and costs associated with inefficiencies/distortion stemming

billion of savings in transfers from Package 1. We thus multiply \$0.3 billion by 74% and 26% to calculate the state and federal portions of the savings, arriving at \$0.222 and \$0.078 billion respectively.

²⁵ According to New York State Office of Children and Family Services (2024), an important federal funding source for foster care services in New York is Title IV-E of the Social Security Act. In New York, this federal funding source offsets 50% of the cost of providing foster care services. We thus roughly assume that 50% of the savings in child protection costs are federal and the other 50% are state and local.

²⁶ To estimate the state share of extra expenditures on the aged as a consequence of them living longer, we need to estimate the state share of both healthcare expenditures and cash transfers on the aged.

First, we calculate what percentage of the longevity expenditures are healthcare versus cash transfer. Healthcare expenditures on the aged include Medicare and Medicaid. According to Centers for Medicare & Medicaid Services (2022), Medicare spending was \$944.3 billion in 2022 and Medicaid spending was \$805.7 billion in 2022. According to the State Comptroller's Office of the New York, 55.1% of Medicaid spending in the fiscal year 2023 was on the elderly, blind and disabled. Thus, total medical expenditures on the aged are about \$1,388 (944.3 + 805.7 * 55.1%) billion. Cash transfer expenditures on the aged include OA & DI payments and SSI payment. Federal data from the Social Security Administration shows that total OA & DI payments equaled \$1,091 billion and total SSI payments equaled \$ 57.561 billion. Thus, total cash expenditures on the aged are about \$1,148.561 (1091 + 57.561) billion. Medical expenditures are thus 54.72% (1388/(1388+1148.561)) of total government expenditures on the aged. Cash expenditures are 45.28% (1148.561/(1388+1148.561)) of total government expenditures on the aged.

We then calculate the distribution of healthcare expenditures between federal and state/local taxpayers. For people aged 65 and above, Medicare and Medicaid are the two main options for public healthcare. According to Centers for Medicare & Medicaid Services (2022), Medicare spending was \$944.3 billion in 2022 and Medicaid spending was \$805.7 billion in 2022. This suggests that of the total Medicare and Medicaid spending, 54% is Medicare while 46% is Medicaid. We assume that 100% of Medicare is federal-funded. According to the State Comptroller's Office of the New York, 55.1% of Medicaid spending in the fiscal year 2023 was on the elderly, blind and disabled and 57% of Medicaid are federally funded (so 43% are state funded). Thus, the state share of healthcare spending on the aged and disabled equals (46% * 55.1% * 43%), or 10.89%.

Next, we calculate the distribution of the cash assistance expenditures between federal and state/local taxpayers. Federal data from the Social Security Administration shows that total OA & DI payments equaled \$1,091 billion and total SSI payments equaled \$ 57.561 billion. Thus, SSI equaled 5% of total cash transfers to the aged. In Table 17 of the 2022 caseload statistics report from the NYS Office of Temporary and Disability Assistance Statistics, total program expenditures on SSI in NYS were \$4.935 billion. The federal contribution was \$4.406 billion, and the state contribution was \$529 million. Thus, 89.3% of total SSI expenditures in NYS are federal, and 10.7% are state in NYS. The share of state expenditures on cash transfers was therefore, 10.7% * 5% or 0.54%.

Using all the information above, the state thus pays for 6.2% (54.72% * 10.89% + 45.28% * 0.54%) of the longevity expenditures, while the federal government pays for 93.8% (54.72 * 89.11% + 45.28 * 99.46%) of the longevity expenditures.

from higher taxes. Benefits from reduced victim costs of crime are substantial, at \$12.1 billion. Using data on New York state mobility, we estimate that 83% of the benefits from reduced victim costs (\$10.04 billion) would accrue to State/local government, while 17% (\$2.06 billion) would accrue to the federal government.²⁷ On the costs associated with inefficiencies/distortion stemming from higher taxes, we estimate that Package 1 would result in \$1.99 billion of inefficiencies accruing to New York State/local government, and \$0.09 billion of inefficiencies accruing to Federal government/taxpayers. The last row of the table presents the total when including the benefits from reduced victim cost of crime and costs associated with inefficiencies/distortion stemming from higher taxes. Due to the substantial benefits of reduced victim costs of crime, we estimate that a net benefit of \$3.3 billion would accrue to New York State/local government, a net benefit of \$2 billion would accrue to the Federal government, and a total net benefit of \$5.3 billion would accrue across all levels of government/taxpayers.

²⁷ In order to estimate the state's share of reductions in the victim costs of crime, we need to estimate what percent of children born in New York move out of the state and the age of the children when they move. Out-migration from New York State means that some of the benefits from reduced victim costs of crime will be realized in other states. To conduct a thorough analysis that incorporates this nuance would be a huge research undertaking that is well beyond the scope of our analysis for the state. Thus, we conduct a very crude analysis:

A 2024 report from the Fiscal Policy Institute suggests that there was a rate of net out-migration of roughly 1% both pre and post pandemic among New York residents who were born in NY (Fiscal Policy Institute 2024). Assuming the rate of net out-migration remains constant over the course of 20 years, meaning that every year 1% of the New York state residents in that year move away, the rate of net out-migration 20 years from now is estimated to be 17.4%. We thus allocate 17.4% of the reduction in victim costs of crime to other states. We know this estimate could be too high because it doesn't take account of the reduction in costs that also result from out-migration. According to the Fiscal Policy Institute out-migration is most common for children below the age of six. On the other hand, the estimate also could be too low because net out-migration is lower than gross out-migration.

**Table E1. Breakdown of taxpayers' benefits and costs from Policy Package 1:
Present discounted value (\$2019 billions).**

	State/Local	Federal	All Taxpayers
Total transfer	\$ -8.9	\$ 0	\$ -8.9
Increased future earnings of children	\$ 0	\$ 0	\$ 0
Increased future tax payments by children	\$ 1.02	\$ 3.68	\$ 4.7
Increased children's health and longevity	\$ 0	\$ 0	\$ 0
Increased parents' and other adults' health and longevity	\$ 0	\$ 0	\$ 0
Avoided expenditures on other cash or near-cash transfers	\$ 0.222	\$ 0.078	\$ 0.3
Avoided expenditures on child protection	\$ 0.2	\$ 0.2	\$ 0.4
Avoided criminal justice expenditures	\$ 3.569	\$ 0.731	\$ 4.3
Increased costs of children's education	\$ -0.9	\$ 0	\$ -0.9
Avoided expenditures on children's health care costs	\$ 0.817	\$ 1.083	\$ 1.9
Avoided expenditures on parents' and other adults' health care costs	\$ 0.0086	\$ 0.0114	\$ 0.02
Increased payment due to increased children's longevity	\$ -0.31	\$ -4.89	\$ -5.2
Increased payment due to increased parents' and other adults' longevity	\$ -0.05	\$ -0.75	\$ -0.8
Decreased tax payments from parents and other adults^a	\$ 0	\$ 0	\$ 0
Administrative costs^b	\$ -0.6	\$ 0	\$ -0.6
Total (without victim cost of crime or tax distortion costs)	\$ -4.92	\$ 0.14	\$ -4.8
Tax distortion costs for taxpayers (inefficiencies due to higher taxation)	\$ -2.08	\$ -0.1	\$ -2.2
Reduced victim cost of crime	\$ 10.04	\$ 2.06	\$ 12.1
Total (with victim cost of crime and tax distortion costs)^c	\$ 3.04	\$ 2	\$ 5.1

Source: Center on Poverty and Social Policy at Columbia University, 2024.

^a Decreased tax payments from parents and other adults are not yet estimated in this analysis as labor supply effects have not yet been provided. While decreased tax payments are typically small, they will be incorporated when these labor supply effects have been estimated.

^b Our estimate of administrative costs is based on the latest report of the Child Poverty Reduction Advisory Council (CPRAC, 2024). In the CPRAC report, it is estimated that the housing voucher proposal (HP1), which is included in all three policy packages analyzed in this report, could incur a maximum of \$546 million of administrative costs. The public benefit proposals included in each of the three packages could incur a total of \$6.1 million of administrative costs. The tax policy proposals impose minimum or no administrative costs. We thus calculate the administrative cost of the three packages as \$546 million + 6.1 million = \$552.1 million, or \$0.6 billion if rounded.

^c Note that, due to rounding procedures, totals might not be an exact sum of their parts.

**Table E2. Breakdown of taxpayers' benefits and costs from Policy Package 2:
Present discounted value (\$2019 billions).**

	State/Local	Federal	All Taxpayers
Total transfer	\$ -6.6	\$ 0	\$ -6.6
Increased future earnings of children	\$ 0	\$ 0	\$ 0
Increased future tax payments by children	\$ 0.7	\$ 2.5	\$ 3.2
Increased children's health and longevity	\$ 0	\$ 0	\$ 0
Increased parents' and other adults' health and longevity	\$ 0	\$ 0	\$ 0
Avoided expenditures on other cash or near-cash transfers	\$ 0.148	\$ 0.052	\$ 0.2
Avoided expenditures on child protection	\$ 0.15	\$ 0.15	\$ 0.3
Avoided criminal justice expenditures	\$ 2.49	\$ 0.51	\$ 3.0
Increased costs of children's education	\$ -0.6	\$ 0	\$ -0.6
Avoided expenditures on children's health care costs	\$ 0.559	\$ 0.741	\$ 1.3
Avoided expenditures on parents' and other adults' health care costs	\$ 0.0086	\$ 0.0114	\$ 0.02
Increased payment due to increased children's longevity	\$ -0.22	\$ -3.38	\$ -3.6
Increased payment due to increased parents' and other adults' longevity	\$ -0.04	\$ -0.66	\$ -0.7
Decreased tax payments from parents and other adults ^a	\$ 0	\$ 0	\$ 0
Administrative costs ^b	\$ -0.6	\$ 0	\$ -0.6
Total (without victim cost of crime or tax distortion costs)	\$ -3.9	\$ -0.08	\$ -4.0
Tax distortion costs for taxpayers (inefficiencies due to higher taxation)	\$ -1.68	\$ -0.13	\$ -1.8
Reduced victim cost of crime	\$ 6.97	\$ 1.43	\$ 8.4
Total (with victim cost of crime and tax distortion costs) ^c	\$ 1.39	\$ 1.22	\$ 2.6

Source: Center on Poverty and Social Policy at Columbia University, 2024.

^a Decreased tax payments from parents and other adults are not yet estimated in this analysis as labor supply effects have not yet been provided. While decreased tax payments are typically small, they will be incorporated when these labor supply effects have been estimated.

^b Our estimate of administrative costs is based on the latest report of the Child Poverty Reduction Advisory Council (CPRAC, 2024). In the CPRAC report, it is estimated that the housing voucher proposal (HP1), which is included in all three policy packages analyzed in this report, could incur a maximum of \$546 million of administrative costs. The public benefit proposals included in each of the three packages could incur a total of \$6.1 million of administrative costs. The tax policy proposals impose minimum or no administrative costs. We thus calculate the administrative cost of the three packages as \$546 million + 6.1 million = \$552.1 million, or \$0.6 billion if rounded.

^c Note that, due to rounding procedures, totals might not be an exact sum of their parts.

**Table E3. Breakdown of taxpayers' benefits and costs from Policy Package 3:
Present discounted value (\$2019 billions).**

	State/Local	Federal	All Taxpayers
Total transfer	\$ -8.5	\$ 0	\$ -8.5
Increased future earnings of children	\$ 0	\$ 0	\$ 0
Increased future tax payments by children	\$ 0.94	\$ 3.36	\$ 4.3
Increased children's health and longevity	\$ 0	\$ 0	\$ 0
Increased parents' and other adults' health and longevity	\$ 0	\$ 0	\$ 0
Avoided expenditures on other cash or near-cash transfers	\$ 0.222	\$ 0.078	\$ 0.3
Avoided expenditures on child protection	\$ 0.2	\$ 0.2	\$ 0.4
Avoided criminal justice expenditures	\$ 3.237	\$ 0.663	\$ 3.9
Increased costs of children's education	\$ -0.8	\$ 0	\$ -0.8
Avoided expenditures on children's health care costs	\$ 0.73	\$ 0.97	\$ 1.7
Avoided expenditures on parents' and other adults' health care costs	\$0.0086	\$ 0.0114	\$ 0.02
Increased payment due to increased children's longevity	\$ -0.28	\$ -4.42	\$ -4.7
Increased payment due to increased parents' and other adults' longevity	\$ -0.05	\$ -0.75	\$ -0.8
Decreased tax payments from parents and other adults^a	\$ 0	\$ 0	\$ 0
Administrative costs^b	\$ -0.6	\$ 0	\$ -0.6
Total (without victim cost of crime or tax distortion costs)	\$ -4.89	\$ 0.11	\$ -4.8
Tax distortion costs for taxpayers (inefficiencies due to higher taxation)	\$ -2.06	\$ -0.09	\$ -2.1
Reduced victim cost of crime	\$ 9.3	\$ 1.9	\$ 11.2
Total (with victim cost of crime and tax distortion costs)^c	\$ 2.35	\$ 1.92	\$ 4.3

Source: Center on Poverty and Social Policy at Columbia University, 2024.

^a Decreased tax payments from parents and other adults are not yet estimated in this analysis as labor supply effects have not yet been provided. While decreased tax payments are typically small, they will be incorporated when these labor supply effects have been estimated.

^b Our estimate of administrative costs is based on the latest report of the Child Poverty Reduction Advisory Council (CPRAC, 2024). In the CPRAC report, it is estimated that the housing voucher proposal (HP1), which is included in all three policy packages analyzed in this report, could incur a maximum of \$546 million of administrative costs. The public benefit proposals included in each of the three packages could incur a total of \$6.1 million of administrative costs. The tax policy proposals impose minimum or no administrative costs. We thus calculate the administrative cost of the three packages as \$546 million + 6.1 million = \$552.1 million, or \$0.6 billion if rounded.

^c Note that, due to rounding procedures, totals might not be an exact sum of their parts.

Table E4 presents the bottom-line net changes in benefits and costs for local and federal taxpayers from all three policy packages, excluding and including both the benefits from reduced victim costs of crime and the costs associated with inefficiencies/distortion stemming from higher taxes. When excluding the benefit of reduced victim cost of crime and the costs associated with inefficiencies/distortion stemming from higher taxes, all three packages result in net losses in State taxes. When including the large benefit from reduced victim costs of crime and the costs associated with inefficiencies/distortion stemming from higher taxes, all three packages lead to net gains in both State and Federal taxes. Changes under Package 3 and 1 are similar because they have similar costs. Changes under Package 2 are smaller because Package 2 has smaller costs.

Table E4. Breakdown of total changes in taxpayers' benefits and costs from all three policy packages (\$2019 billions)

	State/Local	Federal	All Taxpayers
Package 1			
(without victim cost of crime or tax distortion costs)	\$-4.92	\$ 0.14	\$ -4.8
(with victim cost of crime and tax distortion costs)	\$ 3.04	\$ 2.11	\$ 5.1
Package 2			
(without victim cost of crime or tax distortion costs)	\$-3.90	\$-0.08	\$ -4.0
(with victim cost of crime and tax distortion costs)	\$ 1.39	\$ 1.22	\$ 2.6
Package 3			
(without victim cost of crime or tax distortion costs)	\$-4.89	\$ 0.11	\$ -4.8
(with victim cost of crime and tax distortion costs)	\$ 2.35	\$1.92	\$ 4.3

Source: Center on Poverty and Social Policy at Columbia University, 2024.

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